

Fig. 1 – Voxels showing significant reduction of fractional anisotropy (FA) in the region around left dorsolateral prefrontal cortex (DLPFC) in the autism spectrum disorder (ASD) group compared with healthy control (HC) participants (right). Scatter plots depict the negative correlation between FA value around the left DLPFC and Social Responsiveness Scale (SRS) score in the ASD group (left).

### 2.2. Correlation between DTI parameters and SRS and IQ

The SRS scores showed significant differences between the ASD and HC groups (ASD: mean 71.5±22.7 for all participants except one with missing data; HC: mean 27.7±17.2; t=3.9; P=0.004) (Table 1). For the ASD group, there was a significant and negative correlation between the FA value in the region around DLPFC and the total SRS score (t=-5.65, p=0.01) when IQ was co-varied (Fig. 1), and the FA value was not significantly correlated with the full IQ when SRS was co-varied. No significant correlation was observed between FA or  $\lambda_1$  values in the other regions and SRS or IQ in either group when controlling for them.

### 3. Discussion

In the present study we compared the white matter structure of autistic children with that of age- , gender-, and handedness-matched healthy control participants by assessing DTI

parameters FA and  $\lambda$  on a voxel-based, whole brain basis. The major finding of this study was that children with ASD (11–18 years of age) showed alterations reflected in reduced FA and  $\lambda_1$  values in the brain regions that play important roles in social cognition and information integration. Another major finding was that the cerebellar vermis in children with ASD showed structural alterations reflected in the increased  $\lambda_1$  value. These results based on DTI measurements in developing children add to the evidence of cerebral and cerebellar white matter structural abnormalities as reported in previous functional and structural studies on ASD.

The largest cluster among the brain regions showing reduced FA and  $\lambda$  values was the ACC (Table 2). This finding is consistent with previous DTI studies in children and adults with ASD (Barnea-Goraly et al., 2004; Thakkar et al., 2008). The ACC is one of the regions that consistently showed structural and functional abnormalities in ASD in previous studies using MRI, positron emission tomography (Haznedar et al., 1997, 2000), and fMRI (Castelli et al., 2002; Hazlett et al., 2004; Pierce et al., 2004; Kennedy et al., 2006). The ACC is considered to be

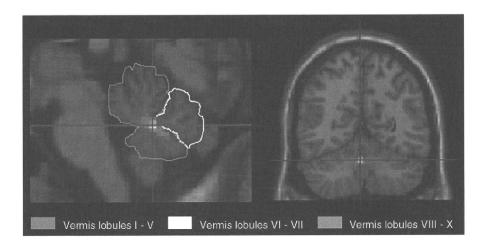


Fig. 2 – Sagittal and coronal views of increased  $\lambda_1$  values in the cerebellar vermis lobules for the autism spectrum disorder (ASD) group compared with healthy control (HC) participants, particularly in the white matter regions between the fastigium and parts of vermis lobules V (orange) and VIII (green), which are near lobules VI and VII (white), respectively.

	Diagnosis	Chronological age	Sex, M/F (6:1)		IQ		CARS-TV score	PARS score	SRS score
Autism spectrum disorder (ASD), $n=7$				VIQ	PIQ	FIQ		(n=6)	(n=6)
A1	HFA	11 y1 m	M	89	92	89	38.5	45	QHIII II
A2	ASP	12 y0 m	M	103	106	104	27.5	25	80
A3	HFA	12 y3 m	M	97	86	91	37.5	43	75
A4	ASP	12 y11 m	M	100	100	100	29	13	31
A5	ASP	14 y6 m	F	95	82	88	36.5	24	65
A6	ASP	16 y7 m	M	81	94	86	39.5		99
A7	HFA	18 y5 m	M	91	92	91	31	23	79
Mean		13.96±2.68 y		93.7±7.41	93.14±8.07	92.71±6.68	34.21±4.91	28.8±12.53	71.5±22.71
Healthy controls (HC)	, n=7					Estimated IQ			(n=7)
H1		10 y9 m	M			112			22
H2		11 y7 m	M			118			53
Н3		11 y10 m	M			124			42
H4		12 y0 m	M			97			39
H5		13 y6 m	F			121			6
H6		15 y3 m	M			124			20
H7		18 y7 m	M			119			12
Mean		13.36±2.74 y				116.43 ± 9.50			27.71±17.2
t-test		n.s.				t=5.40,			t=3.86,
						p=0.0001			p=0.004
$\chi^2$ test			n.s.						

Groups were matched for age, gender and right-handedness. For the estimation of intelligence, the ASD group was administered the Japanese version of the WISC-III or WAIS-III, and the HC group was administered short forms of the Japanese version of the WISC-III or WAIS-III, both of which consist of two subtests (Information and Picture Completion), depending on their age at the time of testing. HFA=High-functioning autism spectrum disorder; ASP=Asperger's disorder; IQ=intelligence quotient; VIQ=verbal IQ; PIQ=performance IQ; FIQ=full IQ; CARS-TV=Childhood Autism Rating Scale-Tokyo version; PARS=Pervasive Developmental Disorders Autism Society Japan Rating Scale; SRS=Social Responsiveness Scale; M=Male; F=Female; n.s.=not significant; y=year; m=month.

associated with social cognitive processes such as self-reflection, empathy and making inferences about others' thoughts (Eisenberger and Lieberman, 2004; Jackson et al.,

2005; Amodio and Frith, 2006). A recently evolved neuronal population in the ACC, the von Economo neurons (VENs), is postulated to serve in socioemotional and higher-order

Brain region	L/R	MNI coordinates					t-value	Cluster size	
			FA		λ <sub>1</sub>				
		X	Y	Z	X	Y	Z		
HC>ASD									
Dorsolateral prefrontal cortex (DLPFC)	L	-40	30	36				14.29	4
Posterior superior temporal sulcus (pSTS)/temporo-parietal junction (TPJ)	L				-52	-66	40	11.97	4
Temporal pole (TP)	R	48	6	-16				12.19	6
Amygdala	R	28	-4	-12				13.86	16
Anterior cingulate cortex (ACC)	R				16	32	28	15.74	44
		0	10	44				14.59	28
Anterior corpus callosum (aCC)		0	16	16				13.57	21
	L	-14	24	14				12.49	12
Superior longitudinal fasciculus (SLF)	R	32	-6	28				12.56	5
Occipitofrontal fasciculus (OFF)	R	22	-4	34				11.24	4
ASD>HC									
Cerebellar vermis lobules	L				-4	-58	-26	22.64	69

cognitive processing (Nimchinsky et al., 1999). Simms et al. (2009) reported alterations in the density of VENs in the ACC in autism. Taken together with previously reported findings using various approaches, our present findings also indicate maldevelopment of the ACC in ASD.

