

Figure 1. Frequencies of CKD according to sex and age.

were older, had higher prevalence of hypertension and hypercholesterolemia, and had a lower frequency of current drinking than those without CKD (Table 1).

During an average 11.7-year follow-up period, we documented 213 strokes and 133 MIs. In men and women combined, compared with subjects for $GFR \geq 90$ mL/min/1.73m² the multivariable HRs (95% confidence intervals; CIs) for CVD incidence were 1.75 (1.22 to 2.50) in $GFR = 50$ to 59 mL/min/1.73m² and 2.48 (1.56 to 3.94) in <50 mL/min/1.73m² (Table 2). In addition, the risks of CVD for each GFR category in men and women separately were similar to the risks for all participants. The multivariable HR (95% CIs) of CVD incidence for CKD was 1.70 (1.30 to 2.23) in all subjects (data not shown).

In Table 3, the multivariable HRs (95% CIs) for strokes were 1.94 (1.26 to 2.98) in the $GFR = 50$ to 59 mL/min/1.73m² and 2.19 (1.18 to 4.06) in the $GFR < 50$ mL/min/1.73m² compared with subjects for $GFR \geq 90$ mL/min/1.73m². Results for cerebral infarction were similar to strokes. Age-adjusted HRs (95% CIs) for intracerebral hemorrhage were 1.93 (0.77 to 4.85) in the $GFR = 50$ to 59 mL/min/1.73m² and 2.52 (0.72 to 8.80) in the $GFR < 50$ mL/min/1.73m² (supplemental Table I, available online at <http://stroke.ahajournals.org>).

In Figure 2, compared with the optimal BP subjects without CKD, the normal BP, high-normal BP, and hypertensive subjects without CKD showed increased risks of CVD, whereas the impact of each BP category on CVD was more evident in subjects with CKD (probability values for interaction between CKD and BP category were 0.04 in men, 0.49 in women, and 0.06 in all subjects). Results of stroke were similar (probability values for the interaction were 0.03 in men and 0.90 in women, data not shown). Supplemental Table II shows the hazard ratios for the association between 10 mm Hg of SBP and the risk of CVD in subjects with or without CKD.

Using the HRs, we estimated the population attributable fraction of CVD to exposure for CKD at baseline by sex. We found that 8.3% in men and 17.6% in men with CVD incidences could be described as excessive incidence attributable to CKD.

Discussion

In this cohort study of a general urban Japanese population, CKD was a risk factor for CVD and its subtypes. A stronger association between BP and the incidence of CVD was

Table 1. Baseline Characteristics of Study Subjects According to Chronic Kidney Disease

Variables	Men			Women		
	CKD (-)	CKD (+)	P Value	CKD (-)	CKD (+)	P Value
No. of subjects	2341	229		2593	331	
Age at baseline, y	55±13	61±12	<0.001	53±13	62±12	<0.001
Body mass index, kg/m ²	22±3	23±3	<0.001	22±3	22±3	0.332
Blood pressure category, %			0.005			<0.001
Optimal	31.7	24.0		43.9	27.2	
Normal	19.2	14.4		16.6	15.4	
High-normal blood pressure	16.2	20.5		14.0	14.8	
Hypertension	32.9	41.1		25.5	42.6	
Present illness, %*						
Hypercholesterolemia	28.1	35.8	0.014	40.7	54.7	<0.001
Diabetes	6.1	6.6	0.791	3.2	5.4	0.036
Smoking status, %			0.007			0.713
Current	51	42		12	12	
Quit	30	40		4	4	
Never	19	18		84	83	
Drinking status, %			0.024			0.017
Current	76	68		34	26	
Quit	3	6		2	3	
Never	21	26		65	71	

*Hypercholesterolemia; antilipidemic drug use or total cholesterol ≥ 5.7 mmol/L (220 mg/dl), diabetes; antihyperglycemic drug use or fasting blood sugar ≥ 7.0 mmol/L (126 mg/dl). Plus-minus values are means±SD.

