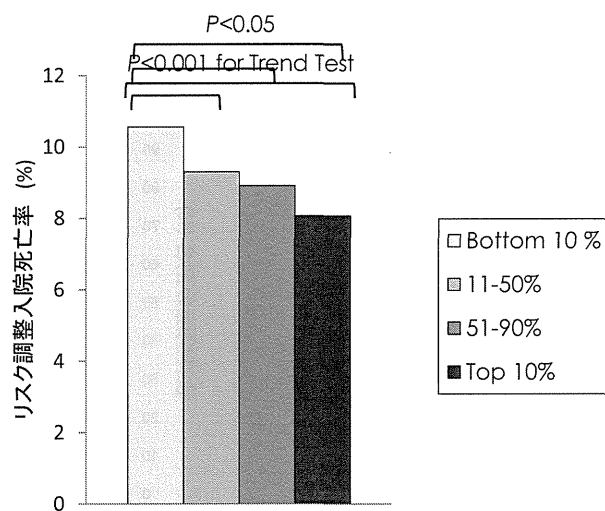
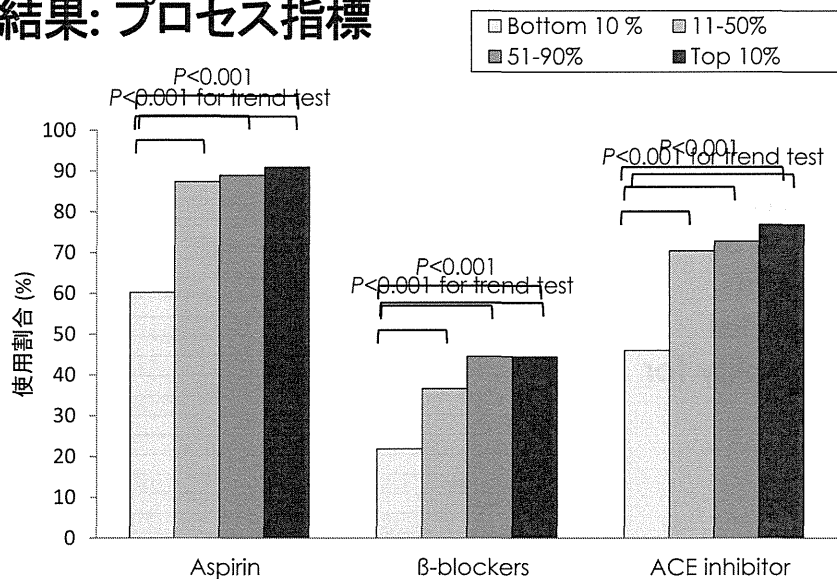


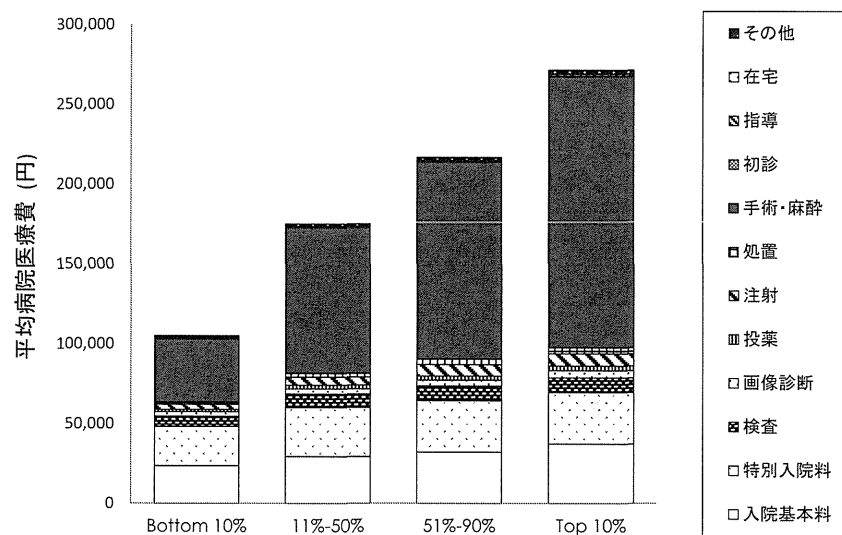
## 結果: 30日以内のリスク調整入院死亡率



## 結果: プロセス指標



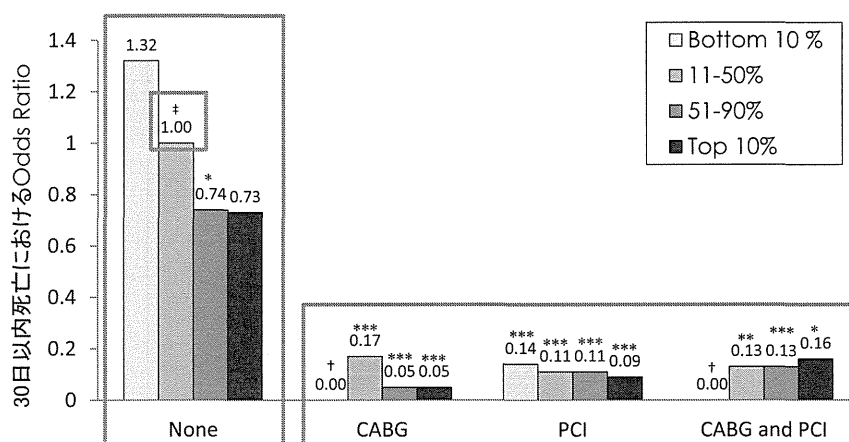
## 結果: グループ間平均医療費の詳細



## 結果: グループ間、手術が行われた件数と割合

|                  | Hospital spending category |               |                |               |
|------------------|----------------------------|---------------|----------------|---------------|
|                  | Bottom 10 %                | 11-50%        | 51-90%         | Top 10%       |
| No. Patients (%) |                            |               |                |               |
| None             | 454 (55.57)                | 2,246 (23.94) | 2,236 (16.97)  | 470 (14.53)   |
| CABG             | 0 (0.00)                   | 127 (1.35)    | 239 (1.81)     | 72 (2.23)     |
| PCI              | 363(44.43)                 | 6,973 (74.34) | 10,571 (80.25) | 2,659 (82.22) |
| CABG and PCI     | 0 (0.00)                   | 34 (0.36)     | 127 (0.96)     | 33 (1.02)     |

## 結果: グループ間、死亡における手術の効果



\* $p < 0.05$ , \*\* $p < 0.01$ , \*\*\* $p < 0.001$ .