The posterior STS/TPJ and TP/amygdala, around which reduced FA and  $\lambda_1$  values in ASD were found in our study and previous DTI studies (Barnea-Goraly et al., 2004; Lee et al., 2007; Cheung et al., 2009), are also implicated in socioemotional processes such as theory of mind function and social perception through face or human action (Adolphs, 2001; Siegal and Varley, 2002; Dolan and Vuilleumier, 2003; Frith and Frith, 2003, 2006; Morris et al., 2005; Fitzgerald et al., 2006; Olson et al., 2007; Saarela and Hari, 2008; Hall et al., 2009). It has been reported that individuals with ASD show less activation of the STS during tasks involving mental attribution to moving geometric figures (Castelli et al., 2002), biological motion cues (Pelphrey and Carter, 2008), and human speech perception (Gervais et al., 2004). In addition, less activation of the amygdala/TP in response to faces (Baron-Cohen et al., 1999; Critchley et al., 2000; Wang et al., 2004; Williams et al., 2006; Ashwin et al., 2007) has been documented in ASD compared to control participants.

In addition, we found reduced FA around the DLPFC, which concurs with the findings of other DTI studies in ASD (Sundaram et al., 2008; Thakkar et al., 2008; Cheung et al., 2009; Ke et al., 2009; Pardini et al., 2009; Pugliese et al., 2009; Sahyoun et al., 2010). The DLPFC has been found to interact with posterior association areas such as STS and to manipulate cognitive representation of socioemotional processes (Taylor and Fragopanagos, 2005; Yurgelun-Todd and Killgore, 2006; Weissman et al., 2008). Our finding is consistent with the fMRI finding indicating abnormal connectivity among the DLPFC, medial PFC, and posterior association areas (Wicker et al., 2008). Furthermore, we found a significant association between FA around the DLPFC and degree of social impairment as measured by the SRS in the ASD group even when IQ was co-varied (Fig. 1), indicating that the more severe the social impairment, the more altered the DLPFC structure. Given our findings, structural alteration of the DLPFC might be an index of social impairment in ASD. The finding might be influenced by the age range of our participants restricted to childhood. There are some indications that prefrontal activity is pronounced during late childhood (Monk et al., 2003; Yurgelun-Todd and Killgore, 2006; Blakemore, 2008). For example, when attention was unconstrained for emotionally engaging face stimuli, children (9-17 years of age) showed greater activation of the prefrontal cortex relative to adults (25-36 years of age) (Monk et al., 2003). Similarly, Yurgelun-Todd and Killgore reported that activity within the prefrontal cortex during fear perception was significantly and positively correlated with age (8-15 years). Therefore, our present data support the concept that efficient connectivity of the prefrontal region during late childhood is specifically important for appropriate social responsiveness.

Reduced FA was also observed around the SLF and OFF in the present study, confirming the findings of previous studies (Pugliese et al., 2009). The SLF contains long association fibers connecting all the lobes, and the OFF connects the occipital and frontal regions of the brain. In addition, reduced FA was also

observed around the aCC, which is important for connecting prefrontal regions (Witelson, 1989; Huang et al., 2005). This finding concurs with previous DTI studies in ASD (Alexander et al., 2007; Thakkar et al., 2008). The finding of disorganization of these commissural and association fibers in autistic children may support the cortical underconnectivity hypothesis of ASD (Belmonte et al., 2004; Geschwind and Levitt, 2007; Müller, 2007).

Another major finding of the present study was the significantly increased  $\lambda_1$  in the cerebellar white matter regions between the fastigium and parts of vermis lobules V and VIII, which are near lobules VI and VII, respectively. To date, the most consistent pathologic finding in ASD is a decreased number of Purkinje cells in the cerebellum vermis or hemisphere (Bauman and Kemper, 1985; Courchesne, 1997; Bailey et al., 1998). In MRI studies, cerebellum vermis lobules VI and VII have been consistently reported to be smaller in individuals with ASD than in typically developing children, adolescents, and adults (Courchesne et al., 1988, 1994, 2001; Ciesielski et al., 1997; Carper and Courchesne, 2000; Kaufmann et al., 2003). It has been shown that the volume of vermis lobules VI and VII is related to stereotyped behaviors (Pierce and Courchesne, 2001) and attention-orienting deficits (Harris et al., 1999; Townsend et al., 1999), which are often found in individuals with ASD. Recent mouse knockout studies showed that the loss of Ca<sup>2+</sup>-dependent activator protein for secretion 2 (CAPS2) function resulted in several cerebellar morphologic abnormalities such as fewer branched dendrites on Purkinje cells and increased granule cell death in vermis lobules VI and VII as well as autistic-like behavioral phenotypes (Sadakata et al., 2007a,b). Given these findings, lobules VI and VII might have been expected to show structural abnormalities in our DTI analysis. Unexpectedly, our whole-brain  $\lambda$  analysis detected structural alterations in neighboring vermis regions (V and VIII). The reason for this is unclear; however, we speculate that the neural connections with neighboring lobules V and VIII might develop excessively to compensate for possible functional abnormalities in lobules VI and VII. Additional studies are necessary to confirm this speculation.

The implications of the present study are as follows. First, our participants were relatively homogenous in that all were high functioning, and both groups were well matched with respect to age, gender, and handedness, indicating successful differentiation of the two groups. Second, our whole-brain analysis of multiple DTI parameters (FA,  $\lambda_1$ ,  $\lambda_2$ , and  $\lambda_3$ ) revealed structural alterations and also demonstrated a brain–behavior relationship in ASD, which has not been reported previously. Potential study limitations include the fact that the neural mechanism in ASD could not be clearly explained by the alteration in DTI parameters because the physiologic meanings of FA and  $\lambda$  are unclear, sample size was small, and IQs were not group-matched.

Behavioral symptoms of ASD manifest early in life and develop under interaction between genetic and environmental factors, resulting in diverse clinical manifestations. To explore the diverse developmental trajectories in ASD in the future, a database of brain structure and their relationship to behavioral phenotypes should be established from infancy and through childhood, adolescence, and adulthood. Whether the brain alterations identified in the present study are specific to autism or common to other developmental disorders is another issue that warrants investigation.

### 4. Experimental procedures

### 4.1. Participants

Children with ASD and HC children were recruited from local schools and the community by advertisement and were enrolled in the study. The ASD group (n=7) consisted of children diagnosed with high-functioning ASD. The HC group (n=7) was selected to be matched for age, gender and righthandedness (Table 1). Diagnosis of ASD, which had been made by independent clinicians in the community, was confirmed according to DSM-IV-TR criteria (American Psychiatric Association, 2000) by our research team, which included an experienced child psychiatrist (Y. K.). To corroborate the ASD diagnosis, autistic symptom severity was evaluated using the Childhood Autism Rating Scale-Tokyo version (CARS, Schopler et al., 1988; Japanese translated and adapted version [CARS-TV], Kurita et al., 1989), and all participants with ASD were confirmed to score higher than the cutoff score of 25.5 for pervasive developmental disorders (PDD) (Tachimori et al., 2003). In addition, in semistructured interviews, the Pervasive Developmental Disorders Autism Society Japan Rating Scale (PARS), which has been validated for the Japanese PDD population (Adachi et al., 2006; Kamio et al., 2006) was completed for all participants except one in the ASD group; all six scored higher than the cutoff score of 9 for PDD (Table 1). Participants in the HC group were confirmed to have no history of neurologic or psychiatric disorder according to interviews with their parents.