Table 2. Age and Multivariable Adjusted Hazard Ratios (95% CIs) for Incidence of Cardiovascular Disease† According to Category of Glomerular Filtration Rate by Sex

Variables	Glomerular Filtration Rate, ml/min/1.73m ²				P for Trend
	≥90	60 to 89	50 to 59	<50	
Men and Women					
Cases, n	94	176	51	25	
Person-years	28 736	29 336	4764	1558	
Age-adjusted	1	1.22 (0.94–1.58)	1.71 (1.20–2.42)	2.49 (1.59–3.90)	<0.001
Multivariable adjusted*	1	1.21 (0.93–1.58)	1.75 (1.22–2.50)	2.48 (1.56–3.94)	<0.001
Men					
Cases, n	50	124	24	11	
Person-years	12 092	14 835	1928	522	
Age-adjusted	1	1.20 (0.85–1.70)	1.63 (1.00–2.68)	2.17 (1.11–4.23)	0.008
Multivariable adjusted*	1	1.21 (0.85–1.70)	1.78 (1.08–2.94)	2.38 (1.21–4.68)	0.004
Women					
Cases, n	44	52	27	14	
Person-years	16 644	14 502	2836	1036	
Age-adjusted	1	1.22 (0.81–1.83)	1.79 (1.09–2.92)	2.81 (1.53–5.18)	<0.001
Multivariable adjusted*	1	1.21 (0.80–1.84)	1.76 (1.05–2.93)	2.31 (1.20–4.43)	0.002

*Multivariable adjusted for age, BMI, smoking, drinking, and present illness (hypertension, diabetes, and hypercholesterolemia).
†Cardiovascular disease includes both stroke and MI.

observed in the presence of CKD. Furthermore, we found that 8% in men and 18% in women of CVD incidence may be derived from CKD cases.

Go et al reported that both severe and moderate renal diseases were risk factors for CVD incidence.⁶ A pooled analysis of community-based studies demonstrated that CKD is an independent risk factor for the composite of all-cause mortality in blacks and whites and CVD incidence in blacks.⁵ In contrast, NHANES I did not provide relationships between mortality and moderately higher serum creatinine levels.⁴ The Framingham Heart Study and Offspring cohorts have shown no significant association between the presence of kidney disease and CVD incidence.³

The results of our study are essentially compatible with previous cohort studies in Japan. The Hisayama study demonstrated that CKD was a risk factor for incidence of coronary heart disease in men and ischemic stroke in women.⁸ The Ohasama study indicated that decreased kidney function increased the risk of first symptomatic stroke events.¹⁹ This study used creatinine clearance rather than estimated GFR. Irie et al showed that subjects with GFR <60 had a higher risk of CVD mortality⁷ but did not examine the risk of GFR 50 to 59 mL/min/1.73m². The NIPPON DATA 90 indicated that CKD was an independent risk factor for cardiovascular death in a community-dwelling Japanese population.²⁰ The end point of these studies was also mortality. Ninomiya et al

Table 3. Age-Sex and Multivariable Adjusted Hazard Ratios (95% CIs) for Incidence of All Strokes, Cerebral Infarction, and Myocardial Infarction According to Category of Glomerular Filtration Rate

Variables	Glomerular Filtration Rate, ml/min/1.73m ²				P for Trend
	≥90	60 to 89	50 to 59	<50	
Person-years	28 258	28 690	4528	1446	
All strokes					
Cases, n	65	99	36	13	
Age and sex adjusted	1	1.02 (0.73–1.41)	1.78 (1.17–2.70)	1.93 (1.05–3.54)	0.004
Multivariable adjusted*	1	1.04 (0.74–1.45)	1.94 (1.26–2.98)	2.19 (1.18–4.06)	<0.001
Cerebral infarction					
Cases, n	42	66	24	9	
Age and sex adjusted	1	0.99 (0.66–1.49)	1.72 (1.03–4.19)	2.01 (0.97–4.19)	0.020
Multivariable adjusted*	1	0.98 (0.65–1.49)	1.81 (1.07–3.07)	2.26 (1.07–4.78)	0.008
Myocardial infarction					
Cases, n	29	77	15	12	
Age and sex adjusted	1	1.68 (1.08–2.61)	1.64 (0.87–3.09)	4.26 (2.14–8.45)	<0.001
Multivariable adjusted*	1	1.60 (1.03–2.49)	1.51 (0.80–2.88)	3.56 (1.73–7.30)	0.002

*Multivariable adjusted for age, sex, BMI, smoking, drinking, and present illness (hypertension, diabetes, and hypercholesterolemia).

