

A Case Control Association Study and Cognitive Function Analysis of Neuropilin and Tolloid-Like 1 Gene and Schizophrenia in the Japanese Population

Masahiro Banno^{1,9}, Takayoshi Koide^{1,9}, Branko Aleksic^{1,2*}, Kazuo Yamada³, Tsutomu Kikuchi^{1,4}, Kunihiro Kohmura¹, Yasunori Adachi¹, Naoko Kawano¹, Itaru Kushima¹, Masashi Ikeda², Toshiya Inada⁵, Takeo Yoshikawa³, Nakao Iwata², Norio Ozaki¹

1 Department of Psychiatry, Nagoya University Graduate School of Medicine, Showa-ku, Nagoya, Aichi-ken, Japan, 2 Department of Psychiatry, Fujita Health University School of Medicine, Kutsukake, Toyoake, Aichi-ken, Japan, 3 Laboratory for Molecular Psychiatry, RIKEN Brain Science Institute, Wako-city, Saitama, Japan, 4 Matsuzaki Hospital, Miyakojima-ku, Osaka, Japan, 5 Department of Psychiatry, Seiwa Hospital, Institute of Neuropsychiatry, Shinjuku-ku, Tokyo, Japan

Abstract

Background: Using a knock-out mouse model, it was shown that NETO1 is a critical component of the NMDAR complex, and that loss of *Neto1* leads to impaired hippocampal long term potentiation and hippocampal-dependent learning and memory. Moreover, hemizyosity of *NETO1* was shown to be associated with autistic-like behavior in humans.

Purpose of the Research: We examined the association between schizophrenia and the neuropilin and tolloid-like 1 gene (*NETO1*). First, we selected eight single nucleotide polymorphisms (SNPs) within the *NETO1* locus, based on the Japanese schizophrenia genome wide association study (JGWAS) results and previously conducted association studies. These SNPs were genotyped in the replication sample comprised of 963 schizophrenic patients and 919 healthy controls. We also examined the effect of associated SNPs on scores in the Continuous Performance Test and the Wisconsin Card Sorting Test Keio version (schizophrenic patients 107, healthy controls 104).

Results: There were no significant allele-wise and haplotype-wise associations in the replication analysis after Bonferroni correction. However, in meta-analysis (JGWAS and replication dataset) three association signals were observed (rs17795324: $p=0.028$, rs8098760: $p=0.017$, rs17086492: $p=0.003$). These SNPs were followed up but we could not detect the allele-specific effect on cognitive performance measured by the Continuous performance test (CPT) and Wisconsin Card Sorting test (WCST).

Major Conclusions: We did not detect evidence for the association of *NETO1* with schizophrenia in the Japanese population. Common variants within the *NETO1* locus may not increase the genetic risk for schizophrenia in the Japanese population. Additionally, common variants investigated in the current study did not affect cognitive performance, as measured by the CPT and WCST.

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* E-mail: branko@med.nagoya-u.ac.jp

⁹ These authors contributed equally to this work.

Introduction

The glutamate hypothesis of schizophrenia (GHS) emerged in the early 1980s as an alternative to the prevailing theory of altered dopamine neurotransmission. The GHS is based on the observation that non-competitive antagonists of the N-methyl-D-aspartate (NMDA) subtype of glutamate receptors, such as phencyclidine (PCP), ketamine and MK-801, induce a psychotic reaction in healthy individuals that resembles schizophrenia (both the positive and negative symptoms). When the same compounds are

administered to patients with schizophrenia, exacerbation of psychotic symptoms can be the outcome [1]. Together, these observations suggest that diminished function of the NMDA receptor (NMDAR) may play a role in the pathoetiology of schizophrenia. Moreover, evidence from morphological, clinical and neuroimaging studies have provided support for the GHS by mapping cognitive impairment, alterations in blood flow and changes in neuronal morphology to particular brain areas, including the frontal and cingulate cortices, both of which are areas with extensive excitatory glutamatergic neurotransmission [2].

The N-methyl-D-aspartate receptor (NMDAR), a major excitatory ligand-gated ion channel in the central nervous system, is composed of a heterotetramer between two NR1 and two NR2 subunits. Moreover, the NMDAR is a principal mediator of synaptic plasticity [3]. It has been shown that corticolimbic NMDAR hypofunction is one of the core molecular mechanisms relevant for phenotypes observed in animal models of schizophrenia [4]. One of the genes that regulate NMDAR function is neuropilin and toll-like 1 gene (*NETO1*). *NETO1* maps to the 18q22-q23 and three alternative splicing variants (mRNA level) have been observed [5]. Specifically, variants 1 and 2 are detected in the retina while variant 3 is specific for fetal and adult brain. NETO1 is a transmembrane protein, which has two extracellular CUB domains, a low-density lipoprotein class A (LDL_A) domain, a transmembrane domain and classical type I PDZ-domain binding motif [5] (Figure 1). Deletion of *Neto1* leads to deficits in synaptic plasticity in mice while stimulation of the AMPA receptor can partially compensate for deficits caused by *Neto1* deletion [6]. NETO1 interacts with the core NMDAR subunits, NR2A and NR2B and a scaffolding protein, postsynaptic density-95 (PSD-95), maintaining the abundance of NR2A-containing NMDARs in the postsynaptic density of the hippocampus. PSD-95 is a protein that is almost exclusively located in the postsynaptic density of neurons, and is important in anchoring synaptic proteins [7]. Increase in surface NR2A, but not NR2B, occurs in hippocampal neurons derived from dysbindin-null mutant mice (*Dys-/-*). Dysbindin controls hippocampal LTP by selective regulation of the surface expression of NR2A [8]. In situ hybridization studies of schizophrenia detected decreased transcript expression of the NR1 subunit, increased transcript expression of the NR2B subunit and unchanged transcript expression of the NR2A subunit in hippocampus [9]. Therefore, regulation of NR2 in hippocampus in schizophrenia may be relevant for the etiology of schizophrenia and NETO1 may play an important role in the molecular mechanism by maintaining the abundance of NR2A-containing NMDARs in the postsynaptic density of hippocampal neurons. Moreover, NETO1 interacts with kainate receptors (KAR), one of the glutamate receptors, in mouse brain. NETO1 modulates the KAR affinity for the endogenous ligand glutamate. NETO1 modulates not only kinetics, but also the amplitude of slow excitatory postsynaptic current in KAR (KAR-EPSC) [10,11]. NETO1 fundamentally alters the function and neuronal localization of GluK1-containing KAR [10,11]. Therefore, NETO1 may influence glutamate neurotransmission through modulation of KAR and NMDAR properties.

Genetic studies suggested that the *NETO1* locus could harbor genetic variants that are relevant for susceptibility to neuropsychiatric disorders. Specifically, hemizyosity of *NETO1* was shown to be associated with autistic-like behaviors in humans [12]. Although unequivocal genome wide evidence ($p < 10^{-8}$) for association at the *NETO1* locus has not been reported, it is of note that in the recent GWASs of Alzheimer's disease and schizophrenia, suggestive association signals were observed (rs1109070; $p = 0.000669$ [13] and rs9962470; $p = 0.000154$ [14] in Alzheimer's disease and schizophrenia, respectively). Moreover, several weak association signals ($P < 0.05$) within the *NETO1* locus were detected in the first GWAS of schizophrenia conducted in the Japanese population (JGWAS) [15]. It is of note that in the JGWAS, genome wide evidence for association was not detected, and the non-genome wide level of statistical significance should be interpreted with caution. However due to the relatively small sample size, type II errors (false negative result) cannot be excluded – especially in the case of small odds ratios (OR) which are expected for common SNPs associated with schizophrenia [16].

Based on the aforementioned biological studies, *NETO1* can be seen as a promising candidate gene for schizophrenia. However, to the best of our knowledge, no genetic association study specifically designed to evaluate the association between *NETO1* and schizophrenia has been conducted. The goal of the present study was to evaluate the association between *NETO1* and schizophrenia based on the JGWAS. Additionally, as deficits were found in LTP and learning and memory in *Neto1*-null mice [7], we performed a cognitive function analysis that targeted the relationship between common SNPs selected based on the JGWAS and cognitive function assessed by the CPT and the WCST.

Results

In the replication sample set, we did not detect any association between eight SNPs and schizophrenia after Bonferroni correction was applied (Table 1). Using the replication sample set, we have conducted haplotype-wise analysis in order to investigate association between haplotypes within the *NETO1* locus and schizophrenia. We did not detect evidence for haplotype-wise association after Bonferroni correction was applied (Table S1). In order to maximize the power, we performed meta-analysis combining results from the JGWAS and the replication dataset. In this analysis we could detect three suggestive association signals (rs17795324: $P = 0.028$, rs8098760: $P = 0.017$, rs17086492: $P = 0.003$). In the test of heterogeneity, we detected four SNPs, which showed significant high heterogeneities (Breslow-Day test; rs17795324: $P = 0.04$, rs6566674: $P = 0.01$, rs8098624: $P = 0.003$, rs1109070: $P = 0.0001$) (Table 2). However, we could not detect any association between these three SNPs (rs17795324, rs8098760 and rs17086492) and CPT and WCST scores (Table S4). We investigated the association between the *NETO1* gene and schizophrenia stratified by gender. We could detect no association signals either for males or females in the meta-analysis after Bonferroni correction (Tables S2 and S3).

Discussion

In this study, we investigated the association between eight SNPs within *NETO1* and schizophrenia in the Japanese population. We detected associations between *NETO1* and schizophrenia in the meta-analysis, however, as the JGWAS dataset was included in the meta-analysis, evidence for association might be overestimated. In order to address this issue we tested the association between candidate SNPs from our meta-analysis and cognitive performance measured by the CPT and WCST. This analysis was of interest for us as *Neto1*-null mice show deficits in LTP and learning and memory [7], and if the three SNPs (rs17795324, rs8098760, rs17086492) were genuinely associated with schizophrenia, then carriers of risk alleles would likely have deficits in cognitive processing assessed by CPT and WCST [6]. However, we could not detect any association between these three SNPs (rs17795324, rs8098760, rs17086492) and the psychological tests we applied. We also performed meta-analysis (Method S1) of rs6566674 (the SNP that was included both in our replication sample set and another Japanese GWAS [17]), however we did not detect evidence for an association with schizophrenia. These results suggest that *NETO1* is not associated with schizophrenia in the Japanese population.