Intellectual function of the participants with ASD was evaluated using the Japanese version of the Wechsler Intelligence Scale for Children-Third Edition (WISC-III, Wechsler, 1991; Japanese translated and adapted version, Azuma et al., 1998) or the Japanese version of the Wechsler Adult Intelligence Scale-Third Edition (WAIS-III, Wechsler, 1997; Japanese translated and adapted version, Fujita et al., 2006), whereas that of the HC group was estimated using the short form of the Japanese version of the WISC-III (Inada and Kamio, in press) or the short form of the Japanese version of the WAIS-III (Dairoku et al., 2009), both of which consist of two subtests, Information and Picture Completion (Table 1).

All participants and their parents provided informed consent to participate in the study. The study protocol was approved by the Research Ethics Committee of the National Center of Neurology and Psychiatry, Japan.

### 4.2. Assessment of social responsiveness

To measure the degree of social responsiveness of our participants, the Japanese version of the SRS (Constantino et al., 2000; Kamio et al., 2009) was completed by parents. The SRS is a 65-item questionnaire that was developed to assess social impairment observed in autism. Completed by a parent or teacher in 15 to 20 min, the SRS provides a clear picture of a child's social awareness, social information processing, reciprocal social communication, social anxiety/ avoidance, and autistic preoccupations and traits. These

items form a single measure of autistic traits rather than distinct scores associated with specific symptom domains (Constantino et al., 2000, 2004; Constantino and Todd, 2005), and can quantify social impairments across the entire child population including both clinical and nonclinical populations (Constantino and Todd, 2003). Total SRS raw scores range from 0 (corresponding to high social competence) to 195 (corresponding to significant social impairment as observed in severe autism). Scores between 60 and 80 are associated with mild forms of autism (Constantino and Todd, 2005).

### 4.3. Magnetic resonance imaging

Magnetic resonance images were obtained using a 3.0 T MRI device (Achieva Series Quasar Dual 3.0 T; Philips Medical Systems, Best, Netherlands) and an 8-channel head array coil for parallel imaging (SENSE). Single-shot echo planar DTI was performed using a twice refocused, spin echo sequence with the following sequence parameters: repetition time/echo time=7420/88 ms; motion probing gradient 32 axes (b value 800 [s/mm²]: 32; b value 0 [s/mm²]: 1); 128×128 matrix; 1.8×1.8 mm in-plane resolution; 60 slices; 2 mm (0 mm gap) slice thickness; scan duration 5′33.5″.

### 4.4. Image analysis

To correct for image distortion owing to eddy current, diffusion-weighted (DW) images obtained at a b value of 800 s/mm<sup>2</sup> were first coregistered with the diffusionunweighted image at b=0 s/mm<sup>2</sup> using software of the MRI system. For each participant, voxel-wise calculation of DTI and subsequent mapping of diffusivity were performed using inhouse software running on Interaction Data Language (IDL) ver. 6.1 (ITT Visual Information Solutions, Boulder, CO). Using 32 DW images ( $b=800 \text{ s/mm}^2$ ) and one diffusion-unweighted image  $(b=0 \text{ s/mm}^2)$ , a 3×3 DTI was obtained for each voxel, from which three eigenvalues,  $\lambda_1$ ,  $\lambda_2$ , and  $\lambda_3$ , were calculated; FA was calculated from the eigenvalues. In addition, for group analyses (HC vs. ASD, ASD vs. HA), the FA,  $\lambda_1$ ,  $\lambda_2$ , and  $\lambda_3$  maps for each participant were analyzed separately using Statistical Parametric Mapping (SPM)2 software (http://www.fil.ion.ucl. uk/spm/software/spm2/). Maps were spatially normalized to the Montreal Neurological Institute (MNI) T2 template and smoothed using a Gaussian filter of 6 mm full width at half maximum. Random-effect analyses were thresholded for FA,  $\lambda_1$ ,  $\lambda_2$ , and  $\lambda_3$  values between the ASD and the HC groups at P<0.05, Family Wise Error (FWE), with an extent threshold of three voxels.

### 4.5. Correlational analysis

For each of the ASD and HC groups, a multivariate regression method (SPSS v17.0) was used for correlation analysis between the DTI parameters and the scores of SRS and IQ (p<0.05). The DTI parameters were those of averaged FA and  $\lambda$  values in the spherical region of interest (ROI; radius, 5 mm), the center of which was the peak voxel in each cluster showing significant difference.

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## Reliability and validity of the Japanese version of the Modified Checklist for autism in toddlers (M-CHAT)

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### ABSTRACT

Early detection and intervention is essential for children with autism spectrum disorders (ASD). Therefore, we examined the reliability and validity of the Japanese version of the Modified Checklist for autism in toddlers (M-CHAT), a 23-item, yes-no questionnaire regarding early autistic symptoms completed by parents of children at 18-24 months of age. Herein, the reliability of the M-CHAT was investigated for children 4-20 months of age. The M-CHAT score (the number of failed items) was found to be significantly correlated among 24 mother-father pairs (Pearson's r = .933), representing good interrater reliability. The test-retest reliability was satisfactory, with 22 mothers providing almost equal M-CHAT scores on two different occasions (r = .990). Significant correlations were observed between the M-CHAT score and the Childhood Autism Rating Scale-Tokyo version score in 25 two-year-old children (r = .581), indicating good concurrent validity. The M-CHAT score was significantly higher in 20 children later diagnosed with ASD compared with reference children (n = 1167), revealing sufficient discriminant validity. A short version of the M-CHAT using 9 items was proposed and effectively differentiated children with ASD from reference children. The efficacy of the Japanese version of the M-CHAT was demonstrated for first-level screening in the general population.

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### 1. Introduction

Autism spectrum disorders (ASD) reportedly affect 1–2% of children (Baird et al., 2006; Baron-Cohen et al., 2009; Kawamura, Takahashi, & Ishii, 2008). Early detection is considered essential for children with ASD based on findings of clinical studies that have shown that early intervention subsequent to early detection can enhance their potential (Dawson et al., 2010) and lead to "optimal outcome" (Sutera et al., 2007). To date, some ASD screening tools, such as the Checklist for autism in toddlers (CHAT) (Baron-Cohen, Allen, & Gillberg, 1992), the Modified Checklist for autism in toddlers (M-CHAT) (Robins, Fein, Barton, & Green, 2001), the Early Screening of Autistic Traits Questionnaire (ESAT) (Swinkels et al., 2006), and the Baby and Infant Screen for Children with aUtIsm Traits (BISCUIT) (Matson et al., 2009) have been developed for the early detection of ASD.