†No case was observed in the category

‡Referent category

## 考察: 病院医療費と医療の質

- 急性心筋梗塞における医療の質と病院医療費の関係
  - 病院医療費とプロセスとアウトカムに関する医療の質との正の関連が見られた
  - 医療費が一番低い病院グループでは、医療の質が低かった
  - ある程度医療費が高い病院グループでは、医療の質に差が大きくなかった

## 考察: 手術の効果

- 死亡における手術(CABG and PCI)の効果
  - 医療費を増加させる要因であり、死亡低下と関連することがわかった
  - 医療費が一番低い病院グループでは、手術が死亡低下に大きく影響すると考えられる
  - 医療費が高いグループでは、手術の効果があるが、医療費の増加と死亡率低下の関連は比較的小さくなると考えられる
  - 本研究での主題ではないため、費用効果的かはわからない

## 考察: 病院ストラクチャー病院医療費、医療の質との関係

- アウトカムと病院ストラクチャーとの関係
  - 急性心筋梗塞症例数<sup>1-4</sup>、教育病院<sup>4-7</sup>、医師当たり症例数と病院症例数<sup>8</sup>、循環器治療設備の有無<sup>4</sup>
- 特に、医療費が低い病院では
  - 医療行為を行うために人的資源(特に専門医)、医療機器、施設(心カテ室やCCUなど)が少ないと考えられる
  - 急性心筋梗塞に特化した治療が提供できる医師と看護師のチームケアがないか少ないと考えられる
  - PCIやCABGなど手術が十分に行われないと考えられる

<sup>1</sup>Canto, et al. 2000. <sup>2</sup>Halm, et al. 2002. <sup>3</sup>Thiemann, et al. 1999. <sup>4</sup>Bradley, et al. 2010. <sup>5</sup>Polanczyk, et al. 2002. <sup>6</sup>Allison, et al. 2000. <sup>7</sup>Ayanian, et al. 2002. <sup>8</sup>Vakili, et al. 2001. <sup>9</sup>Fuchs, et al. 2011.

## 限界

- 自発的な病院の参加からの選択バイアス(Selection bias)
- 計りにくい要因は調節されていない<sup>1</sup>
- リスクが低い患者が病院や医師に受けやすい紹介バイアス(Referral bias)
- 転送された患者も含まれている可能性

<sup>1</sup>Zuckerman et al. 2010.

## 結論

- 急性心筋梗塞においては、病院医療費の低さと医療の質の低さに関係が見られた
- 病院ストラクチャーが医療費と医療の質に影響を及ぼすことがわかった

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## Funding Sources

- Health Sciences Research Grant from the Ministry of Health, Labour and Welfare of Japan
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御静聴、ありがとうございます

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Sungchul Park  
Correspondence to Professor 今中 雄一  
京都大学 大学院医学研究科 医療経済学分野

「医療経済学会」 第7回研究大会

構造化抄録 フォーマット

|       |  |
|-------|--|
| 申込者   | Sungchul PARK, Yuichi IMANAKA, Jason LEE, Hiroshi IKAI, Tetsuya OTSUBO, Naoto UKAWA  |
| 所属・役職 | Department of Healthcare Economics and Quality Management, Graduate School of Medicine, Kyoto University   |
| 一般演題名 | Quality of Care and Hospital Spending in Acute Myocardial Infarction: Evidence from Japan  |
| 1. 背景 | The relationship between health care spending and quality of care has important implications for policy makers as it helps them improve the long-term sustainability of health care financing and quality of care. However, this relationship has received little attention at the level of hospital spending.   |
| 2. 目的 | The overall goal of this study was to determine the relationship between hospital spending and quality of care among patients with acute myocardial infarction (AMI) in Japan, and to investigate hospital structure in regards to spending and quality of care.   |
| 3. 方法 | We utilized administrative data from patients admitted to 180 hospitals in Japan for AMI from 2008 to 2011. Quality of care for AMI was assessed using process as well as outcome measures: 30-day in-hospital risk-adjusted mortality rates, all-cause readmission rates; and the percentage of patients prescribed aspirin, $\beta$ -blockers, and angiotensin-converting enzyme (ACE) inhibitor during hospitalization. Multilevel logistic regression models were developed based on patients clustered within hospitals, with independent variables inclusive of patient-level risk factors such as age, gender, co-morbidities and infarct location; hospital characteristics such as teaching status, hospital ownership, hospital bed size, AMI case volume, and the number of physicians as well as nurses. To compare the quality of care among hospitals, hospitals were classified according to mean hospital spending, which was divided into four categories: bottom 10%, 11-50%, 51-89%, and top 10% of hospital spending. ANOVA with Bonferroni correction for multiple comparisons was conducted to test differences in process as well as outcome indicators among the categories of hospital spending. Also, Jonckheere-Terpstrates were conducted to analyze trends in these indicators across the categories. |
| 4. 結果 | After adjustment for patient and hospital characteristics, the mortality rate decreased from 7.67% for the hospitals in the bottom 10% to 4.20% for those in the top 10% ( $P<0.01$ : test for trend). On the other hand, no statistical significance was found in the readmission rates after adjusting for patient and hospital characteristics across the categories of hospital spending ( $P=0.161$ : test for trend). Hospitals with higher spending were more likely to be associated with better quality of care in process indicators, and the trends across the categories of hospital spending were found to be statistically significant   |
| 5. 考察 | Hospitals with few health care resources were associated with lower spending and poorer quality of care in process and outcome among AMI patients. Our findings may help policy makers become aware of the relationship between hospital spending and quality of care when designing affordable health care reforms. (403 words)   |



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## Clinical Research

# Development and Validation of an Acute Heart Failure-Specific Mortality Predictive Model Based on Administrative Data

Q5 Noriko Sasaki, MD, Jason Lee, PhD, Sungchul Park, BA, Takeshi Umegaki, MD, PhD, Susumu Kunisawa, MD, Tetsuya Otsuq, PhD, Hiroshi Ikai, MD, PhD, and Yuichi Imanaka, MD, PhD

Department of Healthcare Economics and Quality Management, School of Public Health, Kyoto University Graduate School of Medicine, Kyoto, Japan

**ABSTRACT**

**Background:** Acute heart failure (AHF) with its high in-hospital mortality is an increasing burden on healthcare systems worldwide, and comparing hospital performance is required for improving hospital management efficiency. However, it is difficult to distinguish patient

Acute heart failure (AHF) requiring hospitalization is associated with high rates of morbidity and mortality.<sup>1-3</sup> Several recent AHF registries and surveys have reported in-hospital mortality rates in AHF patients ranging from 3.8% to 7.7%.<sup>4-7</sup> Aging of the population, progression of therapeutic intervention, and effective secondary prevention have all led to an increasing burden on heart failure and AHF on health care systems worldwide.<sup>8,9</sup> The comparison of hospital performance and quality of care is an initial step to evaluate, benchmark, and improve hospital management under the growing health care costs associated with AHF. However, it is difficult to distinguish between the influences of patient disease severity from individual hospital care effects, thereby impeding adequate comparison of hospitals.