We detected four SNPs, which showed heterogeneous association (Breslow-Day test; rs17795324: $P = 0.04$, rs6566674: $P = 0.01$, rs8098624: $P = 0.003$, rs1109070: $P = 0.0001$) in the meta-analysis. These high heterogeneities may be derived from flip-flop phenomenon, that is, associations of opposite alleles at the same biallelic locus with the same disease [18]. Although the flip-

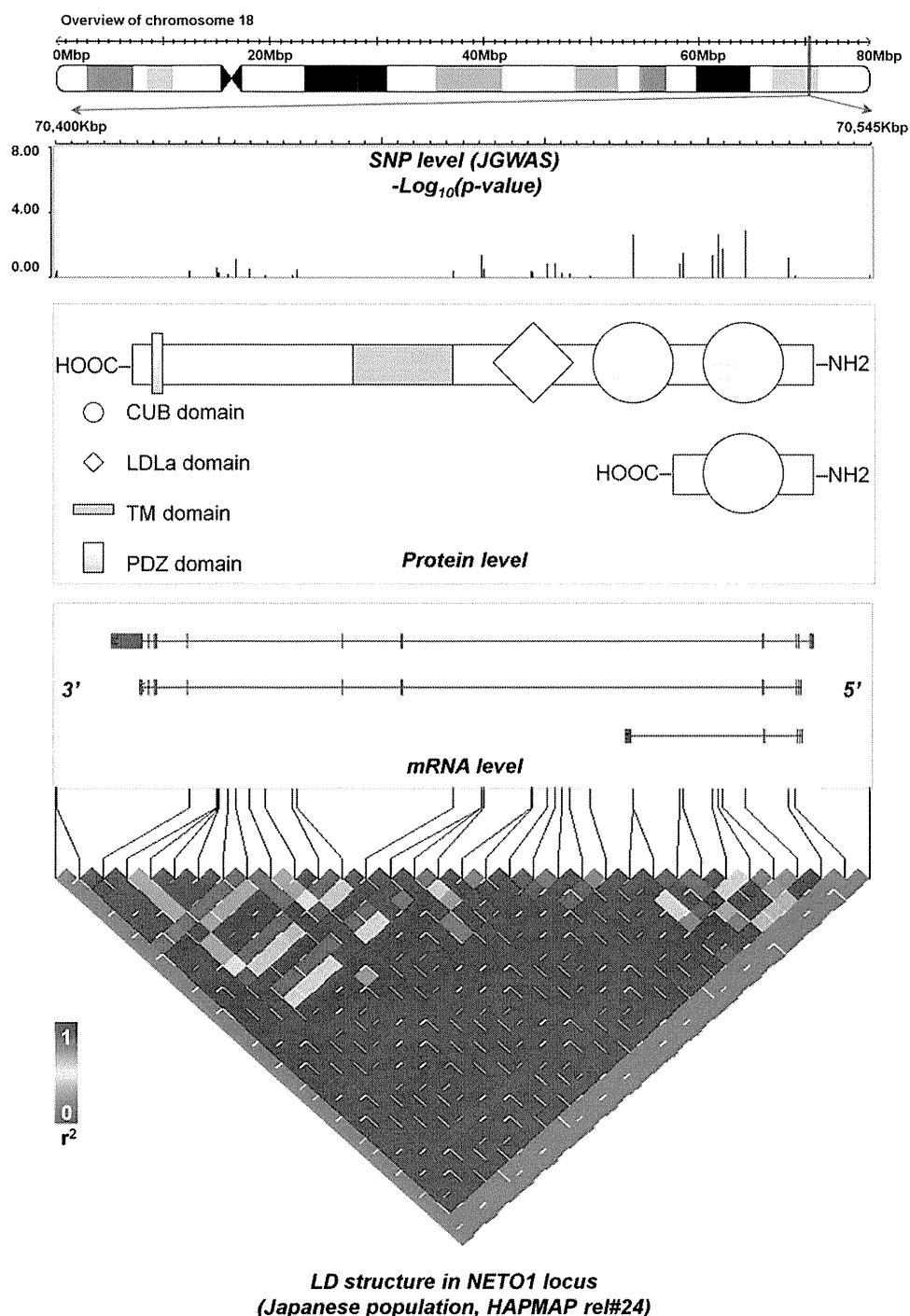


Figure 1. NETO1-gene centric data. LD: linkage disequilibrium. NETO1 maps to the 18q22-q23 and three alternative splicing variants (mRNA level) have been observed. Specifically variants 1 and 2 are detected in retina while variant 3 is specific for fetal and adult brain (mRNA level panel). NETO1 is a transmembrane protein, which has two extracellular CUB domains, a low-density lipoprotein class A (LDLa) domain, a transmembrane domain and a classical type I PDZ-domain binding motif (protein level panel). doi:10.1371/journal.pone.0028929.g001

flop phenomenon may represent a genuine genetic association (i.e., genuinely different LD architectures across populations with different ancestral origins), it may also be an artifact due to sampling variation that leads to variability in observed LD patterns.

Several caveats should be considered when interpreting the results of our study. First, in terms of sample size, the replication

dataset may not have sufficient statistical power to detect associations between SNPs with low genotype relative risk (GRR) and schizophrenia. In other words, our sample has statistical power greater than 0.8 for the detection of association signals at nominal statistical significance, of the polymorphism with a minor allele frequency of 0.1, when the GRR is 1.30. Therefore, the possibility of association between schizophrenia

Table 1. Results of JGWAS (N = 1108).

SNP	Position ^a	Minor allele	Case ^b	Control ^b	p-value ^c	OR	L95 ^d	U95 ^d
rs9962470	68616086	A	0.03	0.04	0.31	0.79	0.50	1.24
rs17086334	68626752	C	0.10	0.08	0.05	1.34	1.00	1.80
rs17795324	68654033	G	0.27	0.33	0.003	0.76	0.63	0.91
rs6566674	68662791	T	0.20	0.16	0.04	1.26	1.01	1.57
rs8098624	68669199	T	0.30	0.24	0.002	1.35	1.12	1.63
rs8098760	68669970	T	0.11	0.08	0.02	1.42	1.06	1.90
rs17086492	68674050	C	0.14	0.10	0.001	1.52	1.17	1.98
rs1109070	68674559	C	0.21	0.16	0.01	1.35	1.09	1.68

^abased on NCBI 36.^bminor allele frequency.^cFisher's exact test.^dLower (L) and upper (U) 95% confidence intervals.

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and common SNPs with $GRR < 1.30$ cannot be excluded. Furthermore, the JGWAS may not have sufficient power to detect associations between SNPs with low GRR and schizophrenia. Therefore, other relevant common variants in the *NETO1* region, which the JGWAS cannot identify, may exist. In addition, association between *NETO1* and schizophrenia may be specific only for the patients with early onset age. However, as in the current study, the frequency of the early onset schizophrenia patients was low (less than 5%). The power to detect such an association is inadequate. Therefore, the effect of onset age on the association between *NETO1* and schizophrenia remains to be investigated in further studies.

The second caveat is that our study design was based on the common disease, common variant hypothesis, based on which we applied a minor allele frequency threshold (>5%) and selected eight SNPs for follow-up. In the best case scenario, common variants detected in GWAS can explain only part of the heritability in cases of schizophrenia (~30%) [19] and missense or nonsense mutations on the one side and structural variations (i.e., copy number variants (CNVs)) on the other side are likely to contribute to the increased susceptibility [20]. Recently, the

concept of synthetic associations has been suggested, though some there are some objections [21]. Uncommon or rare genetic variants can easily create synthetic associations that are credited to common variants. This possibility requires careful consideration in the interpretation and follow up of GWAS signals [22].

The third caveat in our association study is that cases and controls in replication samples were not matched in age. In other words, although highly unlikely, the controls may develop schizophrenia at some point in life, as they were significantly younger than cases.

The fourth caveat is related to the validity of cognitive function analysis. The premise that *NETO1* was associated with cognitive function in humans had been derived from results of a knock-out mice study [7]. *Neto1*-null mice showed impaired spatial learning measured by the Morris water maze task, the delayed matching-to-place version of the Morris water maze task and displaced-object tasks. In the current study, we investigated executive function (WCST) and vigilance/attention (CPT-IP), however, the results of these cognitive tests might not represent similar cognitive dysfunctions that were shown in the *Neto1*-null mouse study. It may be useful to examine different domains of cognitive impairment

Table 2. Allele frequencies of the eight SNPs of *NETO1*.

SNP	Position ^a	Minor allele	Replication (N = 1882)						Meta analysis (N = 2990)				
			Case ^b	Control ^b	p-value ^c	OR	L95 ^d	U95 ^d	p-value ^c	OR	L95 ^d	U95 ^d	BDp ^e
rs9962470	68616086	A	0.02	0.03	0.24	0.79	0.53	1.18	0.126	0.79	0.59	1.07	0.99
rs17086334	68626752	C	0.08	0.08	0.87	0.98	0.77	1.25	0.289	1.11	0.92	1.33	0.09
rs17795324	68654033	G	0.29	0.30	0.61	0.96	0.84	1.11	0.028	0.88	0.79	0.99	0.04
rs6566674	68662791	T	0.17	0.19	0.12	0.88	0.74	1.04	0.926	1.01	0.88	1.15	0.01
rs8098624	68669199	T	0.27	0.29	0.43	0.94	0.82	1.09	0.213	1.08	0.96	1.21	0.003
rs8098760	68669970	T	0.10	0.08	0.18	1.17	0.93	1.47	0.017	1.24	1.04	1.49	0.26
rs17086492	68674050	C	0.12	0.11	0.22	1.14	0.93	1.39	0.003	1.28	1.09	1.50	0.09
rs1109070	68674559	C	0.16	0.19	0.01	0.78	0.65	0.93	0.718	0.98	0.85	1.12	0.0001

^abased on NCBI 36.^bminor allele frequency.^cFisher's exact test.^dLower (L) and upper (U) 95% confidence intervals.^ep-value of Breslow-Day test.

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associated with *NETO1* in schizophrenic patients using broader cognitive assessment tools.

In conclusion, we were not able to detect evidence for an association between *NETO1* and schizophrenia in the Japanese population. Common variants within the *NETO1* locus may not increase the genetic risk for schizophrenia in the Japanese population. Additionally, common variants investigated in the current study did not affect cognitive performance, as measured by the CPT and the WCST.

Materials and Methods

Participants

This study was approved by the Ethics Committees of the Nagoya University Graduate School of Medicine and Fujita Health University, and written informed consent was obtained from each participant. Patients were included in the study if they (1) met DSM-IV criteria for schizophrenia, (2) were physically healthy and (3) had no mood disorders, substance abuse, neurodevelopmental disorders, epilepsy or known mental retardation. A general characterization and psychiatric assessment of subjects is available elsewhere [23]. Controls were selected from the general population. Control subjects had no history of mental disorders, based on questionnaire responses from the subjects themselves during the sample inclusion step, and based on an unstructured diagnostic interview done by an experienced psychiatrist during the blood collection step. The JGWAS sample was comprised of 575 patients with schizophrenia (43.5 ± 14.8 years (mean \pm s.d.), male 50%) and 564 healthy controls with no personal or family history of psychiatric illness (44.0 ± 14.4 years (mean \pm s.d.), male 49.8%). All subjects were unrelated, living in the central area of the Honshu island of Japan and self-identified as members of the Japanese population. Subjects of replication samples consisted of 963 schizophrenic patients (47.7 ± 0.5 years (mean \pm s.d.), male 55.2%) and 919 healthy controls (45.0 ± 0.5 years (mean \pm s.d.), male 51.0%). The JGWAS and replication samples were collected independently at each university hospital.

Genotyping and data analysis

Based on the JGWAS results we initially selected SNPs with probability values, $p < 0.05$ and allelic frequencies, $MAF > 0.05$ within the *NETO1* locus. Then we identified redundant SNPs based on the linkage disequilibrium or LD pattern within the interrogated region. Specifically, if the correlation coefficient between two loci (r^2) was 0.8 or higher, only one of the two loci was selected for the association study [24]. The correlation coefficient between two loci (r^2) was calculated using Haploview v.4.1 based on the HapMap database (release no. 24, population: Japanese in Tokyo). Finally, we selected seven nonredundant SNPs within the *NETO1* locus. Moreover, one common polymorphism (rs9962470), which showed a low p-value ($p = 0.000154$) in a previous GWA study [14], was included for genotyping. All eight SNPs are intronic polymorphisms. DNA was extracted from peripheral blood according to a standard protocol [23]. Genotyping was performed using a fluorescence-based allelic discrimination assay (Taqman, Applied Biosystems, Foster City, CA, USA). Power was calculated according to the methods of Skol et al. [25].

To exclude low-quality DNA samples or genotyping probes, data sets were filtered on the basis of SNP genotype call rate (more than 90%) or checked deviation from Hardy-Weinberg equilibrium (HWE) in the control sample. Subjects whose percentage of missing genotypes was $> 30\%$ or who had evidence of possible DNA contamination were excluded from subsequent analyses.

To reduce the total number of tests, eight associated markers were selected based on the JGWAS results. Next, conditional on the findings of the JGWAS, which used a less stringent nominal level, a meta-analysis was done involving the confirmation sample using the replication data and data from the JGWAS. In the replication sample, Fisher's exact test was used to compare allele frequencies between patients and control subjects. The significance level was set at $p < 0.05$. In the replication sample set, log likelihood ratio tests for assessing haplotype-wise association between schizophrenia and a combination of tagging SNPs was performed using UNPHASED software v3.04. The rare haplotype frequency threshold was set at 5% [26]. In this meta-analysis, p-values were generated by a Cochran-Mantel-Haenszel stratified analysis, and the Breslow-Day test was performed for evaluations of heterogeneous associations as implemented in gPLINK v.2.050 [27].