Japan has established a national health check-up system that provides all children, from infants to 3-year-olds, with regular free routine check-ups. This check-up system is well organized and health check-ups at 18 and 36 months effectively determine the children's language/intellectual development. Because socio-communication abnormalities in ASD begin to manifest at 1 year of age (Kamio, Tobimatsu, & Fukui, in press), the check-up at 18 months of age appears to provide a good opportunity to detect early symptoms of ASD. The M-CHAT, a 23-item, parent-report questionnaire (Robins et al., 2001), may

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enhance the current system, since the checklist was developed for children 18–24 months of age and is easy to administer without increasing the burden on both the families and check-up staff.

For the aforementioned reasons, we developed the Japanese version of the M-CHAT. After a preliminary study (Kamio & Inada, 2006), with the permission of the authors, we added some illustrations (items 7, 9, 17, and 23: see M-CHAT Information, Robins, n.d.) in order to encourage caregivers to notice negative symptoms (attenuated typical development). The primary aim of this study is to establish the reliability and validity of the Japanese M-CHAT.

To the best of our knowledge, the inter-rater reliability and test-retest reliability of the M-CHAT have not been examined. This may be because most children in the targeted age range (18–24 months) are expected to pass almost all the M-CHAT items, thus, yes-no answers are skewed toward positive responses. To avoid such a "ceiling effect" in responses, inter-rater and test-retest reliabilities should be examined for children at different developmental stages. In the current study, the reliability of the Japanese version of the M-CHAT was examined for young children in the first and second years of life who are expected to fail some M-CHAT items (Inada, Kamio, & Koyama, in press).

Although satisfactory internal consistency of the M-CHAT has been reported (Kleinman et al., 2008; Robins et al., 2001), the checklist contains some dummy/buffer items (Baron-Cohen et al., 1992; Robins et al., 2001) that obscure the aim of screening from caregivers. Those items are not directly related to the behaviors of children with ASD; in other words, some items of the M-CHAT may not have face validity. As 6 items were selected from the original M-CHAT as "critical items" (Kleinman et al., 2008; Robins et al., 2001), a more simple but effective short version might be useful. Thus, a second aim of this study is to explore the critical items on the Japanese M-CHAT in order to create a short version.

### 2. Methods

### 2.1. Participants and procedures

### 2.1.1. Reliability sample

Reliability data were collected from voluntary parents recruited from local nursery schools or several communities. Interrater reliability data were collected from the mothers and fathers of 24 children (13 males; mean age = 10.0 months, SD = 4.3 months, range = 4-17 months) and test-retest reliability data were collected from the mothers of 22 children (12 males; mean age = 12.8 months, SD = 5.8 months, range = 4-20 months) with a mean interval of 8.3 days (range = 4-14 days). The caregivers were not particularly familiar with ASD or the general development of children and were not told the aim of the study before giving their answers.

### 2.1.2. Validity sample

The concurrent validity was examined using data from 25 children (17 males; mean age = 24.0 months, SD = .7 months, range = 23–26 months), who lived in Nishi-Tokyo city, Japan, and were consecutively referred to the authors due to developmental concerns between December 2008 and December 2009. The mothers of the children answered the M-CHAT prior to the diagnostic evaluations. Two child psychiatrists and two certificated clinical psychologists who did not know the results of the M-CHAT assessed the children in a team using the Childhood Autism Rating Scale-Tokyo version (CARS-TV), which is the Japanese version of the CARS (Schopler, Reichler, DeVellis, & Daly, 1980) and has been demonstrated to be reliable and valid for evaluating autistic symptoms in children (Kurita, Miyake, & Katsuno, 1989; Tachimori, Osada, & Kurita, 2003). The CARS has been reported to be useful for toddlers 2 years of age (Chlebowski, Green, Barton, & Fein, in press).

The discriminant validity was investigated using M-CHAT data from 1187 children (612 males), which had been collected at free health check-ups for 18-month-olds in Munakata city, Fukuoka prefecture, Japan, between April 2005 and March 2007. Of these children, 20 (16 males) were identified as having ASD at age 3. Based on detailed clinical assessments and comprehensive parental interviews on their developmental history independent of the M-CHAT results, the diagnoses were confirmed by expert consensus among our research team (two experienced psychiatrists and a certificated clinical psychologist) according to the DSM-IV-TR criteria for pervasive developmental disorders (PDD) (American Psychiatric Association, 2000). Of the 20 children with ASD, 7 children (6 males) were diagnosed with autistic disorder (autism) and the remaining 13 children (10 males) were diagnosed with PDD not otherwise specified (PDD-NOS). The remaining 1,167 children (596 males) who completed a 3-year health check-up conducted by health professionals including pediatric neurologists were grouped as a reference group in this study.

### 2.1.3. Ethical issues

The protocol of this study was approved by the ethics committee of the National Center of Neurology and Psychiatry, Japan. Written informed consent to participate in our study was obtained from the caregivers of each of the children.

### 2.2. Data analysis

### 2.2.1. Reliability

For inter-rater reliability, Pearson's correlation coefficient (r) was calculated between the M-CHAT score (the number of failed items) of mothers and that of fathers. In addition, we calculated kappa coefficients ( $\kappa$ ) for each of the 23 items. Similarly, the test-retest reliability was examined using the first (test) and second (retest) answers.

### 2.2.2. Validity

The concurrent validity was examined by calculating Pearson's correlation coefficient (r) between the M-CHAT score and the CARS-TV total score. To examine the discriminant validity, we compared the M-CHAT scores among the autism, PDD-NOS, and reference groups by analysis of variance (ANOVA), and post hoc comparisons were conducted with Bonferroni's correction. The pass rate for each of the 23 items was compared between the ASD (both the autism and PDD-NOS) and the reference groups using Fisher's exact test with Bonferroni's correction.

### 2.2.3. Short version

Based on the above item analysis, we proposed a short version of the Japanese M-CHAT and compared the internal consistency (Cronbach's  $\alpha$ ) between the full and short versions. Similar to the full version (see above), a correlation with the CARS-TV total score and comparison among the autism, PDD-NOS, and reference groups were examined.

### 2.2.4. Screening utility

Using a validation sample from Munakata city (n = 1187), we preliminarily reported the sensitivity, specificity, and positive/negative predictive values for predicting later diagnosis of ASD in children at 18 months of age.

All statistical analyses were performed using SPSS 18.0J for Windows and the level of statistical significance was set at .05 (two-tailed).

### 3. Results

### 3.1. Reliability

### 3.1.1. Inter-rater reliability

A significant positive correlation was observed between the M-CHAT scores (the number of failed items) of mothers and fathers (r = .933, p < .001). As shown in Table 1, kappa coefficients ( $\kappa$ ) for inter-rater reliability were significant for 14 ( $\kappa$  ranges .417–1.000, mean = .712) out of 18 items. For 5 items on which the coefficient could not be calculated, the raw agreement rates were very high, exceeding .917. However,  $\kappa$  were not significant for the following four items: item 8 (Functional play), 14 (Responds to name), 18 (Unusual finger movement) and 22 (Stares at nothing).