Because some hospitals treat sicker patients than others, patient severity should be taken into consideration when comparing hospitals. The comparison of crude mortality rates between facilities would bias evaluations against hospitals with a greater proportion of high risk patients, and risk-adjusted mortality rates can make hospital-level comparisons more meaningful.<sup>10</sup> Risk adjustment accounts for the differences in

**RÉSUMÉ**

**Introduction :** En raison de sa mortalité hospitalière élevée, l'insuffisance cardiaque aiguë (ICA) est un fardeau de plus en plus lourd pour les systèmes de soins de santé à l'échelle mondiale. Ainsi, la comparaison de la performance hospitalière est nécessaire pour améliorer

intrinsic patient health risks at admission. To this end, administrative data are appealing because of its ability to derive numerous variables from a routine work flow, and the relatively large quantity of data available which allow inter-hospital comparisons.

However, because the real-world diagnosis of AHF is highly complex<sup>2</sup> and administrative data have limitations in acquiring clinical variables that influence patient outcomes, the usage of administrative data for risk adjustment among AHF patient groups has been restricted and appears to be challenging thus far.

The aim of our study was to develop an accurate, practical, and reproducible risk adjustment model to predict AHF in-hospital mortality using factors identifiable from administrative data, and to apply this model to an interhospital comparison.

**Methods****Data source**

All data were extracted from the Quality Indicator/Improvement Project (QIP), a project that involves the collection of administrative data from voluntarily participating acute care hospitals, and subsequent analysis of healthcare processes, patient outcomes, and disease management in Japan.<sup>11,12</sup> Participating hospitals vary widely in patient volume, bed numbers, region, and type (publicly- or privately-owned; teaching or nonteaching). Moreover, QIP hospitals provide administrative data based on the Japanese case-mix classification system, known as the Diagnosis Procedure

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Corresponding author: Dr Yuichi Imanaka, Department of Healthcare Economics and Quality Management, School of Public Health, Kyoto University Graduate School of Medicine, Yoshida Konoe-cho, Sakyo-ku, Kyoto 606-8501, Japan. Tel.: +81-75-753-4454; fax: +81-75-753-4455.

E-mail: imanaka-y@umin.net

See page xxx for disclosure information.



severity from individual hospital care effects. The aim of this study was to develop a risk adjustment model to predict in-hospital mortality for AHF using routinely available administrative data.

**Methods:** Administrative data were extracted from 86 acute care hospitals in Japan. We identified 8620 hospitalized patients with AHF from April 2010 to March 2011. Multivariable logistic regression analyses were conducted to analyze various patient factors that might affect mortality. Two predictive models (models 1 and 2; without and with New York Heart Association functional class, respectively) were developed and bootstrapping was used for internal validation. Expected mortality rates were then calculated for each hospital by applying model 2.

**Results:** The overall in-hospital mortality rate was 7.1%. Factors independently associated with higher in-hospital mortality included advanced age, New York Heart Association class, and severe respiratory failure. In contrast, comorbid hypertension, ischemic heart disease, and atrial fibrillation/flutter were found to be associated with lower in-hospital mortality. Both model 1 and model 2 demonstrated good discrimination with c-statistics of 0.76 (95% confidence interval, 0.74-0.78) and 0.80 (95% confidence interval, 0.78-0.82), respectively, and good calibration after bootstrap correction, with better results in model 2.

**Conclusions:** Factors identifiable from administrative data were able to accurately predict in-hospital mortality. Application of our model might facilitate risk adjustment for AHF and can contribute to hospital evaluations.

Combination (DPC).<sup>12</sup> The DPC-based hospital reimbursement system was introduced in 2003, and has been adopted by more than 1400 hospitals by 2011, accounting for more than half of the total 910,000 hospital beds nationwide. This payment scheme is based on per diem charges. The DPC system database includes information on hospital codes, patient demographic characteristics, admission and discharge dates, admission routes, outcomes, primary and secondary diagnoses, comorbidities at admission, complications, surgeries performed, and high cost procedures such as mechanical ventilation, hemodialysis, and cardiopulmonary support device use. Diagnoses including comorbidities and complications are coded by physicians based on International Classification of Diseases, 10th revision (ICD-10) codes. Furthermore, AHF was specifically identified using the 'acute exacerbation' code available in DPC data, which has been determined by the attending physician at admission. Similarly, the reporting of the New York Heart Association (NYHA) functional class at admission by physicians is mandatory within the DPC system.

### Study population

Data were collected from 19,792 patients across 139 hospitals with a primary diagnosis of heart failure (ICD-10 code I50.x). Patients were included in the study if they had been discharged between April 1, 2010 and March 31, 2011; and had been admitted to hospitals that had continuously provided data

l'efficacité de la logistique hospitalière. Toutefois, il est difficile de distinguer la gravité de l'état de santé du patient des effets des soins hospitaliers individuels. Le but de cette étude était de développer un modèle d'ajustement au risque pour prédire la mortalité hospitalière par ICA en utilisant les données administratives couramment disponibles.

**Méthodes :** Les données administratives ont été extraites de 86 hôpitaux de soins de courte durée du Japon. Nous avons sélectionné 8620 patients hospitalisés ayant une ICA d'avril 2010 à mars 2011. Des analyses multivariées de régression logistique ont été menées pour analyser les divers facteurs liés aux patients qui pourraient influencer la mortalité. Deux modèles prédictifs (modèles 1 et 2; sans et avec la classification fonctionnelle de la New York Heart Association, respectivement) ont été élaborés, et l'autoamorçage (*bootstrapping*) a été utilisé pour la validation interne. Les taux de mortalité attendus ont ensuite été calculés pour chaque hôpital par l'application du modèle 2.

