

provided is not detrimentally affected, and further efforts must be made to improve the quality of care in regions with lower health care costs in Japan. This study helps us understand the relationship between spending and quality of health care in Japan.

References

1. Fisher ES, Wennberg DE, Stukel TA, et al. The implications of regional variations in Medicare spending. Part 1: The content, quality and accessibility of care. *Ann Intern Med* 2003;138:273–87.
2. Fisher ES, Wennberg DE, Stukel TA, et al. The implications of regional variations in Medicare spending. Part 1: Health outcomes and satisfaction with care. *Ann Intern Med* 2003;138:288–98.
3. Rudd A, Irwin P, Rutledge Z, et al. Regional variations in stroke care in England, Wales and Northern Ireland: results from the National Sentinel Audit of Stroke. *Clin Rehabil* 2001;15:562–72
4. Fisher ES, Wennberg JE. Health care quality, geographic variations and the challenge of supply-sensitive care. *Perspect Biol Med* 2003;46:69–79.
5. Otsubo T, Imanaka Y, Lee J, et al. Evaluation of resource allocation and supply-demand balance in clinical practice with high-cost technologies. *J Eval Clin Prac* 2011;17:1114–21.
6. Tanihara S, Kobayashi Y, Une H, et al. Urbanization and physician maldistribution: a longitudinal study in Japan. *BMC Health Serv Res* 2001;11:260
7. Ministry of Health, Labour and Welfare, Japan [website]. Analysis by cause of death. Abridged Life Tables For Japan 2010. <http://www.mhlw.go.jp/english/database/db-hw/lifetb10/4.html>

1 **Title:** Derivation and Validation of In-Hospital Mortality Prediction Models in Ischaemic Stroke

2 Patients Using Administrative Data

3

4 **Author Names:**

5 Jason LEE

6 Toshitaka MORISHIMA

7 Susumu KUNISAWA

8 Noriko SASAKI

9 Tetsuya OTSUBO

10 Hiroshi IKAI

11 Yuichi IMANAKA*

12 **Affiliation:** Department of Healthcare Economics and Quality Management, Graduate School

13 of Medicine, Kyoto University

14 **Address:** Yoshida Konoe-cho, Sakyo-ku, Kyoto 606-8501, Japan

15

16 **Running Head:** Stroke Mortality Prediction using Administrative Data

17

18

1 ***Corresponding Author**

2 **Address:** As above

3 **Fax:** +81-75-753-4455

4 **Phone:** +81-75-753-4454

5 **Email:** imanaka-y@umin.net

6

7 **Keywords:** acute ischemic stroke, risk factor, stroke, outcome predictive factors, Japan,

8 mortality, administrative data

9 **Word count:** 3004 words

10

11

12

13

14

15

16

17

18

1 **Abstract**

2 **Background:** Stroke and other cerebrovascular diseases are a major cause of death and
3 disability. Predicting in-hospital mortality in ischaemic stroke patients can help to identify
4 high-risk patients and guide treatment approaches. Chart reviews provide important clinical
5 information for mortality prediction, but are laborious and limiting in sample sizes.
6 Administrative data allow for large-scale multi-institutional analyses but lack the necessary
7 clinical information for outcomes research. However, administrative claims data in Japan has
8 seen the recent inclusion of patient consciousness and disability information, which may allow
9 more accurate mortality prediction using administrative data alone. The aim of this study was to
10 derive and validate models to predict in-hospital mortality in patients admitted for ischaemic
11 stroke using administrative data.

12 **Methods:** The sample consisted of 21 445 patients from 176 Japanese hospitals, who were
13 randomly divided into derivation and validation subgroups. Multivariable logistic regression
14 models were developed using 7-day, 30-day and overall in-hospital mortality as dependent
15 variables. Independent variables included patient age, sex, comorbidities upon admission, Japan
16 Coma Scale score, Barthel index score, modified Rankin Scale score, and admissions after
17 hours and weekend/public holidays. Models were developed in the derivation subgroup, and
18 coefficients from these models were applied to the validation subgroup. Predictive ability was

1 analysed using C-statistics; calibration was evaluated with Hosmer-Lemeshow chi-squared tests.