### 3.1.2. Test-retest reliability

The first and second M-CHAT scores were significantly positively correlated (r = .990, p < .001). As shown in Table 1, 16 of the 18 items showed moderate to high test–retest reliability ( $\kappa$  ranges .645–1.000, mean = .912). For the 5 items for which kappa values could not be calculated, the raw agreement rates were satisfactorily high. However, the  $\kappa$  were not significant for items 11 (Oversensitive to noise) and 14 (Responds to name).

**Table 1** Kappa coefficients ( $\kappa$ ) of the Japanese version of the Modified Checklist for autism in toddlers (M-CHAT).

Item no.	M-CHAT item	Inter-rater $(n = 24)$	Test-retest $(n = 22)$	
1.	Enjoys being swung	.958 <sup>a</sup>	1.000 <sup>a</sup>	
2.	Interest in other children	.700**	1.000**	
3.	Climbs up stairs	.915**	1.000**	
4.	Enjoys peek-a-boo	.917ª	1.000**	
5.	Pretend play	.600**	1.000**	
6.	Imperative pointing	.727**	.909**	
7.	Declarative pointing	.909**	.908**	
8.	Functional play	.111	.909**	
9.	Brings objects to show	.808	.817**	
10.	Eye contact	.958ª	.955 <sup>a</sup>	
11.	Oversensitive to noise	.514**	065	
12.	Responds to smile	1.000 <sup>a</sup>	1.000 <sup>a</sup>	
13.	Imitation of action	.667**	.879**	
14.	Responds to name	-,167	.327	
15.	Point following	.664**	.792**	
16.	Walking	1.000**	1.000**	
17.	Gaze-following	.750 <sup>**</sup>	.817**	
18.	Unusual finger movement	.111	.645**	
19.	Gaining parent's attention	.565**	1.000**	
20.	Wondering hearing	.958ª	1.000 <sup>a</sup>	
21.	Understands what is said	.727**	1.000**	
22.	Stares at nothing	.263	.955ª	
23.	Social reference	.417*	.909**	

p < .05. p < .01.

<sup>&</sup>lt;sup>a</sup> Figure is the raw agreement rate as kappa could not be calculated.

Table 2
The M-CHAT scores (the number of failed items) at 18 months of age.

	Mean	SD	F <sub>(2,1184)</sub>
Full version			73.8**
Autism $(n=7)$	4.71	4.54	
Pervasive developmental disorders			
not otherwise specified (PDD-NOS) $(n = 13)$	2.46	1.56	
Reference (n = 1167)	.58	.99	

*Note*: Any post hoc comparisons with Bonferroni's correction were significant (p < .001).

### 3.2. Validity

### 3.2.1. Concurrent validity

The M-CHAT score was significantly correlated with the CARS-TV total score in 25 children 23–26 months of age (r = .581, p = .002).

### 3.2.2. Discriminant validity

As shown in Table 2, the mean M-CHAT score at 18 months of age differed significantly among the diagnostic groups. The M-CHAT score was the highest in the autism group, followed by the PDD-NOS and the reference groups, respectively.

Table 3 shows that the pass rate at 18 months of age in children later diagnosed with ASD was significantly lower than that of the reference children for the following 11 items: item 5 (Pretend play), 6 (Imperative pointing), 7 (Declarative pointing), 8 (Functional play), 9 (Brings objects to show), 13 (Imitation of action), 14 (Responds to name), 15 (Point following), 17 (Gaze-following), 21 (Understands what is said), and 23 (Social reference).

### 3.3. Short version

Based on the above analyses for each of the 23 items, we developed a short version of the Japanese M-CHAT consisting of 9 items (items 5, 6, 7, 9, 13, 15, 17, 21, and 23). Cronbach's  $\alpha$  for the short version (.752) improved compared to that of the full version (.556), suggesting that the short version has superior internal consistency compared to the full version.

The number of failed items in the short version (M-CHAT-SV score) was significantly correlated with the CARS-TV total score in 25 children 23–26 months of age (r=.615, p=.001). The mean M-CHAT-SV score at 18 months of age differed significantly among the diagnostic groups (F=81.5, d.f. = 2, 1184, p < .001), and all post hoc comparisons with Bonferroni's correction demonstrated a significant difference in the mean M-CHAT-SV score among the autism (mean = 3.43, SD = 3.41), PDD-NOS (mean = 1.62, SD = 1.71), and reference (mean = .20, SD = .72) groups.

Table 3
Pass rate of each item at 18 months of age.

Item no.	M-CHAT item	Autism spectrum disorders (ASD) $(n = 20)$	Reference (n = 1167)	Corrected
1.	Enjoys being swung	95.0%	99.9%	.769
2.	Interest in other children	100.0%	99.7%	1.000
3.	Climbs up stairs	100.0%	99.6%	1.000
4.	Enjoys peek-a-boo	100.0%	99.7%	1.000
5.	Pretend play	75.0%	97.9%	.002**
6.	Imperative pointing	80.0%	98.6%	.006**
7.	Declarative pointing	70.0%	97.9%	.000**
8.	Functional play	70.0%	93.7%	.030*
9.	Brings object to show	75.0%	98.2%	.001**
10.	Eye contact	90.0%	99.1%	.433
11.	Oversensitive to noise	95.0%	83.2%	1.000
12.	Responds to smile	100.0%	99.9%	1.000
13.	Imitation of action	80.0%	99.1%	.001**
14.	Responds to name	85.0%	99.8%	.001**
15.	Point following	85.0%	99.1%	.024*
16.	Walking	100.0%	99.7%	1.000
17.	Gaze-following	65.0%	95.4%	.001**
18.	Unusual finger movement	95.0%	97.9%	1.000
19.	Gaining parent's attention	85.0%	97.7%	.286
20.	Wondering hearing	95.0%	100.0%	.388
21.	Understands what is said	85.0%	99.1%	.030*
22.	Stares at nothing	90.0%	91.7%	1.000
23.	Social reference	60.0%	94.7%	.000**

Fisher's exact test with Bonferroni's correction, p < .05.

Analysis of variance (ANOVA), p < .01.

Table 4
Cut-off value and related measures for predicting later diagnosis of autism spectrum disorders (ASD) in children at 18 months of age.

	Sensitivity	Specificity	Positive predictive value	Negative predictive value
Full version (	M-CHAT score)			
≥1	.800	.609	.034	.994
≥2	.750	.893	.107	.995
≥3	.550	.961	.193	.992
≥4	.350	.978	.212	.989
Short version	(M-CHAT-SV score)			
≥1	.650	.885	.088	.993
≥2	.550	.958	.183	.992
≥3	.400	.981	.267	.990
≥4	.200	.990	.250	.986

Note: Children with ASD, n = 20; reference, n = 1167. Short version of the Japanese M-CHAT consists of 9 items (items 5, 6, 7, 9, 13, 15, 17, 21, and 23).