**Résultats :** Le taux global de mortalité hospitalière a été de 7,1 %. Les facteurs indépendamment associés à la mortalité hospitalière élevée ont inclus l'âge avancé, la classification de la *New York Heart Association* et l'insuffisance respiratoire grave. En revanche, l'hypertension associée à une maladie, la cardiopathie ischémique, et la fibrillation et le flutter auriculaires ont été associés à une plus faible mortalité hospitalière. Les modèles 1 et 2 ont démontré par les statistiques C une bonne discrimination de 0,76 (intervalle de confiance à 95 %, 0,74-0,78) et de 0,80 (intervalle de confiance à 95 %, 0,78-0,82), respectivement, et une bonne calibration après la correction par l'autoamorçage par des meilleurs résultats au modèle 2.

**Conclusions :** Les facteurs identifiables des données administratives ont été en mesure de prédire avec précision la mortalité hospitalière. L'application de notre modèle pourrait faciliter l'ajustement au risque pour l'ICA et peut contribuer aux évaluations hospitalières.

during the 12-month study period. The following selection criteria were also used: (1) patients who had both an 'acute exacerbation' of heart failure code and NYHA functional class II or higher, which were available within the DPC system and (2) patients who were at least 20 years of age at admission. The selection yielded 11,503 patients from 134 hospitals. Patients were excluded from the analysis if they simultaneously had acute myocardial infarction or if they had other conditions indistinguishable from AHF (n = 2211), including cardiopulmonary arrest (ICD-10 codes: I46.1, I46.9, R96), acute respiratory distress syndrome (ICD-10 code: J80), severe pneumonia (ICD-10 codes: J10.0, J11.0, J12-J18, J69), pleuritis (ICD-10 codes: A15.6, A16.5, R09.1, J90, J91, J94.x), and severe renal failure (ICD-10 code: N18.0) with or without dialysis at admission. Patients with a length of stay longer than 3 standard deviations from the mean (n = 169) and an invalid mortality record (n = 1) were also excluded from the analysis.

Because of wide variations in hospital volume and available emergency care, hospitals with fewer than 20 registered cases and those with no recorded utilization of acute mechanical ventilation during the study year were also excluded (46 hospitals with 545 patients). This resulted in a final sample size of 8620 patients from 86 hospitals ranging from 21 patients to 317 patients at the hospital level.

This study was approved by the Institutional Review Board of the Faculty of Medicine at the Graduate School of Medicine of Kyoto University, Japan.

## Statistical analysis

In-hospital mortality rate was used as the primary outcome measure. Two types of mortality prediction models (model 1 and model 2; using identical predictors without and with NYHA functional class, respectively) were constructed with multivariable logistic regression using the original dataset (training set). Discrimination of the logistic regression models was assessed using the *c*-statistic.<sup>13</sup> Bootstrapping was used to assess the internal validation of the model. We used 1000 bootstrap resamples to evaluate the reliability of the regression coefficients and the *c*-statistics.<sup>14</sup> To validate the prediction model, a model in the bootstrap sample dataset (*n* = 8620) was derived and Hosmer-Lemeshow test was performed to evaluate model calibration (*P* > 0.05 is considered favourable).<sup>15</sup>

During model construction, we explored clinically and potentially important predictors available in the database as candidate explanatory variables (Table 1). These variables were categorized into 3 fields of measurement: demographic characteristics, clinical factors associated with patient severity, and comorbidities at admission. Multivariable logistic regression was used to estimate the predictors of in-hospital mortality of the original full dataset. We retained all covariates in the final model.

Clinical factors related to patient severity were defined as follows: admission route was identified using an “emergency” admission code, which allowed for the identification of

**Table 1.** Candidate variables used to develop the in-hospital mortality prediction model

| Candidate variables   | Category   |
|---|--|
| Demographic characteristics   |  |
| Sex   | Male*, female  |
| Age (y)   | 20-59*<br>60-69<br>70-79<br>80-89<br>≥90                                       |
| Hospital admission route  |  |
|   | 1. Emergency with ambulance<br>2. Emergency without ambulance<br>3. Scheduled* |
| Clinical factors  |  |
| NYHA functional class   | II*; III; IV   |
| Severe respiratory failure because of AHF                                       | 0, Absent; 1, present  |
| Comorbidities   |  |
| Ischemic heart disease (ICD-10 codes: I201, I208, I209, I25)                    | 0, Absent; 1, present  |
| Hypertension (including HHD; ICD-10 codes: I10-I15)                             |  |
| Atrial fibrillation/flutter (ICD-10 code: I48)                                  |  |
| Life-threatening arrhythmia (ICD-10 codes: I490, I442, I46)                     |  |
| Chronic renal failure (mild to moderate; ICD-10 codes: N188, N189, N19)         |  |
| Shock (including cardiogenic shock; ICD-10 codes: R570, R571, R578, R579, A419) |  |

AHF, acute heart failure; HHD, hypertensive heart disease; ICD-10, *International Classification of Diseases, 10th revision*; NYHA, New York Heart Association.

\*Reference value.

unplanned admissions with or without ambulance use. Severe respiratory failure because of AHF was identified with procedure codes reflecting acute mechanical ventilation use within 48 hours postadmission. Potential comorbidities were selected from the previous published literature as risk factors for mortality.<sup>1,3,4</sup> However, available diagnoses of comorbidities in each patient were restricted to 4 coding slots. Comorbidities with a prevalence of less than 0.7% were excluded because of possible undercoding.<sup>10</sup>

Finally, the predicted mortality was calculated using the coefficients derived from the average of the bootstrapping datasets. All statistical analyses were conducted using SPSS software, version 19.0J (SPSS Inc, Chicago, IL) and STATA 12 statistical software (StataCorp, College Station, TX).