Neurocognitive assessment

1. CPT. We used the Continuous Performance Test-Identical Pairs Version Release 4.0 (NewCPT.exe, Copyright 1982-2004 by Barbara A. Cornblatt, All Rights Reserved). The size of PC monitor used for the test was 10.4 inches as each letter was at least 2.2×1.5 cm [28]. Stimuli were flashed on the screen at a constant rate of 1 per second, with a stimulus "on" time of 50 ms. Stimuli were four-digit numbers and were presented 150 times. In each 150-trial condition, 30 of the trials (20%) were target trials and required a response. Target trials were those on which the second of a pair of two identical stimuli appeared [28]. The outcome measure was a mean, d' .

2. WCST. The WCST [29] mainly assesses executive function including cognitive flexibility in response to feedback. We used a modified and computerized version of the test: Wisconsin Card Sorting Test (Keio Version) (KWCSST) [30]. The outcome measures were numbers of categories achieved (CA), total errors (TE), and perseverative errors of Milner (PEM) and Nelson types (PEN) in the first trial. We selected outcomes in the WCST, following a prior study, which used KWCSST as a measure of cognitive function [31].

(1) CA: This is the number of categories for which six consecutive correct responses are achieved (eight is the maximum number of categories which can be achieved), and is the sum measure of the level of conceptual shifts in the KWCSST.

(2) PEN: This is the number of incorrect responses in the same category as the immediately preceding incorrect response (maximum of 47 perseverative errors) [32].

(3) PEM: This is the number of incorrect responses in the same category as the immediately preceding correct response after the category changes.

(4) TE: This is the total number of incorrect responses [33].

3. Clinical information. Chlorpromazine (CPZ) equivalent doses were calculated based on the report by Inada [34,35]. The Positive and Negative Symptom Scale (PANSS) was used to evaluate patients [36].

4. Analysis of cognitive performance. From the sample used in the current study, we made a subset of randomly selected participants older than 18 years of age for analysis of cognitive performance. Cognitive data analysis was done for the participants who completed both WCST and CPT-IP. We checked the effect of three SNPs on cognitive performance measured by the Continuous Performance Test and the Wisconsin Card Sorting Test (107 schizophrenic patients, 104 Healthy controls). IBM SPSS statistical software, version 19 was used for all analyses. We compared age, education, CPZ equivalent doses, age at onset, duration of illness, positive scale, negative scale and General

Psychopathology Scale between schizophrenia cases and control subjects using a two-tailed t-test and Welch's t-test. We compared sex between case and control groups using Fisher's exact test. Next, we compared d' in the CPT and CA, PEM, PEN, TE in the WCST between the case and control groups using a two-tailed t-test and Welch's test (Table S4).

Supporting Information

Method S1 Meta-analysis.
(DOC)

Table S1 Haplotype analysis of the eight SNPs of *NETO1*.
(DOC)

Table S2 Allele frequencies of the eight SNPs of *NETO1* in males.
(DOC)

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Resequencing and Association Analysis of the *KALRN* and *EPHBI* Genes And Their Contribution to Schizophrenia Susceptibility

Itaru Kushima^{1,2}, Yukako Nakamura^{1,2}, Branko Aleksić^{1,2}, Masashi Ikeda^{2,3}, Yoshihito Ito^{1,2}, Tomoko Shiino^{1,2}, Tomo Okochi^{2,3}, Yasuhisa Fukuo^{2,3}, Hiroshi Ujike⁴, Michio Suzuki^{2,5}, Toshiya Inada⁶, Ryota Hashimoto^{2,7,8}, Masatoshi Takeda^{7,8}, Kozo Kaibuchi^{2,9}, Nakao Iwata^{*,2,3}, and Norio Ozaki^{1,2}

¹Department of Psychiatry, Graduate School of Medicine, Nagoya University, Nagoya, Japan; ²Core Research for Evolutional Science and Technology, Japan Science and Technology Corporation, Tokyo, Japan; ³Department of Psychiatry, School of Medicine, Fujita Health University, Toyoake, Aichi 470-1192, Japan; ⁴Department of Neuropsychiatry, Graduate School of Medicine, Okayama University, Dentistry and Pharmaceutical Sciences, Okayama, Japan; ⁵Department of Neuropsychiatry, Graduate School of Medicine and Pharmaceutical Sciences, Toyama University, Toyama, Japan; ⁶Seiwa Hospital, Institute of Neuropsychiatry, Tokyo, Japan; ⁷Molecular Research Center for Children's Mental Development, United Graduate School of Child Development, Osaka University, Kanazawa University and Hamamatsu University School of Medicine, Osaka, Japan; ⁸Department of Psychiatry, Graduate School of Medicine, Osaka University, Osaka, Japan; ⁹Department of Cell Pharmacology, Graduate School of Medicine, Nagoya University, Nagoya, Japan

*To whom correspondence should be addressed; tel: +81-562-93-9250; fax: +81-562-93-1831, e-mail: nakao@fujita-hu.ac.jp

Background: Our genome-wide association study of schizophrenia found association signals at the Kalirin gene (*KALRN*) and EPH receptor B1 gene (*EPHBI*) in a Japanese population. The importance of these synaptogenic pathway genes in schizophrenia is gaining independent supports. Although there has been growing interest in rare (<1%) missense mutations as potential contributors to the unexplained heritability of schizophrenia, there are no population-based studies targeting rare (<1%) coding mutations with a larger effect size (eg, OR >1.5) in *KALRN* or *EPHBI*. **Methods and Results:** The present study design consisted of 3 phases. At the discovery phase, we conducted resequencing analyses for all exon regions of *KALRN* and *EPHBI* using a DNA microarray-based method. Seventeen rare (<1%) missense mutations were discovered in the first sample set (320 schizophrenic patients). After the prioritization phase based on frequencies in the second sample set (729 cases and 562 controls), we performed association analyses for each selected mutation using the third sample set (1511 cases and 1517 controls), along with a combined association analysis across all selected mutations. In *KALRN*, we detected a significant association between schizophrenia and P225T (OR = 2.09, corrected $P = .048$, 1 tailed); this was supported in the combined association analysis (OR = 2.07, corrected $P = .006$, 1 tailed). We found no evidence of association of *EPHBI* with schizophrenia. *In silico* analysis indicated the functional relevance of these rare missense mutations. **Conclusion:** We provide evidence that multiple rare (<1%) missense mutations in *KALRN* may be genetic risk factors for schizophrenia.

Key words: synaptogenic pathway/rare missense mutations/GWAS/Japanese population

Introduction

Schizophrenia is a genetically heterogeneous disorder with heritability estimated at up to 80%.¹ According to a recent simulation based on genome-wide association study (GWAS) datasets, a highly polygenic model involving a number of common variants of very small effect may explain more than one-third of the total variation in risk of schizophrenia.² On the other hand, interest has been growing in rare variants as potential contributors to the unexplained heritability of schizophrenia.³ This is partly triggered by recent studies establishing an important role for rare genomic copy number variants (CNVs) in the etiology of schizophrenia.⁴ Another potential genetic variation to explain the remaining heritability is rare missense mutations. Kryukov et al⁵ reported that ~20% of new (de novo) missense mutations in humans result in a loss of function, whereas ~53% have mildly deleterious effects and ~27% are effectively neutral with respect to phenotype by a combined analysis of mutations causing human Mendelian diseases, mutations driving human-chimpanzee sequence divergence, and systematic data on human genetic variation. Their results were supported by an independent study.⁶ Because the pressure of purifying selection acting on the mildly deleterious mutations is weak, their cumulative high frequency in the human population is being maintained

by “mutation-selection balance.” This provides support to a speculation that the accumulation of mildly deleterious missense mutations in individual human genomes can be a genetic basis for complex diseases.⁵ The importance of rare missense mutations in schizophrenia is demonstrated by a study of the *ABCA13* gene in which multiple rare (<1%) coding variants were associated with schizophrenia.⁷

We recently performed a GWAS for schizophrenia in a Japanese population.⁸ Although single locus analysis did not reveal genome-wide support for any locus, a shared polygenic risk of schizophrenia between the Japanese and the Caucasian samples was confirmed. In our GWAS, association signals were detected at the regions of the Kalirin gene (*KALRN*) on 3q21.2 and the EPH receptor B1 gene (*EPHB1*) on 3q21-q23, both of which are in the same synaptogenic pathway⁹ (supplementary figure S1). Associations of each gene with schizophrenia have recently received support from independent GWASs in different populations.^{10,11} Furthermore, a rare de novo CNV overlapping with the *EPHB1* gene locus was detected in a patient with schizophrenia.¹²

KALRN is a large neuronal dual Rho guanine nucleotide exchange factor (GEF) that activates small guanine triphosphate-binding proteins of the Rho family, including Rac1.¹³ This activation enables *KALRN* to regulate neurite initiation, axonal growth, dendritic morphogenesis, and spine morphogenesis. Consistent with its biological function, *KALRN* is a key factor responsible for reduced densities of dendritic spines on pyramidal neurons in the dorsolateral prefrontal cortex (DLPFC)¹⁴ observed in postmortem brains from schizophrenic patients. The messenger RNA expression level of *KALRN* is significantly reduced in DLPFC of patients with schizophrenia and strongly correlated with spine density.¹⁵ In addition, *KALRN*-knockout mice not only exhibit spine loss and reduced glutamatergic transmission in the frontal cortex but also schizophrenia-like phenotypes including robust deficits in working memory, sociability, prepulse inhibition, and locomotor hyperactivity reversible by clozapine, an atypical antipsychotic.¹⁶ These synaptic and behavioral dysfunctions are apparent during young adulthood in mice (12 weeks old), which coincides with the onset of schizophrenia in patients. Notably, Disrupted-in-Schizophrenia 1, a prominent schizophrenia risk factor, was shown to be involved in the maintenance of spine morphology and function by regulating access of *KALRN* to Rac1.¹⁷ *EPHB1* belongs to a receptor tyrosine kinase family and controls multiple aspects of neuronal development, including synapse formation and maturation, as well as synaptic structural and functional plasticity. In neurons, activation of EphB receptors by its ligand B-type ephrins induces the rapid formation and enlargement of dendritic spines, as well as rapid synapse maturation. One of the downstream effectors of ephrinB/EphB signaling is *KALRN*. In

young hippocampal neurons, *KALRN* is reported to play an important role in the maturation of synapses induced by trans-synaptic ephrinB/EphB signaling.¹⁸

According to the above-mentioned study,⁵ most missense mutations with a frequency of <1% are mildly deleterious, indicating that a low frequency of missense mutation per se can serve as a strong predictor of a deleterious effect of variants. Therefore, the working hypothesis of the present study is that rare (<1%) missense or nonsense mutations with a larger effect size (eg, OR >1.5) in *KALRN* and *EPHB1* may be genetic risk factors for schizophrenia. Recently, a DNA microarray-based resequencing method has been developed to enable accurate and rapid resequencing analysis of candidate genes.¹⁹ Using this system, we conducted resequencing analyses for all exon regions of *KALRN* and *EPHB1* in 320 schizophrenic patients and found evidence that rare (<1%) missense mutations in *KALRN* are significantly associated with schizophrenia using the 3-phase study design.