### 3.4. Screening utility (preliminary report)

Table 4 shows the cut-off value and related measures for predicting later diagnosis of ASD in children at 18 months of age. In order to maintain a balance between sensitivity and specificity, failing 2/23 items and 1/9 items was thought to provide the best cut-off; however, low positive predictive values indicated that there were many false positives.

### 4. Discussion

In this study, we demonstrated the reliability and validity of the Japanese version of the M-CHAT, identified its critical items in order to produce a short version, and reported preliminary information regarding the utility of the full/short version as a screening tool for early detection of ASD.

To the best of our knowledge, this is the first study to confirm that the M-CHAT has good inter-rater and test-retest reliabilities for infants and toddlers in the first 2 years of life. The present results suggest that caregivers can provide reliable answers to most of the items on the Japanese version of the M-CHAT. As to the test-retest reliability, a relatively short interval between tests was used in this study because children at this age develop very rapidly, which might have caused memory effect and *false* consistency. Future studies should be conducted to confirm the reliability of the M-CHAT for toddlers with an ASD diagnoses.

The Japanese version of the M-CHAT showed concurrent validity with the CARS, which has been reported to be useful for evaluating autistic symptoms in toddlers 2 years of age (Chlebowski et al., in press). In addition, the M-CHAT score was high in children later diagnosed with ASD, particularly in children with autism. These results reflect the high validity of the Japanese M-CHAT and support its use as a ASD screening tool for toddlers. Further replication studies are necessary using other standardized measurements such as the Autism Diagnostic Observation Schedule (ADOS) (Lord et al., 2000) and/or Autism Diagnostic Interview-Revised (ADI-R) (Lord, Rutter, & Le Couteur, 1994).

Although the children who were evaluated in the present study were slightly younger than those administered the original M-CHAT (Kleinman et al., 2008; Robins et al., 2001), children 18 months of age who are later diagnosed with ASD showed lower pass rates on 5 out of the 6 "critical items" of the original M-CHAT (items 2, 7, 9, 13, 14, and 15) except item 2 (Interest in other children). Item 2 was originally included in the CHAT (Baron-Cohen et al., 1992) and was also found to be sensitive to diagnosis of ASD when asked retrospectively to Japanese caregivers in clinical settings (Koyama et al., in press). Its high pass rate (100%) in children with ASD in this study indicates that caregivers were not aware of the child's disinterest toward other children. At the age of 18 months, Japanese caregivers may still interpret a lack of interest in other children as modesty or shyness rather than a symptom.

Item 14 (Responds to name) is one of the "critical items" on the original M-CHAT (Kleinman et al., 2008; Robins et al., 2001) and similarly showed significant discriminative power for the 18-month-old children in the present study. However, as neither inter-rater nor test-retest reliability were confirmed, we conservatively excluded this item from the short version. Four children were rated as "pass" by their mother but were rated "fail" by their father, while 3 children were rated conversely. Two children who had first failed this item were rated as "pass" the second time by their mother, whereas one child was rated conversely. These inconsistencies were found in older children as well as younger children. Although this item (Does your child respond to his/her name when you call?) was intended to ask about children's ability to respond to caregivers' social approach by calling, it may be ambiguous for parents of children in the age range of this study. Since the item had good validity for the early detection of ASD in our study, it appears that rephrasing for clarity is necessary.

We created a 9-item short version of the Japanese M-CHAT with higher internal consistency and similar validity compared to the full version. This simpler screening tool may have a practical advantage. As there are no reverse items (items 11, 18, 20, and 22) in the short version, the practical load on the check-up staff may be reduced. Although a single screening using parent-report questionnaires can never predict an ASD diagnosis precisely (Bryson, Rogers, & Fombonne, 2003), it can

be the first step in the multistage screening process as health check-ups provide valuable opportunities for screening young children aged 18-24 months.

The low positive predictive value observed in our epidemiological community sample was consistent with the findings of the original M-CHAT for the low-risk sample (Kleinman et al., 2008), indicating many false positives were involved in the early screening process using the M-CHAT. Kleinman et al. (2008) reported that the positive predictive value was improved from .11 to .65 by adding a follow-up telephone interview. Considering that multistage screenings are necessary for effective early detection of ASD (Bryson et al., 2003), the Japanese version of the M-CHAT, which has acceptable sensitivity, may be appropriate for a first-level screening targeting the general population.

Although only a preliminary finding, in terms of the sensitivity and specificity, the best cut-off value for the Japanese version of the M-CHAT was failure on 2/23 items, which is lower than failure on 3/23 items using in the original M-CHAT (Robins et al., 2001). This was due to the higher pass rate in Japanese 18-month-old children who are later diagnosed with ASD compared with their American counterparts (Robins et al., 2001). The 6-month difference in age between the Japanese and American children may in part explain this inconsistency. Alternatively, it may be explained by the tendency of Japanese people to avoid making definite answers, even if they have some concerns. Thus, the yes-no questions of the M-CHAT may not have accurately reflected the Japanese caregivers' observations. Modifying the response style to a Likert scale, as used in the Chinese version (Wong et al., 2004), might produce different results.

Due to the small sample size, we should avoid over generalizing the findings of this study, especially the cut-off score. As for use, some modifications may be necessary according to the age of children and cultural attitudes and values. Furthermore, long-term follow-up is necessary because the current scores could change if new ASD cases are detected in the future. Nonetheless, based on the present results, the Japanese version of the M-CHAT, as well as its short version, are considered reliable and valid for use as a first-level screening tool for early detection of ASD in the general population.

### Acknowledgements

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## 早期発見をめぐって スペクトラム障害

かみお・ようこ

**(3)** はじめに き始めています。 やく社会全体が危機感を持って真剣に取り組もうと動 発達障害」 のある子どもへの対応について、よう

しました。幼児期の早期発見と早期介入は自閉症対策 る全国調査(エコチル調査)に向けて、それぞれ動き出 もの健康に影響を及ぼす遺伝と環境の相互作用を調べ 特別支援教育を、そして環境省は発達障害を含む子ど 厚生労働省は発達障害者支援事業を、文部科学省は

> が始まっています。 の要であり、 括する広い臨床単位) ラム障害 定型自閉症、 (重度の自閉症から軽度のアスペルガー症候群や非 特定不能の広汎性発達障害などの下位群までを包 わが国を含む世界各国で自閉症スペクト の早期発見と早期介入の取り組み

応に、 中心に述べます。 は、 も多様な困難を持ちやすい傾向がみられる一方で、 人にはない素晴らしい回復力も秘めています。 自閉症的特徴の強い子どもは、 自閉症スペクトラム障害のある子どもたちへの 今求められる理解と課題について、乳幼児期を 対人以外の発達 本稿 面 対