## Evaluation of hospitals

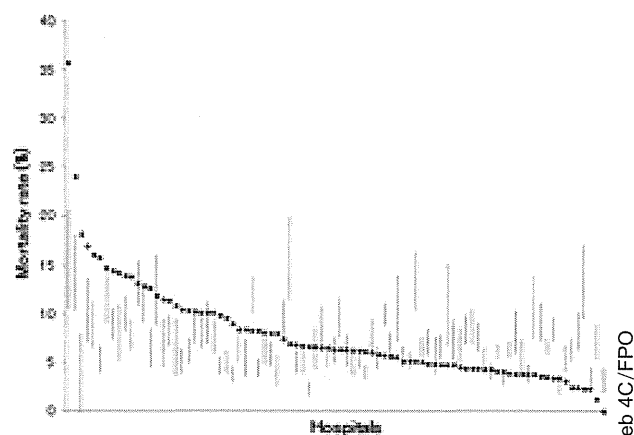
We identified hospitals as having a better or worse observed mortality rate than their expected mortality rate based on 95% confidence interval (CI). The expected mortality rate of each hospital was calculated using model 2 by adding the predicted mortality risk of each patient within an individual hospital and dividing the sum by the number of patients (Fig. 1).

## Results

Hospital characteristics and patient clinical baseline characteristics are shown in Table 2. Overall observed in-hospital mortality rate was 7.1%, which was within the range reported in recent AHF registries.<sup>4-7</sup> The mean age was 78 years, with minimal differences in sex. The prevalence of hypertension (including hypertensive heart disease) was approximately 57%, and that of ischemic heart disease (IHD) and atrial fibrillation/flutter were approximately 34% and 29%, respectively.

## Predictive model of in-hospital mortality in patients with AHF

Table 3 shows the logistic regression coefficients and adjusted odds ratio with corresponding 95% CI for the final



**Figure 1.** The dots represent observed in-hospital mortality rates of acute heart failure in individual hospitals. The lines represent 95% confidence intervals of the expected mortality rate in individual hospitals. Model 2 (with NYHA functional class) was adopted for risk adjustment in the figure.

**Table 2. Demographic characteristics of the original dataset**

|  | Original dataset |
|--|------------------|
| <b>Facility level characteristics</b>            |                  |
| Number of hospitals                              | 86               |
| Teaching, n (%)                                  | 73 (84.9)        |
| Larger beds (>380), n (%)                        | 43 (50.0)        |
| <b>Patient level characteristics</b>             |                  |
| Number of patients                               | 8620             |
| Female, n (%)                                    | 4318 (50.1)      |
| Age (y, mean ± SD)                               | 78.5 ± 12.0      |
| <b>Admission route, n (%)</b>                    |                  |
| Emergency with ambulance                         | 2598 (30.1)      |
| Emergency without ambulance                      | 4396 (51.0)      |
| Scheduled  | 1626 (18.9)      |
| <b>NYHA functional class, n (%)</b>              |                  |
| II   | 2482 (28.8)      |
| III  | 3298 (38.3)      |
| IV   | 2840 (32.9)      |
| <b>Severe respiratory failure because of AHF</b> |                  |
|  | 601 (7.1)        |
| <b>Comorbidities, n (%)</b>                      |                  |
| Ischemic heart disease                           | 2929 (34.0)      |
| Hypertension (including HHD)                     | 4933 (57.2)      |
| Atrial fibrillation/flutter                      | 2534 (29.4)      |
| Life-threatening arrhythmia                      | 193 (2.2)        |
| Chronic renal failure (mild to moderate)         | 956 (11.1)       |
| Shock (including cardiogenic shock)              | 97 (1.1)         |
| <b>Clinical outcome</b>                          |                  |
| In-hospital mortality, n (%)                     | 611 (7.1)        |

AHF, acute heart failure; HHD, hypertensive heart disease; NYHA, New York Heart Association.

equation of the validation dataset. Factors that were independently associated with higher in-hospital mortality included advanced age, NYHA functional class, severe respiratory failure because of AHF, shock, life-threatening arrhythmia including ventricular fibrillation/flutter or sinus arrest, and mild to moderate chronic renal failure. In contrast, hypertension, IHD, and atrial fibrillation/flutter were associated with lower in-hospital mortality.

The c-statistics for the training models of model 1 and model 2 were 0.77 (95% CI, 0.75-0.79) and 0.81 (95% CI, 0.79-0.82), respectively. In the sample dataset, the discriminative ability of the model was maintained with the c-statistics at 0.76 (95% CI, 0.74-0.78) and 0.80 (95% CI, 0.78-0.82), respectively, after bootstrap correction. The Hosmer-Lemeshow test showed no significance ( $P = 0.44$  and  $P = 0.88$ , respectively), indicating good calibration.

### Comparing hospital performance

The observed mortality rates ranged from 0.0% to 35.7% for all 86 hospitals. Figure 1 shows the observed mortality rates for each hospital in descending order. Graph lines represent the 95% CI of the expected mortality rates calculated using the logistic regression model. Model 2 with NYHA functional class was adopted to calculate the expected mortality rate because of its better predictivity. Hospital performance was evaluated based on whether the observed mortality rate was within, higher, or lower than the 95% CI range of the expected mortality rate.

Generally speaking, measured outcome has 4 basic components: (1) intrinsic patient specific risk, (2) quality of

**Table 3. Multivariable predictors of in-hospital mortality**

| Variables                                 | Adjusted odds ratio (95% CI) |                    |
|---|------------------------------|--------------------|
|   | Model 1                      | Model 2            |
| Female                                    | 0.96 (0.80-1.15)             | 0.96 (0.80-1.15)   |
| Age (reference, 20-59 y)                  |                              |                    |
| 60-69                                     | 1.32 (0.72-2.41)             | 1.32 (0.71-2.44)   |
| 70-79                                     | 2.21 (1.29-3.79)*            | 2.24 (1.30-3.86)*  |
| 80-89                                     | 4.10 (2.44-6.87)†            | 4.15 (2.46-6.99)†  |
| ≥90                                       | 7.53 (4.42-12.82)‡           | 7.47 (4.36-12.79)‡ |
| Admission route (reference, scheduled)    |                              |                    |
| Emergency with ambulance                  | 1.39 (1.06-1.84)‡            | 1.09 (0.82-1.45)   |
| Emergency without ambulance               | 1.11 (0.85-1.44)             | 1.00 (0.76-1.31)   |
| NYHA functional class at admission§       |                              |                    |
| III                                       | —                            | 2.28 (1.66-3.12)†  |
| IV  | —                            | 5.67 (4.20-7.65)†  |
| Severe respiratory failure because of AHF |                              |                    |
| Ischemic heart disease                    | 0.58 (0.47-0.71)†            | 0.57 (0.47-0.70)†  |
| Hypertension                              | 0.28 (0.23-0.34)†            | 0.29 (0.24-0.36)†  |
| Atrial fibrillation/flutter               | 0.61 (0.49-0.76)†            | 0.64 (0.52-0.79)†  |
| Life-threatening arrhythmia               | 2.04 (1.34-3.10)*            | 1.93 (1.26-2.95)*  |
| Chronic renal failure (mild to moderate)  | 1.59 (1.25-2.01)†            | 1.53 (1.20-1.95)*  |
| Shock                                     | 3.36 (2.08-5.40)‡            | 2.86 (1.71-4.76)‡  |
| C-statistics (95% CI)                     | 0.76 (0.74-0.78)             | 0.80 (0.78-0.82)   |

Model 1 does not include NYHA class; model 2 includes NYHA class. AHF, acute heart failure; CI, confidence interval; NYHA, New York Heart Association.