Methods and Materials

Subjects

Three sample sets were used in this study. The first sample set, comprising 320 schizophrenic patients (mean age, 54.2 ± 14.1 years, 49.1% male), with long-term hospitalization for severe symptoms, was used to search for rare missense or nonsense mutations. We used the first sample set for mutation screenings because patients with extreme phenotypes (severe symptoms) can be expected to carry more deleterious mutations.²⁰ The second sample set, including 729 cases (45.4 ± 15.1 years, 52.2% male) and 562 controls (44.0 ± 14.4 years, 49.8% male), was used to prioritize detected functional variants for subsequent association analyses. The third sample set, including 1511 cases (45.9 ± 14.0 years, 49.6% male) and 1517 controls (46.0 ± 14.6 years, 49.6% male), was used for association analyses. Age and gender were matched in the second and third sample sets, respectively. All patients were diagnosed according to *Diagnostic and Statistical Manual of Mental Disorders, Fourth Edition*, criteria, and controls were evaluated using unstructured interviews to exclude individuals with history of mental disorders. Detailed information regarding diagnostic procedures is available elsewhere.²¹ All subjects were ethnically Japanese and provided written informed consent. This study was approved by the ethics committees at each participating university.

Array Design for Resequencing Analyses

We used the Affymetrix GeneChip CustomSeq Resequencing Array (Affymetrix, Santa Clara, California) for exon sequencing in the first sample set. These arrays rely on allele-specific hybridization for determining DNA

sequence.¹⁹ Each individual nucleotide of both the sense and the antisense DNA strands is interrogated with four 25-mer probes that differ only with respect to the central position (A, C, G, and T). According to Affymetrix's Custom-Seq Array Design Guide, we designed arrays covering all exon regions of *KALRN* and *EPHBI* (Ensembl release 52 [Human CCDS set]; Transcript: ENST00000360013, ENST00000240874, and ENST00000291478 for *KALRN*; ENST00000398015 for *EPHBI*). Because the principle of the resequencing arrays is based on hybridization, it is necessary to avoid cross-hybridization for accurate resequencing. For this purpose, we removed repetitive elements and highly homologous sequences from the array design.

Array-Based Resequencing

The experiments were conducted according to the manufacturer's instructions (supplementary figure S2). Genomic DNA was extracted from peripheral blood using standard methods. To generate enough target-enriched subject material for hybridization to the arrays, we generated 47 and 14 amplicons per sample for *KALRN* and *EPHBI*, respectively, using long-range polymerase chain reaction (PCR). The PCR conditions were as follows: 94°C for 2 minutes followed by 30 cycles consisting of 94°C for 15 seconds, 68°C for 3 minutes, followed by a final extension of 68°C for 8 minutes, using TaKaRa LA Taq™ (Takara Bio, Otsu, Shiga, Japan). Each PCR product was quantified using PicoGreen (Molecular Probes, Eugene, Oregon), pooled in an equimolar fashion. The PCR products were then purified, fragmented, labeled, and hybridized to the arrays, following the protocol. Finally, the arrays were washed and stained using the GeneChip Fluidics Station 450 (Affymetrix) and scanned using the GeneChip Scanner 3000 (Affymetrix). The data were analyzed using the GeneChip Operating Software (GCOS; Affymetrix), the GeneChip Sequence Analysis Software (GSEQ; Affymetrix), and SeqC (JSI Medical Systems, Kippenheim, Germany; <http://www.jsi-medisys.de/html/products/SeqC/SeqC.htm>) to automate the generation of sequence and genotype calls from the intensity data. In this study, around 17 kb was sequenced per sample, meaning that more than 5.4 Mb was sequenced in total. All missense mutations presented in this study were confirmed using both Sanger sequencing and Custom TaqMan SNP genotyping assays (Applied Biosystems, Foster City, California).

Association Analysis of Each Missense Mutation

Although the rare (<1%) missense mutations were originally discovered among 320 schizophrenic patients, it was possible that a portion of them might have neutral or protective effects.⁵ In addition, it was necessary to reduce the number of statistical tests for multiple comparison problems. To accomplish this, we prioritized rare

(<1%) deleterious variants for subsequent association analyses based on the frequencies in the second case-control sample set because rare deleterious variants relevant to schizophrenia can be assumed to have higher frequency in cases than in controls. The criteria for prioritization were as follows: (1) frequencies of mutations were <1% in controls and (2) frequencies of mutations were higher in cases (ie, OR > 1). Mutations not detected in the second sample set were not followed up in this analysis. The frequencies of such mutations can be so low (<0.0005) that the results of association analyses are unlikely to be statistically significant in our sample size. For mutations meeting the above criteria, we conducted association analyses with schizophrenia using the third sample set. Genotyping was conducted by Custom TaqMan SNP genotyping assays (Applied Biosystems). For quality control, samples with missing call rates of 10% or higher were excluded from the analyses.

Combined Association Analyses

In general, it is difficult to establish an association of a rare mutation with a phenotype because statistical power is limited by low population frequency and because the number of rare variants requires a strict multiple test correction. Therefore, we conducted combined association analyses across rare mutations observed in each gene in the third sample set, comparing the number of mutations in cases with the number in controls. The criteria for mutations included in these analyses were same as the above criteria with 1 exception: Mutations not detected in the second sample set were included in the combined association analyses.

In Silico Analysis

The potential influence of missense mutations was evaluated using PolyPhen-2 (<http://genetics.bwh.harvard.edu/pph2/>) and²² PMut (<http://mmb2.pcb.ub.es:8080/PMut/>)²³ softwares. PolyPhen-2 uses 8 sequence-based and 3 structure-based predictive features and compares a property of the wild-type allele and the corresponding property of the mutant allele. PolyPhen-2 trained on HumDiv datasets is reported to achieve true positive prediction rates of 92% with a false-positive rate of 20%.²² A mutation is appraised qualitatively as benign, possibly damaging, or probably damaging based on naive Bayes posterior probability that a given mutation is damaging. PMut also allows the fast and accurate prediction (~80% success rate in humans) of the pathological character of missense mutations based on the use of neural networks. The final output is a pathogenicity index ranging from 0 to 1 (indexes >0.5 signal pathological mutations).

We also examined evolutionary conservation of the mutated residues and surrounding amino acids. Multiple sequence alignment of human *KALRN* or *EPHBI* with 6 orthologs was performed for this purpose.

Power Calculation

Power calculation was performed with a power calculator called CaTS (<http://www.sph.umich.edu/csg/abecasis/CaTS/>).²⁴ Power was estimated under the following parameter assumptions with respect to association test statistics: genetic relative risk = 2, prevalence of disease = 0.01, risk allele frequency = the values frequency observed in controls, and $\alpha = .05$; a multiplicative model was used.

Statistical Analysis

For the association analysis of each variant, Fisher exact test was used to examine whether rare deleterious variants were significantly overrepresented in the patient group rather than the control group.

A combined association test was performed following a previous study.⁷ In brief, to account for variable sample size, sample size was adjusted to $N=n/(\sum(1/N_i))$, where N_i is the sample size at the i th variant, and n is the number of variants. The number of observed variants was adjusted as $\sum(pi) \times N$, where pi is the frequency of the i th variant. Fisher exact test was used in this test as well to examine an overrepresentation of rare deleterious missense mutations in the patient group rather than control group.

All statistical tests were 1 tailed, and a P value less than 0.05 was considered significant. Bonferroni correction was used for solving multiple testing problems.

Results

Discovery of Mutations

We detected 12 and 6 missense mutations with a frequency of <5% in *KALRN* and *EPHBI*, respectively, among 320 cases in the first sample set (table 1). All but 2 mutations (N2973S in *KALRN* and T981M in *EPHBI*) were novel. All mutations were validated by both Sanger sequencing and Custom TaqMan SNP genotyping assays. In the first sample set, 2 patients were compound heterozygotes for rare missense mutations in the 2 genes. One patient had R410H in *KALRN* and R905C in *EPHBI*. The other had A2382V in *KALRN* and D375N in *EPHBI*. There were no clinical characteristics shared between these patients. No nonsense mutations were identified in this study.

Association Analysis of Each Missense Mutation

In the prioritization phase using the second sample set, T1207M and P2255T in *KALRN* and R637C and R905C in *EPHBI* showed a higher frequency in cases than in controls (table 1). Seven missense mutations (R410H, Q770K, and A2382V in *KALRN* and F151S, D375N, D577N, and T981M in *EPHBI*) were not detected. The frequency of P1695Q was more than 4% both in cases and in controls. Based on our criteria, we selected 4 missense mutations (T1207M and

P2255T in *KALRN* and R637C and R905C in *EPHBI*) for subsequent association analyses using the third sample set.

In the third phase, P2255T showed a nominally significant association with schizophrenia (OR = 2.09, $P = .012$) in the third sample set (table 2). This remained significant after correction for multiple testing of 4 variants (corrected $P = .048$). T1207M in *KALRN* and R637C and R905C in *EPHBI* were also more frequent in cases, although differences were not significant.

We excluded mutations not detected in the second sample set from this analysis. This was supported by a power analysis showing that the third sample set had only 10% power in analysis of very rare mutations.

Combined Association Analysis

In addition to 4 mutations (T1207M and P2255T in *KALRN* and R637C and R905C in *EPHBI*), 7 very rare mutations (R410H, Q770K, and A2382V in *KALRN* and F151S, D375N, D577N, and T981M in *EPHBI*), which were not detected in the second samples set, were included in the combined association analysis. A global comparison of the frequencies of 5 selected mutations in *KALRN* between cases and controls in the third sample set showed a significant increase in frequency in schizophrenic patients (OR = 2.07, $P = .003$) (table 3). This remained significant after correction for multiple testing (corrected $P = .006$). On the other hand, a global comparison of the frequencies of 6 selected mutations in *EPHBI* did not show a significant difference (OR = 1.09, $P = .438$).

In Silico Analysis

Results of *in silico* analysis are shown in table 4. All missense mutations but A2382V in *KALRN* were predicted to have functional relevance by PolyPhen-2 or PMut software.

A multiple alignment of the region of *KALRN* or *EPHBI* containing rare missense mutations with 6 orthologs is shown in table 4. Most of the rare missense mutations showed a high degree of amino acid conservation in different species.

Discussion

In this study, we conducted resequencing analyses for the 2 synaptogenic pathway genes (*KALRN* and *EPHBI*) in schizophrenia using a DNA microarray-based method. After resequencing more than 5.4 Mb, we discovered 17 rare (<1%) missense mutations in *KALRN* or *EPHBI* and detected a significant association between schizophrenia and P2255T in *KALRN*, as well as in the combined association analysis for *KALRN*. These findings are consistent with an estimation that most rare (<1%) missense mutations are mildly deleterious and are associated with a heterozygous fitness loss.⁵

Table 1. *KALRN* And *EPHB1* Missense Mutations Identified in The First Sample Set And Their Frequencies in The Second Sample Set

Gene	Genomic Position	Base Change	dbSNP Reference	AA Change	First Sample Set		Second Sample Set				OR >1
					Homo	Hetero	Genotype Counts		Mutation Frequency		
							SCZ	CONT	SCZ	CONT	
KALRN	125527659	G → A	ss250607852	R410H	0	1	0/0/701	0/0/541	0	0	
KALRN	125531474	T → A	ss250607853	L452Q	0	1	0/1/709	0/2/541	0.0007	0.0018	
KALRN	125600376	C → A	ss250607854	Q770K	0	1	0/0/706	0/0/544	0	0	
KALRN	125656787	C → T	ss250607855	T1207M	0	1	0/2/705	0/1/542	0.0014	0.0009	+
KALRN	125764534	C → A	ss250607856	P1695Q	0	1	0/59/636	1/44/492	0.0425	0.0428	
KALRN	125764599	A → T	ss250607857	M1717L	0	1	0/0/705	0/1/540	0	0.0009	
KALRN	125860927	G → A	ss250607858	R2049K	0	1	0/1/696	0/1/540	0.0007	0.0009	
KALRN	125873259	C → A	ss250607859	P2255T	0	7	1/14/684	0/7/536	0.0114	0.0064	+
KALRN	125873289	C → T	ss250607860	P2265S	1	0	0/6/701	0/7/533	0.0042	0.0065	
KALRN	125873382	G → T	ss250607861	G2296C	0	1	0/1/703	0/1/542	0.0007	0.0009	
KALRN	125876103	C → T	ss250607862	A2382V	0	1	0/0/697	0/0/540	0	0	
KALRN	125920964	A → G	rs16835896	N2973S	0	3	0/3/698	0/6/538	0.0021	0.0055	
EPHB1	136153231	T → C	ss252863894	F151S	0	1	0/0/710	0/0/543	0	0	
EPHB1	136334407	G → A	ss252863895	D375N	0	1	0/0/708	0/0/544	0	0	
EPHB1	136368508	G → A	ss252863896	D577N	0	1	0/0/707	0/0/544	0	0	
EPHB1	136394134	C → T	ss252863897	R637C	0	2	1/1/707	0/2/541	0.0021	0.0018	+
EPHB1	136450890	C → T	ss252863898	R905C	0	3	0/9/695	0/1/543	0.0064	0.0009	+
EPHB1	136460639	C → T	rs56186270	T981M	0	2	0/0/706	0/0/541	0	0	

Note: Genomic position based on NCBI build 36, chromosome 3. Amino acid changes based on NCBI Reference Sequence NP_001019831.2 (2986 aa) for *KALRN* and NP_004432.1 (984 aa) for *EPHB1*. All but N2973S (rs16835896) and T981M (rs56186270) are novel. AA change, amino acid change; dbSNP, Single Nucleotide Polymorphism Database; Homo, homozygote; Hetero, heterozygote; SCZ, schizophrenia; CONT, control; NCBI, National Center for Biotechnology Information.