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和書店、二〇〇九年)、『発達障害の臨床心理

より現職。著書に『自閉症』(分担執筆、 大学院人間環境学研究院を経て、二〇〇六年 卒業。京都大学医学部精神神経科、

博士。専門は児童精神医学。京都大学医学部 究所児童・思春期精神保健研究部部長。 国立精神・神経医療研究センター精神保健

学」(共著、東京大学出版会、二〇一〇年)、

医学書院、二〇一〇年

## 撃早期診断の時期とは?

歳、 は社会に出てから自ら気づき診断を求めるケースも多 には就学後に初めて周囲に気づかれる場合や、 れの有無によって幅があります。技術上は、早くて一 クトラム障害という診断でも、 のが実情です。 ここで「早期」としている時期は、 遅くても三歳で診断可能と考えられますが 全般的発達や言 同じ自閉 なかに 症スペ 語 実際 0 遅

される高機能タイプでは、 動を伴わないおとなしいタイプの子どもは見逃され 自閉症やことばをよく話すアスペルガー症候群 が一歳六カ月健診で発見されているわけではなく、多 とはいっても必ずしもすべての典型的自閉症の子ども 動特徴を示すので、たいてい一歳代で気づかれます。 心が乏しく感覚遊びに没頭する、多動など、顕著な行 最も早く診断可 スもありますが、三歳になるまでは判断に迷うこ 傾向にあります。 言語の遅れや呼びかけても反応がない、人への関 能 なのは、 一方、知能に遅れがない 一歳六カ月で診断 知的障害を伴う自 高機能 に代表 可 能な 閉 80 症

とも少なくありません。

ほしいものです。

このように、自閉症スペクトラム障害の診断時期にこのように、自閉症スペクトラム障害の診断時期にこのように、自閉症スペクトラムの特性を把握けつなら、遅くとも三歳健診では子どもの特性を把握ような軽度群であっても自閉症スペクトラムの特性をような軽度群であっても自閉症スペクトラムの特性を出るかなら、遅くとも三歳健診では子どもの特性を把握けつなら、遅くとも三歳健診では子どもの特性を把握けつなら、遅くとも三歳健診では子どもの特性を把握けつなら、遅くとも三歳健診では子どもの特性を把握けつなら、遅くとも三歳健診では子どもの特性を把握けつなら、遅くとも三歳健診では子どもの生活しにくさいます。

もあれば、支援のない状況が長く続いてなかなか回復可能です。これも個人差があり、回復力の高いケースるのは事実ですが)がなされるように支援できれば、それは遅し、的確な対応(それは遅くなればなるほど難しくないができるだけ速やかに発達段階に応じたニーズを時点でできるだけ速やかに発達段階に応じたニーズをで、必ずしも悲観する必要はありません。気づかれたなますが、就学前に支援がなされなかったからといっただし、早いほうが望ましいという理由は以下で述

への糸口が見つからないケースがあるのは事実です。

# **◎早期診断・早期支援はなぜ重要なのか**

なります。 かった、子どもの優れた特徴の発見も同時にできれ たものになります。 もより納得のゆく、本当の意味で家族のニーズに沿っ 確に評価したうえでの家族への説明は、 育サービスの紹介、 セスメントで得られる特性把握は、育児への助言、 活をサポートしていくこと、にあります。子どものア な目的は、 要な重要な情報です。一人ひとりの子どもの特性を的 ている能力を最大限に発揮しすこやかに成長できる生 診断は手段であって、 意欲を引き出す豊かな環境を整える際の助けにも 子どもが安定した気持ちで一人ひとり持っ 幼稚園や保育所での支援などに必 また、日々のなかでは気づかな 目的ではありません。 家族にとって 最終的 療

の療育プログラムでは短期的な効果も報告されていま重要な手段です。米国からは、自閉症に専門的な内容ない現在、早期療育は子どもの社会的発達を促進する自閉症スペクトラム障害に対する生物学的治療法が

児への動機づけを高めることが示唆されています。の一部ではありますが、こだわりを軽減し、母親の育兼ねて早期から母子へ支援することで、子どもの行動不十分と言わざるをえませんが、それでも育児支援もす。日本の公的地域サービスはまだ質的にも量的にも

関係にはない、ということです。 短期効果だけでなく、成人後の生きがいやQOL短期効果だけでなく、成人後の生きがいやQOL短期効果だけでなく、成人後の生きがいやQOL短期効果だけでなく、成人後の生きがいやQOL短期効果だけでなく、成人後の生きがいやQOL短期が見た。このことは高機能の自閉症スペクトラム障害の成た。このことは高機能の自閉症スペクトラム障害の成まざまな困難を乗り越えていく力が育っていく可能性まざまな困難を乗り越えていく力が育っていく可能性まざまな困難を乗り越えていく力が育っていく可能性まざまな困難を乗り越えていく力が育っていく可能性まざまな困難を乗り越えていく力が育っていく可能性を示唆しているのではないでしょうか。つまり、自閉を示唆しているのではないでしょうか。つまり、自閉を示唆しているのではないでしょうか。つまり、自閉を示唆しているのではないでしょうか。

暮らす自閉症スペクトラムの子どもの大半は、医療機るリスクが大きいのです。最近の研究からは、地域でタルな問題の頻度が高く、そのためにQOLが低下す度の場合には、むしろ合併するうつや不安などのメンアスペルガー症候群のように自閉症状それ自体が軽

51 自閉症スペクトラム障害の早期発見をめぐって

関にかかっているかどうかにかかわらず、 ども自身が、大きな困難でも周りの人々に助けても 校での支援が迅速かつ適切に行われ、 メンタルな問題も合併していることがわ らって乗り越えられたという自信を得て、 問題も合併したアスペルガー症候群のケースでも、 いうことを知りました。そして一番大事なことは、 スムーズになされた場合には、子どもの回復は早いと く成長していったということです。 私たちはこうした調査をするなかで、メンタルな 医療との連携が かってい さらに大き なんらか 子 学 ま 0)

## 地域での早期発見の実際

いないこと、児童精神科医の数が決定的に不足してい害の子どもの発達について十分な知識や経験を持っては、健診に携わる保健師や小児科医が必ずしも発達障れています。これを活用した自閉症スペクトラム障害れています。これを活用した自閉症スペクトラム障害れています。これを活用した自閉症スペクトラム障害かが国には乳幼児健診という世界に誇るべき制度が

り、 どといった多くの問題があります。多動やかんしゃく 機が長すぎて、サービスを受けることができない 多職種の専門家が全国的には乏しいことも関係して ること、 せることになります。 言を受けていない母親の不安を高め、 人遊びに没頭し母親を無視する自閉症状は、 が強いと外出もままならず、室内で遊ばせていても一 時間が過ぎてしまったり、あるいは療育サービスの待 ます。このため現状では、 評価する言語療法士、 懸念を持たれた場合でも診断に至らずいたずらに があると思われます。さらに発達を多面: 臨床心理士、作業療法士 早期兆候が見逃されてい ストレスを募ら 診断や助 的