2 **Results:** All three models showed predictive abilities similar or surpassing that of chart
3 review-based models. The C-statistics were highest in the 7-day in-hospital mortality prediction
4 model, at 0.906 and 0.901 in the derivation and validation subgroups, respectively. For the
5 30-day in-hospital mortality prediction models, the C-statistics for the derivation and validation
6 subgroups were 0.893 and 0.872, respectively; in overall in-hospital mortality prediction, these
7 values were 0.883 and 0.876.

8 **Conclusions:** In this study, we have derived and validated in-hospital mortality prediction
9 models for three different time spans using a large population of ischaemic stroke patients in a
10 multi-institutional analysis. The recent inclusion of Japan Coma Scale, Barthel index, modified
11 Rankin Scale scores in Japanese administrative data has allowed the prediction of in-hospital
12 mortality with accuracy comparable to that of chart review analyses. The models developed
13 using administrative data had consistently high predictive abilities for all models in both the
14 derivation and validation subgroups. These results have implications in the role of
15 administrative data in future mortality prediction analyses.

16

17

18

1 INTRODUCTION

2 Stroke and other cerebrovascular diseases are a major cause of death and disability [1,2], and
3 represent the third leading cause of death in Japan [3]. Despite a decrease in stroke incidence in
4 the past few decades, factors such as the ageing population of Japan and the saturation of
5 antihypertensive therapy may have contributed to the disease burden of stroke [4].

6 Within a backdrop of rising health care costs, hospitals are under pressure by
7 governments and third-party payers to provide high standards of care with limited resources.
8 These standards are frequently evaluated with quality indicators, which can be categorized
9 according to the Donabedian structure, process and outcome facets of health care [5]. In
10 addition to determining quality of care, these evaluations provide benchmarking opportunities
11 that can be used in incentive systems or payment system reform [6].

12 A common outcome measure in hospital evaluations is in-hospital mortality; although
13 not all deaths can be prevented, appropriate process of care should reduce unnecessary and
14 preventable deaths, particularly in the early stages of admission [7]. Various timespans of
15 in-hospital mortality are used as indicators, ranging from the short-term to the entire admission
16 period [8–12]. But mortality rates are not easily interpretable, as failure to properly account for
17 variations in case mix and patient severity would bias evaluations against hospitals that treat a
18 more severe class of patients.

1 Predicting in-hospital mortality upon admission may help to provide accurate
2 prognoses to patients and relatives, guide therapeutic goals, and identify high-risk patients [13–
3 15]. Mortality prediction models that use data from chart reviews tend to be more accurate than
4 those using administrative data, as chart reviews provide critical predictors of mortality such as
5 patient level of consciousness and disability levels [16-20]. On the other hand, hospital
6 performance evaluations become more meaningful when interpreted in the context of
7 multi-institutional data, and analysing regional variations can also contribute to understanding
8 the mortality rates of a single hospital. Large-scale data from numerous hospitals are required
9 for such studies, but the laborious nature of chart review analyses can limit sample sizes. Also,
10 the importance of standardized data for proper data collection and determining predictive factors
11 had been shown in previous stroke registries [21,22].

12 The advent of computerized hospital claims data has brought about more opportunities
13 for researchers and policymakers to analyse health care quality, trends, and outcomes on a scale
14 larger than previously possible. Administrative data has become extensively used in health
15 services research, particularly in process measures of care [23–25]. However, these data
16 generally lack clinical information reflecting patient severity upon admission and are therefore
17 limited in their practical applications for predicting mortality. Previous mortality prediction
18 models that have used administrative data alone have done so with a level of discrimination

1 lower than that of clinical data [26,27].

2 Recent additions to Japanese administrative claims data include information on patient
3 consciousness and disability levels, which may improve the predictive ability of mortality
4 prediction models based on these data. Additionally, information on the date of stroke
5 occurrence has also become available. This allows researchers to distinguish acute stroke from
6 chronic or subacute stroke, previously a major limitation of administrative data-based stroke
7 analyses.