Schizophrenia is a genetically heterogeneous disorder, with both very rare variants with a high effect size (eg, CNVs in 1q21.1, 15q13.3) and common variants with a low effect size (eg, rs1344706 in *ZNF804A*) involved in its genetic architecture. In this frequency-effect size spectrum, P2255T (OR: ~2, risk allele frequency in controls: ~0.005) is located between the CNV in 1q21.1 (OR: ~10, frequency in controls: ~0.0001)²⁵ and rs1344706[T] in *ZNF804A* (OR: ~1.1, risk allele frequency in controls: ~0.6),²⁶ both of which have been recently associated with schizophrenia. The relatively modest effect size of P2255T compared with that of the above CNVs can be attributable to the difference in the effect of each variant on gene(s): Although CNVs strongly influence the

expression of multiple genes, missense mutations in *KALRN* are presumed to have limited effects on *KALRN* function. P2255T is located in the evolutionally conserved proline-rich region between the C-terminal GEF and SH3 domains²⁷ and is surrounded by 2 nearby phosphorylation sites (S2237 and S2262), according to Human Protein Reference Database (http://www.hprd.org/index_html)²⁸ (figure 1). *In silico* analysis with PhosphoMotif Finder²⁹ shows that T2255 itself can be recognized and phosphorylated by many kinases, suggesting functional implications of P2255T (figure 1). In addition, *in silico* analysis predicts that phosphorylation of T2255 will induce that of nearby S2253. Thus, P2255T may greatly change the phosphorylation status in a narrow

Table 2. Association Analyses of Each Missense Mutation in the Third Sample Set

	AA Change	Third Sample Set		OR	<i>P</i> Value		
		Genotype Counts				Mutation Frequency	
		SCZ	CONT			SCZ	CONT
KALRN	T1207M	0/7/1477	0/3/1482	2.34	.171		
KALRN	P2255T	0/31/1448	0/15/1473	2.09	.012		
EPHB1	R637C	0/4/1477	0/4/1478	1.00	.636		
EPHB1	R905C	0/15/1458	0/12/1466	1.26	.347		

Note: Abbreviations are explained in the first footnote to table 1. *P* values were calculated by Fisher exact test (1 tailed).

Table 3. Combined Association Analysis in The Third Sample Set

Gene	AA Change	Third Sample Set				Combined Analysis	
		Genotype Counts		Mutation Frequency		Gene Based	
		SCZ	CONT	SCZ	CONT	OR	P value
KALRN	R410H	0/0/1481	0/0/1484	0	0	2.07	.003
KALRN	Q770K	0/0/1486	0/0/1490	0	0		
KALRN	T1207M	0/7/1477	0/3/1482	0.0024	0.0010		
KALRN	P2255T	0/31/1448	0/15/1473	0.0105	0.0050		
KALRN	A2382V	0/7/1473	0/4/1480	0.0024	0.0013		
EPHB1	F151S	0/0/1478	0/0/1484	0	0	1.09	.438
EPHB1	D375N	0/0/1483	0/0/1490	0	0		
EPHB1	D577N	0/0/1486	0/2/1483	0	0.000673		
EPHB1	R637C	0/4/1477	0/4/1478	0.0014	0.0014		
EPHB1	R905C	0/15/1458	0/12/1466	0.0051	0.0041		
EPHB1	T981M	0/5/1481	0/4/1484	0.0017	0.0013		

Note: Abbreviations are explained in the first footnote to table 1. *P* values were calculated by Fisher exact test (1 tailed).

region between the C-terminal GEF and SH3 domain. A protein with multiple phosphorylated sites like *KALRN* can be assumed to have an exponential number of phospho-forms, and individual phospho-forms may have distinct biological effects. The diffuse distribution of these phospho-forms at steady state enables the phosphoproteome to encode information and flexibly respond to varying demands.³⁰ Thus, it is conceivable that P2255T may influence such plasticity in *KALRN* by changing the number of phosphorylated sites. Interestingly, detailed examination of clinical information from the first sample set, which was uniquely available to us, revealed that con-

genital or early-onset vascular disease was observed in 5 of 7 cases with P2255T (supplementary table S1). Because *KALRN* may represent a candidate gene for vascular diseases,^{31,32} it is tempting to speculate that P2255T may be a potential risk factor for vascular disease.

In addition to P2255T, we detected multiple rare (<1%) missense mutations in *KALRN* or *EPHB1*. Such variants are not sufficiently frequent to be covered by GWAS nor do they have sufficiently large effect sizes to be detected by linkage analysis in family studies. For modest effect sizes, it is suggested that association testing may require composite tests of overall mutational load,

Table 4. Results of *In Silico*/Conservation Analysis

KALRN		R410H	Q770K	T1207M	P2255T	A2382V	
Analysis		Probably	Probably	Probably	Benign	Benign	
PolyPhen-2		damaging	damaging	damaging			
PMut		Pathological	Neutral	Pathological		Neutral	
Conservation	Human (NP_001019831.2)	LDERSTI	IFLQLRI	IHATEIR	RSQPARL	SILAPLT	
analysis	Chimpanzee (XP_516703.2)	LDERSTI	IFLQLRI	IHATEIR	RSQPARL	SILAPLT	
	Dog (XP_535768.2)	LDERSTI	IFLQLRI	IHATEIR	RSQPSRV	SVLAPLT	
	Cattle (XP_001790302.1)	LDERSTI	IFLQLRI	IHATEIR	RSQPARV	SILTPLT	
	Mouse (XP_001481079.1)	LDERSTI	IFLQLRI	IHATEIR	RSQPPRV	SILAPLA	
	Rat (NP_114451.1)	LDERSTI	IFLQLRI	IHATEIR	RSQPPRV	SILAPLT	
EPHB1		F151S	D375N	D577N	R637C	R905C	T981M
Analysis		Benign	Probably	Possibly	Probably	Probably	Probably
PolyPhen-2			damaging	damaging	damaging	damaging	damaging
PMut		Pathological	Neutral	Neutral	Pathological	Pathological	Pathological
Conservation	Human (NP_004432.1)	QVDFGGR	RCDDNVE	VYSDKLQ	YKGRLLK	LLDRSIP	QSPTAMA
analysis	Chimpanzee	QVDFGGR	RCDDNVE	LLVEQWQ	YKGRLLK	LLDRSIP	QSPTAMA
	(XP_001150963.1)						
	Dog (XP_542791.2)	QVDFGGR	RCDDNVE	VYSDKLQ	YKGRLLK	LLDRSIP	QSPTTMA
	Cattle (XP_614602.4)	QVDFGGR	RCDDNVE	VYSDKLQ	YKGRLLK	LLDRSIP	QSPTAMA
	Mouse (NP_775623.2)	QVDFGGR	RCDDNVE	AYSJKLQ	YKGRLLK	LLDRSIP	QSPSVMA
	Rat (XP_217250.1)	QVDFGGR	RCDDNVE	VYSDKLQ	YKGRLLK	LLDRSIP	QSPSVMA

Note: The bold are the mutated amino acids.

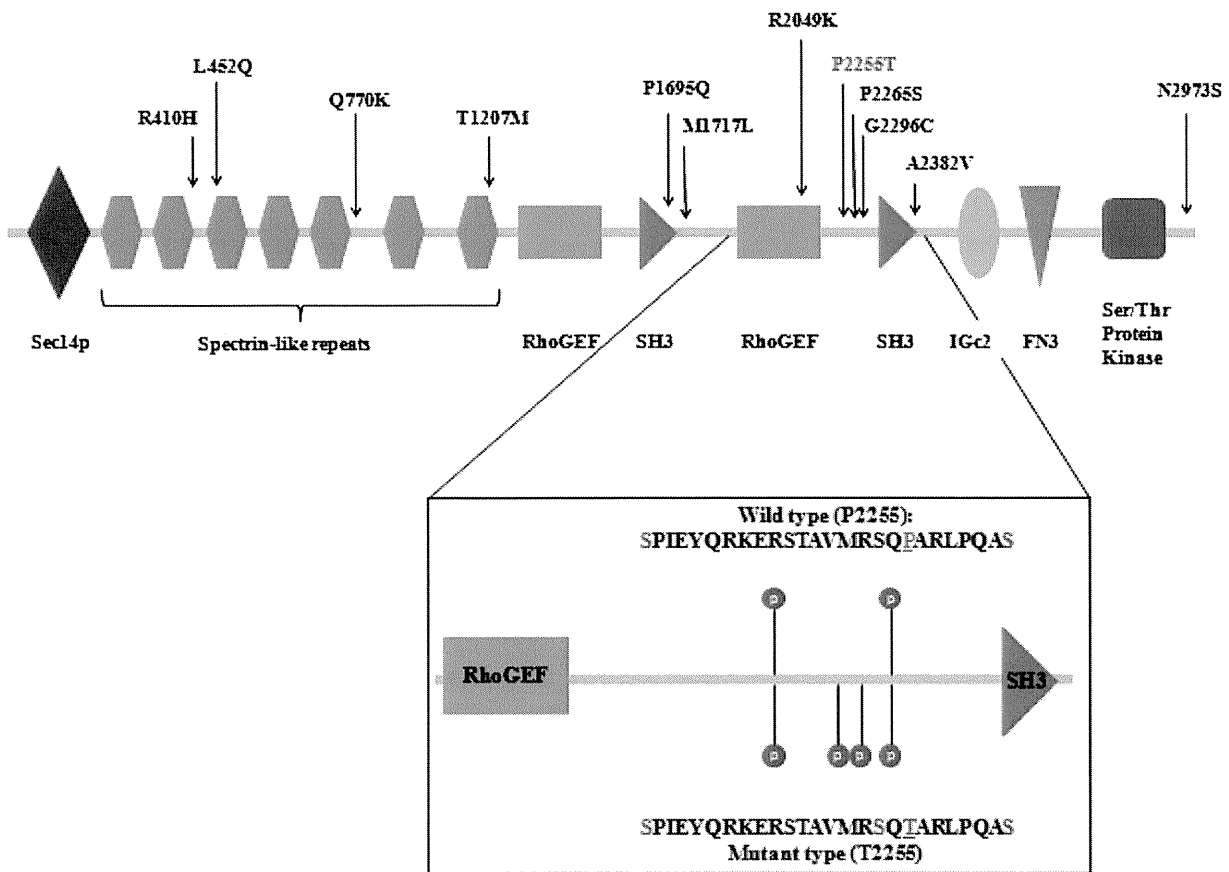


Fig. 1. Rare Missense Mutations in *KALRN* and Change in Phosphorylation Status by P2255T.

comparing frequencies of mutations of potentially similar functional effect in cases and controls. Thus, we also performed combined association analyses for *KALRN* or *EPHBI* and found evidence that multiple rare (<1%) missense mutations in *KALRN* as a whole are associated with schizophrenia. This finding is supported by *in silico* analyses showing that most of the mutations are predicted as being of functional relevance and that they are located in evolutionally conserved regions. In contrast, there were no significant differences in the cumulative frequencies of rare missense mutations in *EPHBI*. This might be due to a type II error. The cumulative frequency of rare mutations of *EPHBI* in controls is almost same as the one of *KALRN* in controls (0.0075 vs 0.0073), indicating that cumulative effect size of rare missense mutations in *EPHBI* may be smaller than the one in *KALRN*. In the mammalian genome, there are 5 different EphB receptors (EphB1, EphB2, EphB3, EphB4, and EphB6), with a high similarity at the amino acid level. Analysis of double and triple knockout mice lacking EphB1, EphB2, and EphB3 in different combinations revealed that EphBs have functional redundancy even though all these EphBs are responsible for spine morphogenesis and synapse formation to varying degrees.³³ This is in contrast with the drastic phenotypes observed in *KALRN*-knockout mice.¹⁶ Therefore, biological effects

of rare missense mutations in *EPHBI* may be compensated for by other intact *EPHBI*s. This might lower the ORs of rare missense mutations in *EPHBI*. Given that all the mutations detected in *EPHBI* were predicted to have pathogenicity by PolyPhen-2 or PMut, a larger-scale case-control study with sufficient power may provide a significant result in a combined analysis for *EPHBI*.