リー ŧ 症 クリーニングは、慣れれば難しいものではなく、 ングを加えるべきと考えます。自閉症の早期兆候 が推奨するように、 することは難しいという現実からも、米国小児科学会 の遅れのない自閉症スペクトラム障害の子どもを発見 児以外の社会的 既存の全般的な発達スクリーニングだけでは、 ニング尺度が開発されていますが、本稿では世界 有用と思われます。こうした目的で多数のスク 発達に困難を持つ子どもに対 自閉症の発見に有効なスクリーニ のス 知

### 日本語版 M-CHAT (The Japanese version of the M-CHAT)

お手さんの直弧のご様子について、もっとし質能にあてはまるものを〇で始んでください。ナベマの質能に二面をくださるおにお願い。小にます。もし、質能の行動をかったにしないと思われる場合は伏とせば、1.2歳しか見た覚えがない、お手さんはそのような行動をしない(れい)と、を選ぶように)とご適なください。知言から、17.23 についてはらさご参考べたさい。

1.	お子さんをブランニのように終らしたり、ひざの上で紹するとなってますか?	他们的说
2	他の手どもに関係がありますか?	はい・リャル
3.	強敵など、何かの上に違い上がることが好きですか?	1411111111111
4	イナイイナイバーをすると翼びますか?	はいりは火
S.	電流の発達を楽しかててしゃべるまれをしたり、人形やその他のモノを使ってごっこ 速びをしますか?	はい・ロス
6	のでは、レンモノがある時、指をさして要求しませか。?	はいいえ
7.	何かに実践を持った時、指をさして伝えようとしますか?	はい・いは
<b>1</b> .	クルマキ種者などのオモチャを、首に入れたり、さかったり、落としたりする遊びではなく、 オモチャに含った遊び答をしますか?	はい・以東
9.	あなたに見てほしいモノがある時、それを見せに持ってきますか?	はい・以後
<del>10</del> .	1, 20上9長く、おなたの音を見つめますか?	はい・以次
11.	ある様の行に、とくに返数に見なして不軽値になりますか?(耳をふさくなど)	はいいりん
12	かなたがお子さんの頭をみたり、美いかけると、美術を近してきますか?	1811-14次
13.	おなたのすることをまねしますか?(たとえば、ロをといらせてみせると、前またをしようとしますか?)	はいいえ
14	おなたがも前を呼ぶた。反応しますか?	はい・いれた
15.	あがたが家園の中の離れたところにあるオモチャを指でさすと、お手さんはそのが確定 見ますか?	はい・いい
16.	おする人は最後ますか?	1211-14
17.	あなたが見ているそとを、お子さんも一格に見ますが?	12.1.14
18.	前の近くて倒をひらひら動かすなどの変わった噂がありますか?	12t1 - 64
19.	わなたの注意を、自分の方にひこうとしますか?	(#14 + 14)
20.	おうさんの質で描こえないのではないかとこれをおれたことがありますが、?	1211-14
21.	<b>言われたことはをわかっていますか?</b>	(\$65 - 63)
22	例とない音をに、一っと見つめたり、自然なびたすらうろうろすることがよりますか?	1269 - 64
23.	いつしと違うことがある時、あなたの数を見て見りを充ちめますか?	12 4

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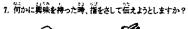
MCHAT の影響はDisce Robins, Deboth Feis, Mariente Baron こかります。この日本意思は、独立精神・神秘センター科神(M) 所見象・思る時間神(株式正成の体では大きな大きな呼が) およっておいままだけられることのです。

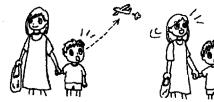
### 図1 日本語版M-CHAT

(The Japanese version of the M-CHAT)

日本語版M-CHATは、独立行政法人国立精神・神経医療研究センターの児童・思春期精神保健研究部のホームページからPDFをダウンロードすることができます。

http://www.ncnp.go.jp/nimh/jidou/research/mchat.pdf

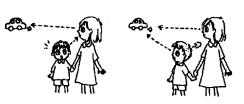




9. あなたに見てほしいモノがある時、それを見せに持ってきますか?



17. あなたが見ているモノを、お子さんも一緒に見ますか?



23. いつもと違うことがある時、あなたの顔を見て反応を確かめますか?



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### 53 自閉症スペクトラム障害の早期発見をめぐって

います。 for Autism in Toddlers)の日本語版をご紹介したいと思各国で広く用いられているM-CHAT(Modified Checklist

ncnp.go.jp/nimh/jidou/research/mchat.pdf)。 日本語版M-CHATは、二十三項目から成る親記入式日本語版M-CHATは、二十三項目から成る親記入式日本語版M-CHATは、二十三項目から成る親記入式日本語版M-CHATは、二十三項目から成る親記入式

六カ月から二歳までという月齢で用いることに意味がたか月から二歳までという月齢で用いることに意味がたんどの子どもが通過します。こうした社会的発達をないまま覚える言語は、高機能自閉症児の膨大な語彙ないまま覚える言語は、高機能自閉症児の膨大な語彙ないまま覚える言語は、高機能自閉症児の膨大な語彙ないまま覚える言語は、高機能自閉症児の膨大な語彙ないまま覚える言語は、高機能自閉症児の膨大な語彙ないまま覚える言語は、高機能自閉症児の膨大な語彙ないまま覚える言語は、高機能自閉症児の膨大な語彙ないまま覚える言語は、高機能自閉症児の膨大な語彙ないまま覚える言語は、高機能自閉症児の膨大な語彙ないまま覚える言語は、平均的な子どもではゼロ歳から芽生の社会的行動は、平均的な子どもではゼロ歳から芽生の社会的行動は、平均的な子どもではゼロ歳から芽生の社会的行動は、平均的な子どもではゼロ歳から芽生の社会の行動は、平均的な子どもではゼロ歳から芽生の社会の行動は、平均的な子どもではゼロ歳から芽生の社会の行動は、

とが発達上の問題として把握されるのです。あり、その時点でこれらに通過できていないというこ

ニーズの質と量をより詳細に判断するのがよいと思わ 進めていきます。面接や親子グループに誘って、 ケースについては親のニーズを発見し寄りそう方向で す。基本的には、一回だけで判断をせず、継続的 認を行うなど、工夫の余地もいろいろあると思い 項目については健診場面で保健師あるいは心理士 げるためには、親回答に加えて、短時間で観察可能 エーションを経験しています。また、さらに精度を上 場合、二歳時の相談で行う場合など、 タイミングや方法は、一歳六カ月の集団健診で用い 話フォローを行うなど経過をモニターしながら、 自治体も少しずつですが、増えてきています。 ており、最近では全部あるいは一部を健診に活用する これまで私たちは、地域によってスクリーニングの M-CHATは、このように日本でも有用性が確認され いくつかのバ が確 ま IJ

健診でM-CHATで陽性となるケースを数年追跡した結が、長期の追跡による検証が待たれます。一歳六カ月どの方法が優れているかは一概に比較できません

れます。