9 **Aims**

10 The main aim of this study was to utilize the recent additions of patient consciousness and
11 disability levels to administrative claims data in order to derive and validate models to predict
12 7-day, 30-day and overall in-hospital mortality in patients admitted for acute ischaemic stroke.

14 **METHODS**

15 **Data Source and Subjects**

16 Administrative claims data were obtained from hospitals voluntarily enrolled in the Quality
17 Indicator/Improvement Project (QIP), which utilizes large-scale data analysis and feedback for
18 the monitoring, evaluation, and improvement of health care quality. Hospitals in the QIP

1 provide data in a standardized format under the Diagnosis Procedure Combination (DPC)
2 case-mix payment system. These hospitals represent a variety of hospital ownership, teaching
3 statuses, size and patient casemix.

4 We identified the discharge records of patients who had been admitted into the sample
5 hospitals with acute ischaemic stroke occurring on the date of admission as the main diagnosis
6 or the condition for which the most resources had been used in that admission. We selected only
7 patients who had stroke onset on the day of admission because patients with stroke occurring
8 one or more days earlier would introduce an added layer of uncertainty in the factors affecting
9 mortality, and patients with stroke occurring after admission would imply admission for other
10 reasons. Patients admitted between July 1st, 2010 and June 30th, 2011 were included in analysis.
11 Ischaemic stroke was identified using the International Classification of Diseases, 10th Revision
12 (ICD-10) code for cerebral infarction (I63x). Patients were excluded from analysis if they were
13 below 18 years of age upon admission. Furthermore, patients with extremely short length of
14 stay durations (transferred or discharged alive within 2 days of admission) presented difficulties
15 in post-analysis interpretation of the results. Similarly, patients with extremely long length of
16 stay durations may not reflect mortality in acute ischaemic stroke. As such, both these groups of
17 patients were also excluded from analysis. Finally, patients who had been admitted to hospitals
18 with fewer than 20 cases per year were also excluded.

1

2 **Mortality Prediction**

3 Study outcomes were defined as in-hospital death for the patients within 7 days from admission,
4 30 days from admission, and any time within the entire hospitalization period. In this study, the
5 terms “7-day mortality” and “30-day mortality” are used to refer to in-hospital mortality that
6 occurs within the stated timespans, and “in-hospital mortality” is used to refer to mortality that
7 occurs at any point during the entire admission period. Multivariable logistic regression
8 analyses were conducted using the various outcome measures as dependent variables.
9 Independent variables included patient age (as a continuous variable), sex, comorbidities upon
10 admission, patient consciousness and disability levels, and arrival period. Comorbidities
11 included acute myocardial infarction, atrial fibrillation, dyslipidemia, hypertension, peripheral
12 vascular disease, chronic pulmonary disease, connective tissue disease, liver disease, renal
13 disease, and cancer. These comorbidities were identified according to the criteria as stipulated in
14 the Dartmouth-Manitoba version of Charlson comorbidity index [28]. Other comorbidities such
15 as diabetes and paraplegia were also considered, but ultimately excluded from the final model
16 due to low incidence. This may be related to a possible undercoding issue that limits the number
17 of simultaneous diagnoses in the restrictive formats of administrative data.