One important aspect of the present study is that we found rare mutations associated with schizophrenia in the *KALRN* gene, in which GWASs detected association signals for schizophrenia. Several studies have recently reported the 1 gene may harbor both rare and common variants associated with the same diseases, including schizophrenia,³⁴ type 2 diabetes,³⁵ and hypertriglyceridemia.³⁶ Given that the cost of whole-genome sequencing is still high to search for rare mutations, resequencing analyses for genes with support from GWAS might be a better strategy for detection of rare mutations with larger effect size.

There are several limitations to this study. First, we could not conduct segregation analyses for mutations due to limited access to family members. Furthermore, given the modest risk (OR ~2), these mutations would show incomplete penetrance. In fact, it is reported that penetrance estimates of CNVs at 1q21.1 and 15q13.3,

both of which show higher ORs, are 0.061 and 0.074, respectively.²⁵ Therefore, a population-based study is a better choice to evaluate genetic associations for missense mutations with modest risk.³⁷ The second limitation is population stratification. Although a Japanese population is considered relatively homogenous, small population stratifications may have influenced our findings.³⁸ However, we believe that the recruitment of subjects in local regions minimized this concern. Third, we did not conduct functional analyses for detected missense mutations. The detailed effects of these mutations on the pathophysiology of schizophrenia need to be examined in a future study. Fourth, our resequencing analyses were not comprehensive in terms of the kind of variants and the number of genes. In other words, the present study did not cover indels or CNVs because of the methodological limitation of the DNA microarray-based method. Because these classes of variants could have a more profound effect on protein function, their genetic contribution to schizophrenia might be revealed in future studies. Also, as shown in *EPHBI*, it is assumed that a variety of molecules or pathways have a role in spine formation or synapse plasticity, which are impaired in patients with schizophrenia, to compensate for each other. A combined analysis of a large number of genes relevant for synaptic function might provide more robust evidence that rare missense mutations as a whole contribute to pathomechanisms of schizophrenia.

In conclusion, we provide the first evidence that multiple rare (<1%) missense mutations in *KALRN* may be genetic risk factors for schizophrenia. Further studies will be needed to examine the pathogenicity of these mutations from a biologic point of view.

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Supplementary Material

Supplementary material is available at <http://schizophreniabulletin.oxfordjournals.org>.

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Genome-Wide Association Study of Schizophrenia in a Japanese Population

Masashi Ikeda, Branko Aleksic, Yoko Kinoshita, Tomo Okochi, Kunihiro Kawashima, Itaru Kushima, Yoshihito Ito, Yukako Nakamura, Taro Kishi, Takenori Okumura, Yasuhisa Fukuo, Hywel J. Williams, Marian L. Hamshere, Dobril Ivanov, Toshiya Inada, Michio Suzuki, Ryota Hashimoto, Hiroshi Ujike, Masatoshi Takeda, Nick Craddock, Kozo Kaibuchi, Michael J. Owen, Norio Ozaki, Michael C. O'Donovan, and Nakao Iwata

Background: Genome-wide association studies have detected a small number of weak but strongly supported schizophrenia risk alleles. Moreover, a substantial polygenic component to the disorder consisting of a large number of such alleles has been reported by the International Schizophrenia Consortium.

Method: We report a Japanese genome-wide association study of schizophrenia comprising 575 cases and 564 controls. We attempted to replicate 97 markers, representing a nonredundant panel of markers derived mainly from the top 150 findings, in up to three data sets totaling 1990 cases and 5389 controls. We then attempted to replicate the observation of a polygenic component to the disorder in the Japanese and to determine whether this overlaps that seen in UK populations.

Results: Single-locus analysis did not reveal genome-wide support for any locus in the genome-wide association study sample (best $p = 6.2 \times 10^{-6}$) or in the complete data set in which the best supported locus was *SULT6B1* (rs11895771: $p = 3.7 \times 10^{-5}$ in the meta-analysis). Of loci previously supported by genome-wide association studies, we obtained in the Japanese support for *NOTCH4* (rs2071287: $p_{\text{meta}} = 5.1 \times 10^{-5}$). Using the approach reported by the International Schizophrenia Consortium, we replicated the observation of a polygenic component to schizophrenia within the Japanese population ($p = .005$). Our trans Japan–UK analysis of schizophrenia also revealed a significant correlation (best $p = 7.0 \times 10^{-5}$) in the polygenic component across populations.

Conclusions: These results indicate a shared polygenic risk of schizophrenia between Japanese and Caucasian samples, although we did not detect unequivocal evidence for a novel susceptibility gene for schizophrenia.

Key Words: Genome-wide association study, *NOTCH4*, polygenic component, schizophrenia, *SULT6B1*

Epidemiologic studies show that genetic factors account for more than 80% of the population variance in susceptibility for schizophrenia; however, as with virtually all other relatively common disorders, it has historically proven difficult to identify the specific genetic variants involved (1).

The application of genome-wide association technology to large case–control samples of mainly European ancestry has recently implicated a number of risk loci for which the evidence is strong. These include loci defined by single nucleotide polymorphisms (SNPs) in which the effects are weak (odds ratios [ORs] 1.1–1.25) among which the strongest supported loci are *zinc finger protein 804 A (ZNF804A)* (2–5), a broad region including the major histocompatibility complex (MHC) on chromosome 6p21.3–22.1 (6–8), *neurogranin (NRGN)*, and *transcription factor 4 (TCF4)* (8).

From the Department of Psychiatry (MI, BA, YK, ToO, KuK, IK, YI, TK, TaO, YF, NI), School of Medicine, Fujita Health University, Toyoake, Aichi, Japan; Medical Research Council (MI, HJW, MLH, DI, NC, MJO, MCO), Centre for Neuropsychiatric Genetics and Genomics, Department of Psychological Medicine and Neurology, School of Medicine, Cardiff University, Cardiff, United Kingdom; Department of Psychiatry (BA, IK, YI, YN, NO), Graduate School of Medicine, Nagoya University, Nagoya, and Japan Science and Technology Agency (BA, YK, ToO, KuK, IK, YI, YN, TK, TaO, YF, MS, RH, MT, NO, NI), the Core Research for Evolutional Science and Technology, Kawaguchi, Japan; Biostatistics and Bioinformatics Unit (MLH, DI), School of Medicine, Cardiff University, Cardiff, United Kingdom; Seiwa Hospital (TI), Institute of Neuropsychiatry, Tokyo; Department of Neuropsychiatry (MS), Graduate School of Medicine, University of Toyama, Toyama; Department of Psychiatry (RH, MT), Graduate School of

Although the robust support for a number of recently implicated loci represents something of a break from the past inconsistencies, little of the genetic variance of schizophrenia can be explained by the loci identified thus far. One explanation for this is that much of the risk is conferred by common but weak genetic effects that require larger samples. Another explanation is that most of the risk cannot be readily detected by genome-wide association studies (GWAS), the missing genetic component being conferred by mutations that exert substantial individual effects that are rare or even unique to individual pedigrees.

Although the relative contributions of these classes of variant awaits empiric resolution, the GWAS of the International Schizophrenia Consortium (ISC) provided strong support for a substantial polygenic contribution (at least 30%) to the population risk of schizophrenia, much of which is conferred by common alleles with small effect sizes (6,9,10). The basic principle of their analysis was that in the presence of a substantial common polygenic component, although most of the individual genetic effects will not be

Medicine, Osaka University, Osaka; Molecular Research Center for Children's Mental Development (RH, MT), United Graduate School of Child Development, Osaka University, Kanazawa University, and Hamamatsu University School of Medicine, Osaka; Department of Neuropsychiatry (HU), Okayama University, Graduate School of Medicine, Dentistry and Pharmaceutical Science, Okayama; and Department of Cell Pharmacology (KoK), Graduate School of Medicine, Nagoya University, Nagoya, Japan.

Authors MI and BA contributed equally to this work.

Address correspondence to Nakao Iwata, M.D., Ph.D., Department of Psychiatry, School of Medicine, Fujita Health University, Toyoake, Aichi, 470-1192; Japan. E-mail: nakao@fujita-hu.ac.jp.

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detectable in current sample sizes, the sum of many such effects across multiple SNPs might differ between cases and controls. After discounting the influence of various potential sources of bias, the authors concluded that the findings were best explained by the existence of an important polygenic component to the disorder comprising a large number of common alleles, although some contribution from low-frequency alleles was not excluded or deemed unlikely (6).

There were two additional striking findings in the ISC article (6). The first was that those alleles selected as “risk” alleles for schizophrenia were also enriched in people with bipolar disorder, supporting the hypothesis of shared genetic susceptibility between these disorders (11,12). The second was that sets of “risk” alleles defined from white individuals of European origin were better at predicting affected status in other white European subjects than they were in African Americans, although an attenuated effect was seen in an African American sample. This may be attributable to differences in allele frequencies and linkage disequilibrium between Europeans and African Americans, although genetic heterogeneity remains a possibility. In this article describing a study that sought novel susceptibility variants, we report the first GWAS for schizophrenia in a Japanese sample. Although the Japanese population is considered relatively homogeneous (13), GWAS studies in other populations strongly suggest that our study of 575 cases and 564 controls is underpowered to detect any findings at genome-wide levels of significance. Thus, we attempted to enhance power by following up the top 150 of the most strongly supported SNPs from the GWAS in an independent sample of 1511 cases and 1517 controls drawn from the Japanese population as well as 479 cases and 2938 controls from the United Kingdom (2). We also sought to examine whether the Japanese population shares with Europeans a polygenic component for schizophrenia and bipolar disorder using schizophrenia and bipolar case–control samples from the United Kingdom that have been previously subjected to GWAS (2,14). Because it is unlikely that stratification effects would bias the allele distributions en masse in samples ascertained in Japan in the same direction as in a European sample, confirmation of a shared polygenic effect argues strongly against the idea that residual uncontrolled stratification is responsible for the effect. Moreover, because rare alleles of large effect are expected to reflect an ongoing process of new mutation (to compensate for their removal by selection), the existence of transcontinental effects also argue against the idea that rare alleles alone can drive this effect, it being unlikely that relatively new variants would be carried on the same ancestral haplotypes in both populations.

Methods and Materials

Participants

We selected 575 patients with schizophrenia (43.5 ± 14.8 years) and 564 healthy controls (44.0 ± 14.4 years) for genome-wide association analysis (our screening GWAS: [JPN_GWAS]). All subjects were unrelated, living in the Tokai area of the mainland of Japan, and self-identified as Japanese. The details of the sample and copy number variation analysis of this GWAS data set have been reported previously (15), and see also Supplement 1.