\*  $P < 0.01$ .

†  $P < 0.001$ .

‡  $P < 0.05$ .

§ Reference: NYHA II.

care provided, (3) random variation, and (4) bias introduced by systematic errors in measurement.<sup>16</sup> Because our model showed high predictivity based on intrinsic patient risk, random variation is considered to be minimized. Systematic error in measurement was assumed to be negligible when comparing hospital performance. Consequently, a difference observed outside the 95% CI range of the expected outcomes measured for a single organization is considered to reflect the real differences between the organization and the reference standard in the quality of care provided. In other words, unexplained differences between expected outcomes and observed outcomes might reveal unwarranted institutional variations.

## Discussion

### Predictive model based on routinely available administrative data

In the present study, a risk adjustment model for AHF in-hospital mortality was developed using DPC administrative data in Japan. Our model was designed to account for differences in intrinsic patient health risks for assessing clinical performance of acute care hospitals.

A number of risk stratification models or scores for AHF using clinical data have been reported, mainly as beneficial tools for supporting clinical decision-making including initial triage or effective treatment.<sup>1,4,17-19</sup> However, there is a lack of clinically plausible and feasible risk adjustment methods in the interest of evaluating and comparing multiple hospital

performance in this field, particularly when using administrative databases.

With the nationwide spread of the DPC administrative data system in Japan, disease-specific risk adjustment methods will become more useful and practical for hospital management intending to improve quality of care. It would be labourious, costly, and time-consuming for physicians, researchers, insurers, and policymakers to collect laboratory data or other clinical findings in addition to administrative data. Although we had only used data available from the administrative database in this study, our model was shown to reliably predict in-hospital mortality in AHF patients.

There are several previous studies in which in-hospital mortality of acute myocardial infarction patients or patients with other disease was accurately predicted by complementing administrative data with present on admission (POA) modifiers for secondary diagnosis.<sup>18</sup> The high predictivity of our model was assumed to be partly because of the fact that this POA information is contained and routinely available in the DPC data system.

Furthermore, in-hospital mortality predictors detected in the current study were partially consistent with previous clinical reports. Advanced patient age and mild to moderate chronic renal failure were found to be associated with increased risk of in-hospital mortality in both this study and in previous studies.<sup>1,4,5</sup> Advanced NYHA class at admission and shock were also reported to be associated with higher mortality.<sup>3,19</sup>

Our model differs from other models used to predict AHF mortality in a number of important ways. A unique difference is that essential predictors of patient severity such as NYHA functional class and severe respiratory failure because of AHF on admission were included as independent variables despite using only administrative data, as the DPC system provides POA information as part of the database. Mechanical ventilation use within the first 48 hours of admission could reflect, in part, care provided in the hospital, and the inclusion of this variable might be debatable. However, we assumed that the clinical decision to intubate might not be extremely different among physicians facing critically ill patients. Considering the inevitable limitations in specifying the exact severity only from diagnoses, we used this variable included in administrative data as a surrogate for severe respiratory failure because of AHF. Moreover, the variables can be easily obtained, because they are continuously generated through a routine work flow.

A second unique characteristic is that despite the lack of precise clinical data and the possibility of undercoding, the model was able to reveal that hypertension was highly associated with a lower risk of in-hospital mortality. Hypertension is usually considered to be 1 of the most common precursors and the most frequent underlying disease in patients with AHF. Although elevated blood pressure (BP) is an increased risk of developing heart failure in the general population, recent studies have shown that higher BP on admission is associated with lower risk of dying.<sup>1,4,20,21</sup> Indeed, high BP on admission does not always imply antecedent hypertension, and considering that previous hypertension had no independent influence on in-hospital mortality in another report,<sup>20</sup> the result might require future examination.

In addition, IHD and atrial fibrillation/flutter were also found to be associated with lower in-hospital mortality in

this study. This might be because of the following reasons. In the case of IHD, there are previous reports consistent with our study, indicating coronary revascularization status might be associated with improved early survival.<sup>22</sup> Next, although new-onset atrial fibrillation has been reported to increase in-hospital mortality,<sup>23</sup> there is no compelling evidence to show the prognostic value of previous atrial fibrillation in patients with AHF. It would be debatable to interpret these factors as solely having protective effects, because the results might also reflect undercoding of patients who died in hospitals.<sup>24</sup> However, hypertension and history of coronary angioplasty have shown possible protective effects in administrative claims and chart-based models,<sup>25</sup> which might partly support our results. Because our sample has limitations in obtaining precise clinical information, further studies are also required to evaluate this issue.

Finally, despite the relatively small number of variables used in the present study, the model performance was noteworthy. The *c*-statistics of our model was approximately 0.8 after bootstrap correction. The value was at least equal or superior to previous studies using variables derived from chart reviews including comorbidities and clinically-extracted data such as symptoms, vital signs, physical examination findings, laboratory test results and multiple therapies.<sup>1,3,4,17,25</sup> The *c*-statistics of these studies based on chart reviews ranged from 0.71 to 0.84. The results from this study might imply possible applications of our model in the risk adjustment of wider AHF populations in the future.

### Implications of hospital performance evaluation using the predictive model

The accurate prediction of hospital mortality rates for AHF using routinely available administrative data would lead to increased use of risk-adjusted outcome as a quality indicator. The staff of individual hospitals can assess their quality of care by analyzing the disparity between the 95% CI of expected mortality rates and observed mortality rates, or by comparing their outcomes with other hospitals. Periodical and continual measurement will help hospitals self-monitor their quality of care. If a facility's performance is consistently an outlier when compared with other hospitals, that facility would require greater scrutiny.