18 Patient consciousness and disability/dependence levels included Japan Coma Scale

1 (JCS) scores, Barthel activities of daily living (ADL) index scores, and modified Rankin Scale
2 (mRS) scores on admission; these variables have been included in DPC data from July 2010
3 onward. These scores are measured according to the respective standardized criteria of each of
4 the 3 variables by clinicians and transcribed to the claims data by administrative staff. The JCS
5 is the most widely used clinical tool for evaluating consciousness level in Japanese emergency
6 care, and consists of a scale of ten levels categorised into four groups: (i) JCS level 0, indicating
7 a completely alert state; (ii) JCS levels 1–3 (disoriented: awake without stimulation); (iii) JCS
8 levels 10–30 (somnolent: arousable only in the presence of stimulation); and (iv) JCS levels
9 100–300 (comatose: unarousable despite stimulation) [29]. For this analysis, we utilized JCS
10 level 0 as the referent, with the remaining class groups as binary variables. Next, during the
11 model construction process, we decided to use a cut-off point for Barthel ADL index score of 20
12 (out of a possible 100); an index score that was less than 20 was included as the independent
13 variable, and scores that were 20 or more were used as the referent. The cut-off of 20 was
14 selected due to as patients with scores below this value represent low functional ability, and
15 accounted for approximately half of our sample. In the case of mRS, scores of 0 to 3 were used
16 as the referent, as patients with these scores generally require little assistance in their daily
17 activities; mRS scores of 4 and 5 were included as dummy variables to indicate patients with
18 severe disabilities.

1 Factors concerning the arrival period included whether a patient was admitted after
2 hours (defined as the period between 6pm and 8am), and on weekends and public holidays.

3 The predictive abilities of the models were assessed using the C-statistic.

4 The patients were randomly assigned into derivation and validation subgroups, with
5 each group comprising approximately 50% of the overall sample. Prediction models for the
6 three mortality outcomes were constructed using the derivation dataset and evaluated using the
7 validation dataset. We conducted this validation by applying the regression coefficients from
8 each independent variable obtained from the test subgroup to the validation subgroup, and
9 calculated the C-statistic from the resulting analysis. Calibration of each model was evaluated
10 using the Hosmer-Lemeshow chi-squared test (with $P > 0.05$ considered favourable). All
11 statistical analyses were conducted using SPSS software, version 17.0.0 (SPSS
12 Inc., Chicago, IL, USA). A two-tailed P -value below 0.05 indicated statistical significance.

13

14 **RESULTS**

15 After exclusions, overall sample size for the analysis was 21 445 patients from 176 hospitals.

16 The derivation subgroup consisted of 10 774 patients, and the validation subgroup consisted of
17 10 671 patients. Table 1 shows the demographics of the overall sample, and the derivation and
18 validation subgroups. The derivation and validation subgroups were statistically similar to each

1 other insofar as the independent variables were concerned: a two-sided *t*-test for age (given as a
2 continuous variable in years) as well as chi-squared tests for the other variables (given as binary
3 variables) showed non-significant *P*-values between the two subgroups for all the variables used,
4 as well as in all three outcomes.

5 Patients were approximately 75 years of age upon admission, and there were more
6 men than women represented in the sample. There were higher incidences of atrial fibrillation,
7 dyslipidemia and hypertension when compared to the other comorbidities. Approximately 50%
8 of the patients were alert according to the JCS levels, with decreasing incidence in the higher
9 JCS levels (indicating less alert states). On the other hand, there was increasing incidence in the
10 higher mRS scores, and more than half of the patients had low functional activity (Barthel index
11 <20).

12 The results of the prediction models developed in the derivation subgroup are shown
13 in Table 2. All three models showed good calibration, with *P*-values from the
14 Hosmer-Lemeshow tests for the 7-day mortality model, 30-day mortality model, and in-hospital
15 mortality model being 0.716, 0.280, and 0.119, respectively. Although age did not have
16 significant association with short-term mortality indicator of 7-day mortality, this factor gained
17 significance when the time-line for mortality was extended to 30 days and to overall in-hospital
18 mortality. Atrial fibrillation was significantly associated with increased mortality for all three

1 timespans. Peripheral vascular disease was significantly associated with increased in-hospital
2 mortality, and marginally non-significant in 30-day mortality (*P*-value: 0.051). Although renal
3 disease and cancer had no significant association with 7-day mortality, they were associated
4 with increased 30-day mortality and in-hospital mortality. On the other hand, dyslipidemia and
5 hypertension were both significantly associated with reduced mortality for all three timespans.