For follow-up studies, we used an independent Japanese sample comprising 1511 cases (aged 45.9 ± 14.0 years) and 1517 controls (aged 46.0 ± 14.6 years) diagnosed and ascertained in the same way as the GWAS data set. These samples were recruited from three areas on the Japanese mainland, comprising the Kansai and Chugoku areas in addition to the Tokai area. To enhance the sample in the replication analysis, data were added from 934 Japanese

controls genotyped by Illumina550 (Illumina, San Diego, California) as part of the Japanese Single Nucleotide Polymorphisms (JSNP) project (<http://snp.ims.u-tokyo.ac.jp/index.html>). If SNP data were available in the JSNP sample, we merged the two sample sets to form a final Japanese replication sample (we refer this as “Rep_JPN”) comprising 1511 cases and 2451 controls (SNPs genotyped in both samples can be seen in Table S1 in Supplement 2).

We additionally included data from a UK schizophrenia GWAS data set of 479 cases and 2938 controls genotyped using the Affymetrix 500K array (Santa Clara, California), details of which have been reported before (2,14).

For the polygenic component analysis, we also included the Wellcome Trust Case-Control Consortium (WTCCC) bipolar disorder data set of 1868 cases and 2938 shared controls, details of which are reported elsewhere (2,14).

After complete description of the study to the subjects, written informed consent was obtained. This study was approved by the ethics committees of each university participating in this project.

GWAS and Quality Control

Genotyping was performed using the Affymetrix Genome-Wide Human SNP Array 5.0 according to the manufacturer’s protocol. After applying several quality control (QC) criteria (e.g., call rate $\geq 95\%$, autosomal chromosomes, Hardy–Weinberg equilibrium (HWE) $\geq .0001$ and minor allele frequency [MAF] $\geq 5\%$; Supplement 1), the final GWAS consisted of 1108 samples (560 cases and 548 controls) and 297,645 SNPs (MAF $\geq 5\%$).

Q-Q plots were generated on the basis of allele-wise analysis of SNPs that passed QC (Supplement 1), and our observed value of λ is consistent with those generally reported in well-matched samples ($\lambda = 1.065$ and $\lambda_{1000} = 1.117$).

Follow-Up Genotyping

Follow-up genotyping in our independent Japanese case–control sample was performed by Sequenom (San Diego, California) using the Sequenom iPLEX Gold System. Markers that could not be assayed on this platform were genotyped using a TaqMan assay (Applied Biosystems, Foster City, California).

Candidate SNPs were selected for replication as follows. First, the top 200 SNPs were identified (corresponding to $p \sim < 5 \times 10^{-4}$). Highly correlated markers based on $r^2 > .9$ to a more significant marker within 100 kb (r^2 was based on HapMap information [release Number 24, October 2008] and our own GWAS from controls) were then removed. From this list, we included the following: 1) SNPs with $p < 5 \times 10^{-5}$ ($n = 15$ after 11 redundant SNPs removed. Total number = 26. Of these, two SNPs failed for primer design. 2) Under the premise that in GWAS analysis, power favors more common alleles and that the enrichment for true associations is greater in this category of alleles (6), SNPs with MAF $\geq 10\%$ surpassing a more relaxed threshold ($P < \sim 3.5 \times 10^{-4}$) were selected, corresponding to the top 150 SNPs ($n = 76$ after 12 low MAF SNPs and 36 redundant SNPs removed. This resulted in a total of 124. Of these, 5 SNPs failed primer design. We additionally included 13 SNPs that ranked from 151st to 200th on the grounds that they could be included in the Sequenom panels of markers without compromising the design of the higher-priority SNPs. Consequently, 97 SNPs were genotyped in the replication sample, of which 5 did not pass QC on the basis of genotype call rate ($> .95$) and HWE ($p > .001$). All genotype calls were confirmed by visual inspection of cluster plots.

SNP-Based Association Analysis

Consistent with most other GWAS, our study is based upon allele-wise association analysis which assumes an additive model.

Genomic control adjusted p values were also calculated based upon median chi-square statistics. This was performed using PLINK v1.07 (16).

Combined analysis across data sets (Meta_JPN: JPN_GWAS + Rep_JPN, Meta_ALL: JPN_GWAS + Rep_JPN + UK schizophrenia) were conducted using the Cochran–Mantel Haenszel (CMH) approach conditioned by sample as implemented in PLINK v. 1.07.

Polygenic Component Analysis

Discovery (for selecting “score alleles” based on association statistics) and targeting (for calculation of polygenic score) samples are summarized in Table S2 in Supplement 1. Briefly, we examined five discovery and target pairs:

1. Japanese: A set of 280 cases and 274 controls were selected for discovery, and the results were tested in an additional set of 280 cases and 274 controls. The discovery/target samples were selected at random (on the basis of random number generation) from the Japanese GWAS data set. This procedure was repeated 1000 times to ensure the results of this analysis were representative of random divisions of the data set.
- 2, 3. Each of the UK schizophrenia (479 schizophrenia and 2938 controls) (2) and bipolar (1868 cases and 2938 controls) (14) samples were used separately as a discovery data set to generate lists of “risk” alleles that were tested in the full Japanese GWAS sample.
- 4, 5. The full Japanese GWAS sample was used as a discovery data set to generate lists of “risk” alleles that were tested in the UK schizophrenia and bipolar data sets.

For the UK data sets, we used the QC criteria applied in the primary manuscripts (2,14) in which SNPs that deviated from HWE ($p < 1 \times 10^{-5}$ in cases or .001 in control) and had a low call rate ($< 97\%$) were excluded. Note that the criteria for HWE exclusion in the UK data set is slightly different from that in the Japanese GWAS. The precise choice of HWE filter is arbitrary, but we note that both data sets criteria are on the more stringent side of customary practice.

Following the ISC (6), we reduced the set of SNPs by removing SNPs that are in linkage disequilibrium (LD) using the same criteria applied by the ISC (r^2 threshold at .25, window size 200 SNPs). In the tests of the split Japanese data set, we used LD-pruned SNPs selected on the basis of the metrics in the full set of Japanese controls. For all comparisons between Japanese and European data sets, we pruned SNPs sequentially first on the basis of the LD metrics in the discovery data set and second on those in the target data set. Polygenic score was calculated by weighting scores for “risk” alleles by the logOR observed in the discovery data set according to the method used by the ISC (6).

Nominally associated alleles were selected on the basis of the genomic-control adjusted p value in the allele-wise association analysis from the discovery samples at the following liberal significance thresholds (P_T) ($P_T < .5$, $P_T < .4$, $P_T < .3$, $P_T < .2$ and $P_T < .1$). The polygenic score was calculated using PLINK v. 1.07. Nagaelkerke’s pseudo R^2 (a measure of variance explained by a particular factor) was calculated by logistic regression analysis using R (<http://www.r-project.org>) with covariation for “nonmissing SNPs” according to the ISC study (6).

Results

Single Marker Association Analysis

A summary plot of the GWAS (MAF $\geq 5\%$) is presented in Figure S1 in Supplement 1. We did not observe any associations at a widely

used approximate benchmark for genome-wide significance ($p = 7.2 \times 10^{-8}$) (17). The strongest associations were observed at rs12218361, which maps to chromosome 10 at 126.06 Mb and is 3’ of *ornithine aminotransferase* (OAT, $p_{\text{allele}} = 6.2 \times 10^{-6}$, two-tailed), and rs11895771, which maps to chromosome 2 at 37.27 Mb within *sulfotransferase family, cytosolic, 6 B, member1* (SULT6B1, $p_{\text{allele}} = 8.0 \times 10^{-6}$, two-tailed). The most significant 200 markers are given in Table S1 in Supplement 2.

We genotyped 97 LD-pruned SNPs mainly from the top 150 GWAS findings in an independent Japanese replication sample (1511 cases and 1517 controls). For 22 of these, it was possible to expand the control sample size using data from the Japanese population based on the public database (JSNP). Data for 81 SNPs were also available in the UK data set (Affymetrix 500 K chip) and were included in the association analysis. On the basis of the replication sample from Japanese (Rep_JPN) alone, rs9880957 showed the most significant association ($p = 2.8 \times 10^{-3}$, two-tailed, OR = 1.2), but the associated allele was not the same as in the GWAS. Additionally, we undertook set-based analysis (using PLINK) to investigate whether there was an excess of association signals for these top GWAS findings in the replication data set that surpassed nominal p thresholds (e.g., $p < .1$, .05, .01, .001) in the Rep_JPN and UK data sets (10,000 permutation without lambda correction for all SNPs that passed the p threshold). However, no significant enrichment was observed (data not shown). That finding is compatible with the polygenic analysis we describe subsequently and with the now widely accepted hypothesis that common alleles that might be detectable in principle by GWAS exert effects that are too weak to be substantially enriched for associations that surpassed the threshold we specified for follow-up.

In the CMH analysis of the complete Japanese sample (Meta_JPN: JPN_GWAS + Rep_JPN), the best p was found at rs1011131 in LOC392288 ($p = 1.2 \times 10^{-4}$, two-tailed), which is weaker than in the initial GWAS ($p = 2.5 \times 10^{-5}$, two-tailed). Further expanding the sample size by including UK samples (Meta-ALL: JPN_GWAS + Rep_JPN + UK schizophrenia) did not provide convincing support for any locus (Table S1 in Supplement 2). The strongest association signal in Meta-ALL was rs11895771 ($p = 3.7 \times 10^{-5}$, two-tailed) in *SULT6B1*, which had been ranked second in the screening GWAS (Table 1).

Excluding *ZNF804A* (the Japanese data for which were included in the paper by O’Donovan *et al.*) (2), we additionally tested regions containing schizophrenia candidate loci supported by genome-wide significant associations in previous GWAS data sets (6–8). Specifically, we focused on three regions: the MHC region (Chr6 25 ~ 33 Mb), *NRGN*, and *TCF4*. In this analysis, we first imputed ungenotyped SNPs in these regions (boundaries ± 1 Mb) for fine mapping (the imputation method is presented in Supplement 1). None of the specific SNPs at these loci that have been reported by others (6–8) as genome-wide significant were imputable in our Japanese GWAS sample (Figures S2–S4 in Supplement 1). However, interestingly, we did observe a strong, fairly well circumscribed association signal on chromosome 6 in the region of *NOTCH4* (Figure S2 in Supplement 1). Furthermore, genetic association within *NOTCH4* has been reported (18) in another Japanese study (non-overlapping with the present sample) at rs2071287 (Figure S2 in Supplement 1), which is in complete LD ($D' = 1$, $r^2 = .56$) with rs2071286, the best SNP tested in our GWAS data. Because that previously supported SNP (rs2071287) is also associated in our GWAS ($p = 2.1 \times 10^{-3}$), we then followed up this SNP in the Rep_JPN sample; rs2071287 was again significantly associated ($P_{\text{allele}} = .018$, two-tailed, Figure S5 in Supplement 1; note: we could not impute this SNP with high confidence in the UK schizophrenia