The mortality prediction model in this study might provide a feasible and low-cost alternative to the labour-intensive chart review approach for the evaluation of multiple hospital performance, especially in the management of patients with AHF.

### Limitations

There are several limitations in the present study. First, AHF has been referred to as "heterogeneous syndromes,"<sup>2</sup> varying in case identification with multiple types of data sources, leading to difficulties in straightforward comparisons with previous studies. There are also several detailed aspects of AHF identification that remain unclear because we could not collect clinical data such as left ventricular systolic function, serum BNP level or other factors that are considered to be critical to the heart failure prognosis.<sup>3</sup>

Second, the designation of NYHA functional class at admission by the attending doctors might not be completely

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reliable in all cases, because these attending doctors might include noncardiologists. However, because NYHA class is rarely available in administrative data in other countries, the effect and applications of NYHA class shown here might be informative for people involved in the development and analyses of these databases.

Third, the study population was restricted to AHF patients from acute care hospitals voluntarily participating in the QIP. A selection bias might have occurred by only comparing hospitals willing to participate in this program. However, the large number and diverse characteristics of QIP participant hospitals might reduce the effect of this bias.

Fourth, the coding slots for comorbidities are limited to only 4 slots in the DPC system, which might result in possible undercoding. Refinement of the coding system will be required in order to further improve subsequent research quality.

Finally, there are still concerns with using administrative data as the sole data source, as opposed to including any kind of clinical data. The inability to obtain and describe in detail the specific clinical conditions of each individual patient is a fundamental limitation of administrative data. Therefore, the validity of risk adjustment using administrative data alone has been repeatedly challenged, and the results of several model comparisons have been reported.<sup>25,26</sup> Although these reports have advocated the addition of clinical data to administrative data-based analyses, it has also been shown that difficult-to-obtain key clinical findings add little to predictive power or risk-adjustment equations.<sup>18</sup> In the present study, POA information that is already included in the DPC administrative database proved to be a useful alternative source of clinical data.

### Conclusions

Despite the relatively small number of variables used in the current models, the factors identifiable from routinely-available administrative data were able to predict in-hospital mortality for AHF with a high level of discrimination. Our models facilitate risk adjustment of AHF patients and might contribute to evaluating quality of care among multiple hospitals related to AHF.

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### q3 Disclosures

The authors have no conflicts of interest to disclose.

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631 **SUMMARY**  
632 We developed feasible and accurate risk adjustment models of in-  
633 hospital mortality in acute heart failure (AHF) patients using only  
634 administrative data. The accuracy of our models was excellent, espe-  
635 cially when adjusted for New York Heart Association class. Our model  
636 facilitates risk adjustment of AHF patients and might contribute to  
evaluating quality of care among multiple hospitals related to AHF.

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## 病院管理学会 2012

題名：急性心不全患者の院内死亡予測モデル開発とリスク調整死亡率の病院間比較

佐々木典子、Jason Lee、國澤進、大坪徹也、猪飼宏、今中雄一

【背景】急性心不全は死亡率・再入院率とも高く、高齢化に伴いますます患者が増加すると予測され、医療費増加の大きな一因と考えられている。しかしながら、急性心不全の病態は多彩であり、これまでDPCデータのみから対象患者を同定することは困難で、病院間の医療提供内容やアウトカム指標を比較検討することは不可能だった。【目的】新規に導入された病名付加コードから急性心不全患者を同定し、大規模施設を中心とした29病院が参加する急性心不全の臨床的な多施設症例登録ATTENDレジストリー結果(2010年報告)と比較して、対象集団の疫学的位置づけを確認する。その後、多数の病院の医療提供内容・質につき妥当な比較を行うために、院内死亡率をアウトカムとして、重症度など患者因子でリスク調整を行うモデルを作成し、モデルの妥当性につき検討する。また、同モデルを用いたリスク調整死亡率の病院間比較を行う。

【方法】「医療資源を最も投入した傷病名」が「心不全」(DPCコード050130)で、病勢を表す病名付加コード(急性:30101、慢性の急性増悪:30102)があり、2010年4月1日から2011年3月31日の12ヶ月間データのあるQuality Indicator/Improvement Project(QIP;自発的参加のベンチマーキング事業)登録90急性期病院より、同期間内に退院した9557件の患者レコードを抽出した。尚、「主傷病名」「入院時併存症」に来院時心肺停止、成人呼吸促迫症候群、重症肺炎、胸膜炎、末期腎不全がある症例、手術病名に透析導入関連手術のある症例、在院日数が極端に長い症例、データ不整合症例、NYHA分類欠損症例及び20歳以下の症例は除外した。また20症例以下の施設および急性期に人工呼吸器使用歴のない施設は除外した。まず、90施設を病床規模で対応させてランダムに2群に分け、全9557症例をtest dataset(n=4861)とvalidation dataset(n=4696)に分割した。年齢、性別のほか、入院時の重症度を反映すると考えられるNYHA分類、救急搬送の有無、48時間以内人工呼吸器使用の有無、各種併存症などDPCデータで得られる変数から、ロジスティック回帰分析(ステップワイズ法)を用いて前者で院内死亡予測モデルを作成し、後者で妥当性を検証した。全病院の粗死亡率とリスク調整死亡率を算出し、病院間比較を行った。【結果】全集団における院内粗死亡率は7.0%で、基礎疾患の約3割を占める虚血性心疾患(急性心筋梗塞は除く)割合とともに、ATTENDレジストリーと類似の結果となった。今回のデータでは大学病院は含まれなかったが、ATTEND参加病院よりも多数、病院規模も多様で、構成患者はより高齢で、性差が少なく、よ



り軽症な患者が多かった。test dataset, validation dataset の院内死亡予測モデルの C-statistics は各々0.82,0.80 で、Hosmer-Lemeshow  $\chi^2$  検定で 1.07 (P = 0.96)を示し、適合度は良好だった。病院間比較結果は供覧する。【考察】「急性」を示す病名付加コードにより、従来困難だった急性心不全を同定することができた。対象集団の疫学的位置づけは、比較した ATTEND レジストリーが大学病院など積極的に加療する施設を主とした集団であり、解釈に注意を要する。本研究で開発した院内死亡予測モデルの予測力は高く、今後の応用が期待できる。QIP 参加病院は医療の質改善に積極的な傾向があるなどの交絡については留意が必要である。【結論】本研究で開発した急性心不全患者の院内死亡予測モデルの予測力は高く、高額医療につながる急性心不全医療の質について病院間比較を行う際、今後も広く利用できると考えられる。(1594)