6 Reduced patient consciousness was significantly associated with increased mortality:
7 the JCS levels showed that even mild disorientation (levels 1 to 3) was associated with
8 increased mortality. However, the adjusted ORs for JCS levels 10 to 30 and 100 to 300 were
9 highest in 7-day mortality, but decreased with longer timespans. Disability and dependence
10 levels were also significantly associated with increased mortality, with Barthel index scores of
11 less than 20 and an mRS score of 5 significantly associated with all three mortality indicators.
12 The mRS score of 4 was significantly associated in 30-day mortality and in-hospital mortality,
13 but not in 7-day mortality.

14 The Figure shows the predictive ability of the models for the three mortality outcomes
15 in the derivation subgroup and the validation subgroup. The C-statistics were highest in 7-day
16 mortality, and were all higher than 0.872 for the various measures.

17

18 **DISCUSSION**

1 The accuracy of predicting in-hospital mortality in ischaemic stroke patients using
2 administrative claims has been limited by the lack of information reflecting disease severity of
3 patients upon admission. However, administrative data in Japan have recently included
4 information on patient consciousness and disability levels. In this study, we have derived and
5 validated in-hospital mortality prediction models for three timespans based on administrative
6 data, and shed light on predictors of in-hospital mortality in patients admitted for ischaemic
7 stroke.

8 Predictive ability was highest in the short-term 7-day mortality indicator, with the
9 C-statistic reaching 0.906 in the derivation subgroup. This predictive ability is similar to or
10 higher than existing mortality prediction models, whether in chart review analyses or
11 administrative data complemented with clinical information [8–11,19]. Although predictive
12 ability was observed to be reduced as the timespans increased, the C-statistics for both 30-day
13 mortality and in-hospital mortality were still noticeably higher than previous attempts to predict
14 mortality using administrative data [26,27]. Validation of all three prediction models showed a
15 high degree of consistency between the discriminatory ability of the models in both the
16 derivation and validation subgroups.

17 By analysing different mortality rates at various timespans in the same population, we
18 were able to observe the shift in predictors that influence mortality. Age was not observed to

1 have a significant association with 7-day mortality, but was associated with mortality in the
2 longer timespans. Furthermore, short-term mortality was found to be most influenced by low
3 levels of patient consciousness, which is similar to the results found using clinical information
4 [17]. Atrial fibrillation was consistently associated with increased mortality across all three
5 timespans, a result corroborated by previous studies [30–32]. Higher mRS and Barthel index
6 scores, in addition to JCS levels, were significantly associated with mortality in general;
7 however, only the highest mRS score of 5 was significantly associated with 7-day mortality.
8 Although we did not have access to a stroke severity scale, such as the National Institutes of
9 Health Stroke Scale (NIHSS), the combination of patient consciousness levels (which is a
10 component of the NIHSS) and patient disability and dependency scales accorded a high level of
11 discrimination to in-hospital mortality prediction.

12 Renal disease and cancer were associated in 30-day mortality and in-hospital mortality,
13 indicating that these diseases had less impact on short-term mortality when compared to patient
14 consciousness and disability levels, which was congruent with results found in a previous study
15 [9,33]. Hypertension and dyslipidemia were consistently associated with lowered mortality for
16 all three timespans. Although hypertension is associated with an increased risk of stroke [34–36],
17 the link between hypertension and stroke mortality may be similar to the results seen in heart
18 failure, in which higher blood pressure has been shown to be associated with reduced risk of

1 dying [37–38]. Furthermore, another study has shown that hypertension was not an important
2 predictor of death in stroke patients [39]. In the case of dyslipidemia, these results may be an
3 indication of reverse epidemiology, which has been previously observed in stroke patients [40–
4 41].

5 Whether patients had been admitted during office hours or after hours had no bearing
6 on survivability, but admissions on weekends and public holidays were associated with
7 increased 7-day mortality. This may indicate that a decreased availability of resources—possibly
8 that of manpower—during weekends and public holidays may have resulted in poorer processes
9 of care leading to increased short-term mortality. However, if a patient should survive past the
10 first seven days, weekend and public holiday admissions ceased to be a significant predictor of
11 in-hospital mortality.