Table 1. Top Single Nucleotide Polymorphisms Based on GWAS and Meta-Analysis

CHR	SNP	BP	Closest Gene	Meta_ALL (JPN_GWAS+Rep_JPN+UK_SCZ)			Meta_JPN (JPN_GWAS+Rep_JPN)			JPN_GWAS		Rep_JPN		UK_SCZ						
				A1	MAF	A2	P _{CMH}	OR ^a	L95	U95	P _{CMH}	OR ^a	L95	U95	P _{allele}	OR ^a	P _{allele}	OR ^a	P _{allele}	
2	rs11895771	37266439	SULT6B1	T	.49	G	3.7×10^{-5}	.84	.77	.91	4.1×10^{-4}	.84	.76	.92	8.0×10^{-6}	.64	.14	.92	.033	.84
7	rs1011131	19474460	LOC392288	G	.07	C	1.2×10^{-4}	1.30	1.14	1.48	1.2×10^{-4}	1.31	1.14	1.50	2.5×10^{-5}	1.78	.054	1.17	.63	1.14
14	rs1176970	40505514	LOC644919	G	.15	C	1.4×10^{-4}	1.22	1.10	1.35	3.0×10^{-4}	1.27	1.12	1.44	3.2×10^{-4}	1.58	.041	1.17	.14	1.14
1	rs4908274	103162502	COL11A1	A	.28	T	3.1×10^{-4}	1.20	1.09	1.32	3.1×10^{-4}	1.20	1.09	1.32	1.1×10^{-4}	1.45	.067	1.12	NA	NA
6	rs2294424	11860537	C6orf105	T	.41	C	5.0×10^{-4}	1.15	1.06	1.24	5.0×10^{-4}	1.17	1.28	1.07	1.2×10^{-4}	1.40	.081	1.1	.41	1.08
2	rs13010889	40617519		A	.15	C	.0011	.85	.77	.94	.0016	.85	.77	.94	8.7×10^{-5}	.67	.17	.92	.40	.84
2	rs17026152	40611159		A	.26	G	.0012	.85	.77	.94	.0012	.85	.77	.94	1.3×10^{-4}	.69	.15	.92	NA	NA
6	rs2787566	101985455	GRIK2	A	.04	G	.0014	1.34	1.12	1.61	.0014	1.39	1.1	1.7	2.8×10^{-4}	2.03	.15	1.19	.49	1.16
6	rs2071286	32287874	NOTCH4	T	.19	C	.0014 ^b	.87	.79	.95	.0049 ^b	.86	.78	.96	3.3×10^{-4}	.68	.23 ^b	.93	.13	.87
8	rs17462248	29426926		G	.2	T	.0017	1.16	1.06	1.27	.020	1.14	1.0	1.3	2.1×10^{-4}	1.52	.60	1.04	.030	1.2

p values were calculated on the basis of the allele-wise test (two-tailed).

A1, minor allele based on whole sample; A2, major allele based on whole sample; BP, base position; CHR, chromosome(hg18); GWAS, genome-wide association study JPN_GWAS; our screening GWAS; L95, lower bound of 95% confidence interval for OR; MAF, minor allele frequency based on whole sample; NA, not analyzed; OR, odds ratio; SNP, single nucleotide polymorphism; U95, upper bound of 95% confidence interval for odds ratio; UK-SCZ: UK schizophrenia.

^aOR was calculated on the basis of A1 in Meta-ALL as reference.

^bControls from Japanese SNPs (JSNP) were merged into the replication sample.

data set because of the high missing rate of 12%). Next we conducted a meta-analysis based on Meta_JPN (imputed data from JPN_GWAS was down-weighted using PROPER-INFO from SNPTEST by METAL: <http://www.sph.umich.edu/csg/abecasis/metal/>) and the sample of Tochigi (18). This provided fairly strong evidence for association ($P_{meta} = 5.1 \times 10^{-5}$, two-tailed, Figure S5 in Supplement 1).

Polygenic Component Analysis

p values and pseudo-*R*² statistics (Nagaelkerke's *R*²) for the analysis based on the split Japanese sample are presented in Figure 1 and in Table S3 in Supplement 1. The polygenic scores in the target data were higher in the cases than the controls and, in most cases, significantly so. As in the ISC study, the evidence became stronger and the pseudo-*R*² improved at more liberal *P*_T values. The most significant correlation was found at *P*_T < .5 (*p* = .005). In this condition, the pseudo-*R*² was slightly lower (*R*² = .021) compared with the ISC study (6) in which *R*² ≤ .032 were observed in the Caucasian samples (Figure 1), although we note that the ISC study used information from a greater number of SNPs, with the larger sample available to that group allowing the inclusion of SNPs with MAF as low as 2%.

The results of the analysis based on discovery in the UK schizophrenia data set and targeting the JPN_GWAS are shown in Figure 2 (Table S3 in Supplement 1). Again, as with the ISC data, the signal and predictive power improved at the more liberal thresholds, with only the most relaxed threshold (corresponding to the optimal threshold from the ISC study) attaining significance (*p* = .029). However, the analysis using the WTCCC bipolar sample for discovery and the Japanese as the target did not reveal significant support for shared risk across disorders (Figure 2 and Table S3 in Supplement 1).

Following are the results of the analyses based on discovery in the JPN_GWAS and testing in the UK schizophrenia and bipolar data sets. Alleles trained in this direction were highly significant, but weakly predictive, of schizophrenia status in the UK sample (*p*_{min} = 7.0×10^{-5}) than those analyses based on training in the UK data sets. Again, no significant effect was observed for bipolar disorder. In the schizophrenia analysis, we observed no clear relationship

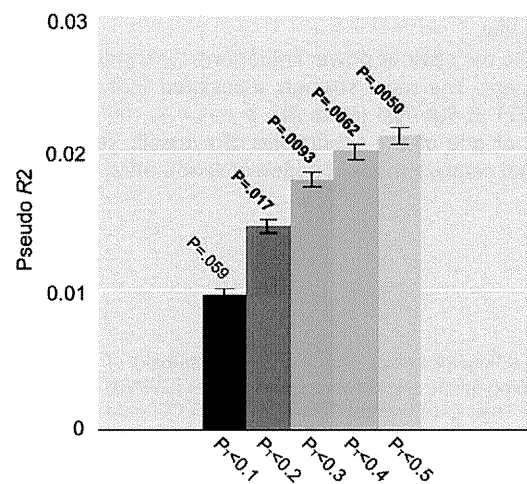


Figure 1. Polygenic component analysis for the pair within screening genome-wide association studies samples. *p*_T = *p* threshold. Pseudo *R*² and *p* values represent the mean and median values, respectively, from 1000 random divisions of the data set. Error bars represent the 95% confidence intervals for *R*² from those repeat analyses. Bold numbers represent significant *p* values (< .05).

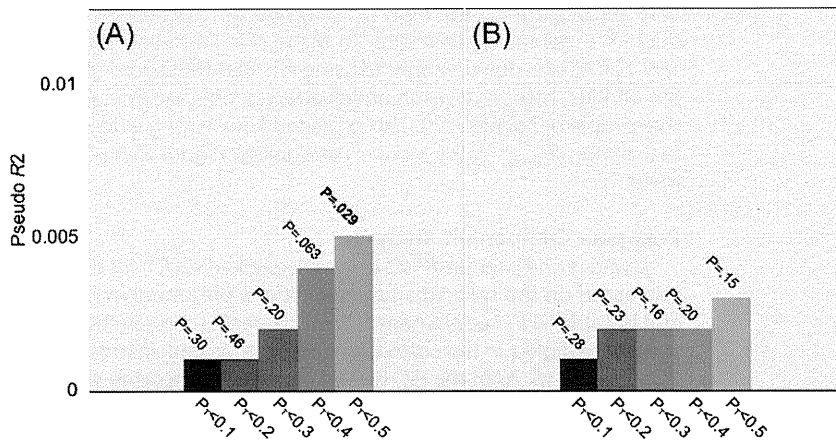


Figure 2. Polygenic component analysis for the pairs of Wellcome Trust Case-Control Consortium (WTCCC) data sets/screening genome-wide association studies (GWAS). **(A)** UK schizophrenia/screening GWAS discovery/target pair. **(B)** WTCCC bipolar/screening GWAS discovery/target pair. $p_T = p$ threshold. Bold numbers represent significant p values ($< .05$).

between the test allele significance threshold (P_T) and either the statistical support or the pseudo- R^2 (Figure 3 and Table S3 in Supplement 1).

Discussion

In this study, we did not detect unequivocal evidence for a novel susceptibility gene for schizophrenia, although our results do provide weak support for association between *SULT6B1* and schizophrenia, and our analyses of previously implicated regions and candidate genes provide support for the hypothesis that previous findings at the MHC region of chromosome 6 may point to *NOTCH4*. The absence of association at genome-wide levels of significance is not surprising given the relatively small size of our GWAS. Recent large-scale GWAS of schizophrenia suggest that the effect sizes of common risk alleles are small (ORs < 1.25). Power analysis suggests that our GWAS has only .18% power under an additive model to detect at $\alpha = 7.2 \times 10^{-8}$, a susceptibility variant with an allele frequency of .3 conferring an OR of 1.25. Clearly, with power like this, it would be extremely unlikely that any one locus would be detected at strong levels of support; however, in the presence of a thousand or more loci as has been suggested (6), the power to detect at least one of these would be considerably greater, albeit the subsequent power to replicate that specific locus would once again be low.

Despite the obvious power limitations, two findings are worthy of comment. The most strongly associated individual SNP was rs11895771 at *SULT6B1* (Meta-ALL $p = 3.7 \times 10^{-5}$). *SULT6B1* is a member of one of the subfamilies of cytosolic sulfotransferases (SULT) that catalyze the sulfonation of xenobiotics, hormones, and

neurotransmitters, including 17β -estradiol and corticosterone (19), functions that are at least plausibly related to schizophrenia (20–22), and brain function (23–25) more widely.

The second locus of interest was *NOTCH4*. *NOTCH4* has been reported to be associated with schizophrenia in a small UK sample (26) (not overlapping with the present sample), but replication data from candidate gene studies have not been strongly supportive. However, a recent synthesis of GWASs as well as a large number of additional subjects reported a genome-wide significant association at rs3131296 (8), which is located within *NOTCH4* (Figure S2 in Supplement 1), although the extensive LD across the MHC region makes pinpointing the source of that signal to a specific gene impossible. It is therefore of interest in our evaluation of the MHC region that the signal clearly maximized to the *NOTCH4* region (Figure S2 in Supplement 1), lending support to the hypothesis that this may be the relevant susceptibility gene in the region. We are unable to evaluate the specific SNP (rs3131296) reported in the SGENE study for the Japanese population because of the failure of imputation. In the Japanese population, the MAF of rs3131296 differs considerably from that in Europeans (MAF = 10% and 2.3% for CEU and JPT populations, respectively, in HapMap Phase 3 data, 13% reported in SGENE), which means the ability of this marker to tag a common functional variant is likely to differ significantly between populations. Given the evidence for association observed in our study and the prior genetic evidence for *NOTCH4*, this locus warrants further detailed analysis in larger and more ethnically diverse samples.

This study provides the first independent (of the samples used by the ISC) replication of the polygenic score analysis reported by

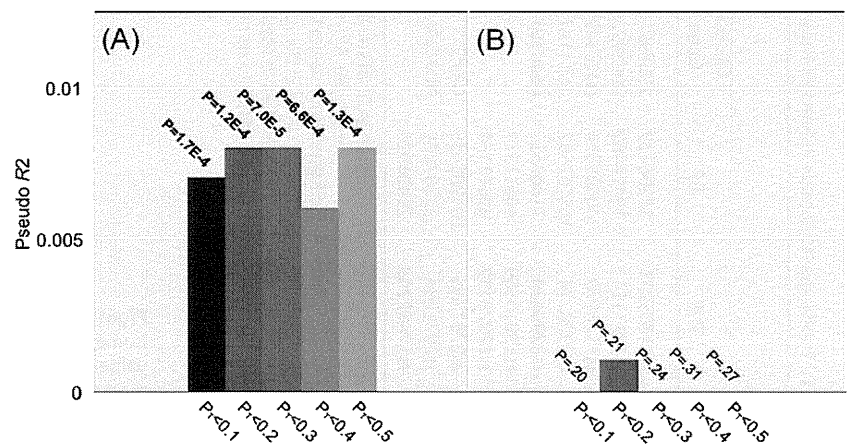


Figure 3. Polygenic component analysis for the pairs of the screening genome-wide association studies (GWAS)/Wellcome Trust Case-Control Consortium (WTCCC) data sets. **(A)** Screening GWAS/UK schizophrenia discovery/target pair. **(B)** Screening GWAS/WTCCC bipolar discovery/target pair. $p_T = p$ threshold. Bold numbers represent significant p values ($< .05$).