【キーワード】アウトカム指標、リスク調整死亡率、急性心不全

*Possible categories* (2 つ) : 01)医療政策 02)地域医療 03)医療経済 04)看護管理 05)医療安全 06)質管理・ガイドライン 07)医療情報管理 08)薬剤・資材管理 09)病院経営 10)人材育成・管理 11)在宅介護・ケアマネジメント 12)診療報酬 13)患者参加・患者支援 14)国際貢献 15)その他

# 急性心不全患者の院内死亡予測モデル 開発とリスク調整死亡率の病院間比較

京都大学大学院 医学研究科 社会健康医学系専攻  
医療経済学分野

佐々木 典子、國澤 進、大坪 徹也、猪飼 宏、今中 雄一

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## 背景

- 急性心不全は死亡率・再入院率とも高く、高齢化による患者数増加は医療費圧迫の大きな一因となることが世界的に懸念されている。(Dickstein K et al, Eur J Heart Failure 2008;10:933)
- 多彩な病態を呈する急性心不全患者をこれまでDPCデータで同定することは困難で、病院間の医療提供内容やアウトカム指標を比較検討することは不可能だった。
- 急性心不全の国内疫学データは乏しく、急性心不全の臨床的な多施設症例登録 The acute decompensated heart failure syndromes (ATTEND) レジストリー結果が参考になる。しかし、この研究は大規模施設を中心とした29病院の限られたデータである(2007年4月～2009年5月の中間報告) ことを考慮する必要がある。(Sato N et al, Am Heart J 2010;159:949)

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# 目的

- 新規に導入された病名付加コードから急性心不全患者を同定し、ATTENDレジストリーと比較して対象集団の疫学的位置づけを確認する。
- 院内死亡をアウトカムとして、病院の医療提供内容の妥当な比較を行うため、重症度など患者因子でリスク調整を行うモデルを作成し、そのモデルの妥当性を検討する。
- 同モデルを用いて、観察死亡率とリスク調整予測死亡率につき病院間比較を行う。

## 方法 (1)

- 研究デザイン：観察横断研究
- 本研究で使用したデータは、当教室で主催する Quality Indicator/Improvement Project (QIP) のデータベースである。QIPは全国の急性期病院から自発的に提供されたDPCデータの分析・比較を通して、医療の質・効率性・経済性について検討し、解析結果を参加病院にフィードバックするプロジェクトである。  
施設の種類は、教育病院/非教育病院、国公立病院/公的病院/民間病院からなる。

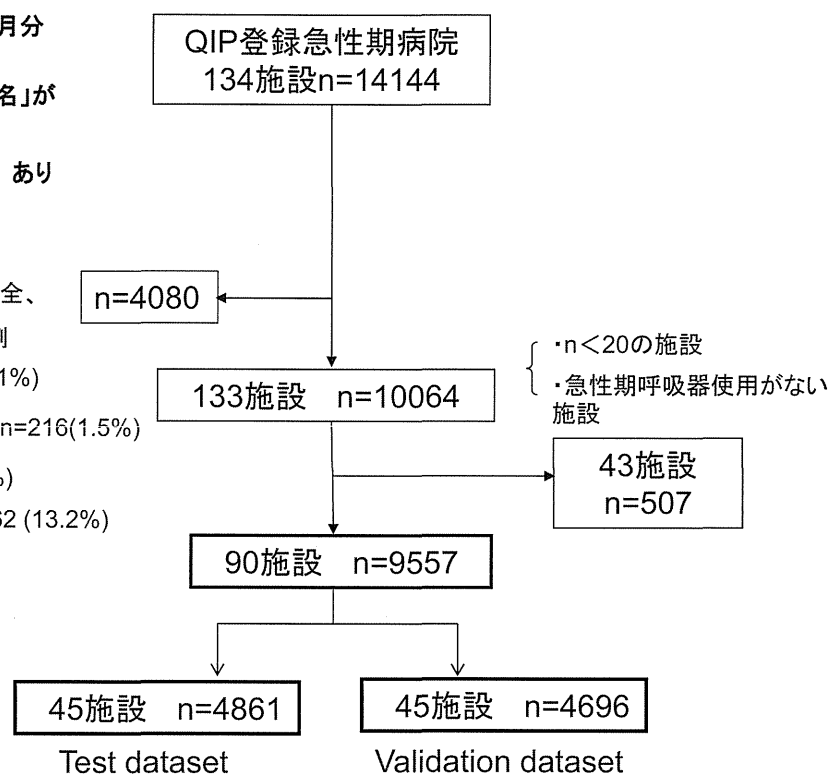
## 方法(2)～Selection process of patients～

### <選択基準>

- ・2010年4月～2011年3月(12ヶ月分データ提出病院)
- ・「医療資源を最も投入した傷病名」が「心不全」DPCコード 050130
- ・急性増悪コード 30101 / 30102 あり
- ・年齢≥20才

### <除外基準>

- ・来院時に心肺停止、末期腎不全、ARDS、胸膜炎、肺炎併存症例 n=2421(17.1%)
- ・在院日数が極端に長い症例 n=216(1.5%)
- ・データ不整合症例 n=1 (0.01%)
- ・入院時NYHA欠損症例 n=1862 (13.2%)



## 方法 (3)

### ■ モデル作成方法

- ・対象90施設を病床規模で対応させてランダムに2群に分け(Split Sample法)、一方をtest dataset (n=4861), もう一方をvalidation dataset (n=4696)とした。
- ・年齢、性別のほか、入院時に重症度を反映すると考えられる変数(次表)を用いてロジスティック回帰分析(ステップワイズ法)を行い、院内死亡予測モデルを作成した。予測精度としてC-statisticsを算出した。

### ■ モデル妥当性検証方法

test datasetで作成した院内死亡予測モデルの回帰係数を用いてvalidation datasetで予測死亡率を計算した。Validation datasetにおけるC-statisticsを求め、その後モデル適合度判定のためHosmer-Lemeshow  $\chi^2$ 検定 ( $p>0.05$ で適合度良好)を行った。

### ■ 病院間比較

全病院の粗死亡率(観察死亡率)とリスク調整予測死亡率(95%信頼区間)を算出し、病院間比較を行った。