12 Although accurately predicting mortality and its predictors in ischaemic stroke patients
13 is unlikely to substantially improve the survivability of patients, it can strengthen the
14 attributional validity of hospital level evaluations. Fairer evaluations of outcome measures such
15 as mortality would then have more meaningful applications in incentive systems or contributing
16 to payment system reform.

17 In addition to random categorisation at the patient level, we also conducted an
18 additional analysis with randomisation conducted at the hospital level (data not shown). Here,

1 hospitals were randomly divided into 2 groups and tested with the same mortality prediction
2 models. Each of the random groups comprised of 88 hospitals, with 11,314 patients in 1 group
3 and 10,131 patients in the other. The model produced similar results to patient-level randomized
4 groups: in 7-day in-hospital mortality, the C-statistics were 0.894 and 0.911; in 30-day
5 in-hospital mortality, the C-statistics were 0.885 and 0.879; and in overall in-hospital mortality,
6 the C-statistics were 0.874 and 0.871. This showed that the models had comparable predictive
7 abilities in different groups of hospitals.

8 The limitations of this study are as follows: 7-day mortality and 30-day mortality were
9 both indicators of in-hospital mortality on stipulated timespans in this study. We were unable to
10 track mortality that occurred post-discharge, but the acute nature of stroke would generally
11 preclude the transfer or discharge of patients who had not been stabilized; we do not expect
12 this limitation to have a substantial impact on our analysis. Next, the hospitals included in this
13 analysis were all volunteer participants in the QIP, and therefore represent a group of hospitals
14 that have an active aim to improve the quality of the health care that they provide. However, the
15 wide variations in size, ownership and geographic location of the hospitals, as well as the large
16 sample size may make a substantial bias unlikely.

17 Administrative databases are a product of hospital reimbursement systems, and not
18 designed for the purposes of health outcomes research. As such, disease severity information

1 has been relatively less important from a hospital's perspective when compared to information
2 on health services and goods consumed. The novelty of this study shows that the recent addition
3 of the JCS, Barthel ADL Index, and mRS scores to ischaemic stroke patients to the DPC
4 database in Japan has allowed the use of this administrative database to be applied effectively to
5 predict in-hospital mortality Our results have implications in the use of administrative data in
6 future mortality prediction analyses, as well as to provide fairer risk-adjusted mortality rate
7 evaluations. As the quality of administrative claims data improves, health care researchers must
8 endeavour to make the most of this resource.

9
10 **Funding:** This study was supported in part by a Health Sciences Research Grant from the
11 Ministry of Health, Labour and Welfare of Japan, and a Grant-in-Aid for Scientific Research
12 from the Japan Society for the Promotion of Science

13
14
15
16
17
18

1 **REFERENCES**

- 2 1. Murray CJ, Lopez AD: Mortality by cause for eight regions of the world: Global
3 Burden of Disease Study. *Lancet* 1997;349:1269–76.
- 4 2. Kim AS, Johnston SC: Global variation in the relative burden of stroke and
5 ischemic heart disease. *Circulation* 2011;124:314–23.
- 6 3. Ministry of Health, Labour and Welfare, Japan [website]. Analysis by cause of
7 death. Abridged Life Tables For Japan 2010 [Cited 13 September 2012] Available
8 from: <http://www.mhlw.go.jp/english/database/db-hw/lifetb10/4.html>
- 9 4. Shinohara Y: The changing face of the burden of stroke in Japan. *Int J Stroke*
10 2007;2:133–5.
- 11 5. Donabedian A: The quality of care. How can it be assessed? *JAMA*
12 1988;260:1743–8.
- 13 6. Murray CJ, Frenk J: A framework for assessing the performance of health systems.
14 *Bull World Health Organ.* 2000;78:717–31.
- 15 7. AHRQ quality indicators—guide to inpatient quality indicators: quality of care in
16 hospitals—volume, mortality and utilization. *AHRQ Quality Indicators.* Rockville,
17 MD: AHRQ; 2004:1–95.