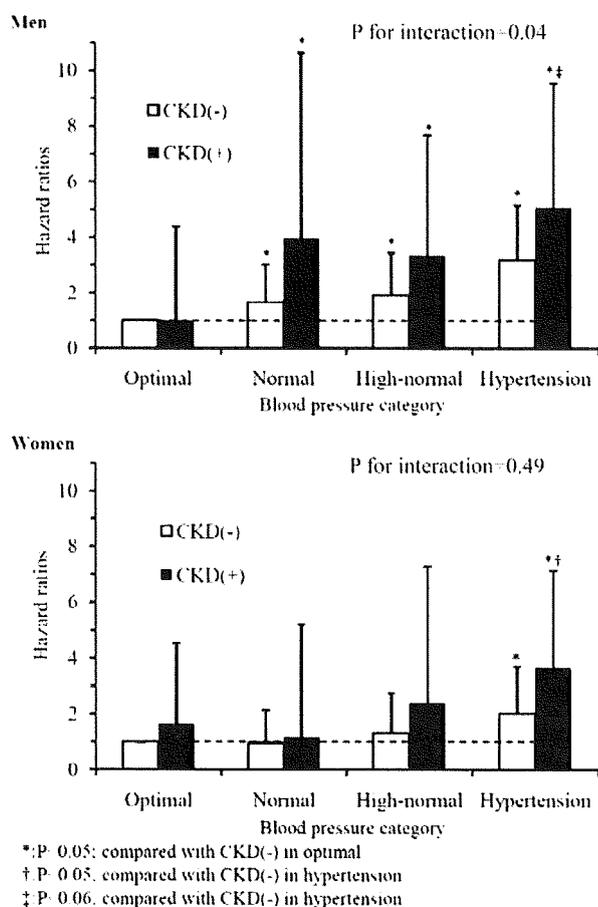


Figure 2. The combination of CKD and BP categories on multivariable hazard ratios for CVD. Data for men and women are presented separately. Multivariable analyses are adjusted age in 5-year increments as stratified variables and other potential confounding factors of hypercholesterolemia, diabetes, and smoking and drinking status.

has recently reported that CKD was risk factors for CVD and stroke in women and that CKD increased the association between BP category and CVD in all subjects from 10 combined different cohort studies using different methods of creatinine measurement.¹⁰ All of our samples were measured using the same analyzer at one laboratory.

Compared with the previous studies, our study has several methodological strengths. First, we could perform subanalysis by age and CVD subtype, because we evaluated a large cohort of participants. Second, each participant's health status was checked during a clinical visit at the National Cardiovascular Center every 2 years. In addition, each year, a health questionnaire was given to each participant via mail or telephone. We could evaluate the registry of CVD incidence with the data obtained from clinical visits, annual questionnaires, or death certificates. Finally, our cohort population was selected at random from an urban population, in contrast to most other cohort studies in Japan, which have relied on rural populations.^{7,8,19}

There may be some reasons why CKD is more positively associated with CVD in blacks or Japanese than in whites. Blacks and Japanese are more likely to have hypertension at

an earlier age.^{9,21} Therefore, the period of hypertension exposure tends to be longer in blacks and Japanese than in whites. The GFR estimation has been adjusted by a factor suitable for Japanese populations.¹⁵

Reduced kidney function is associated with increased levels of inflammatory factors,^{22,23} abnormal apolipoprotein levels,²² elevated plasma homocysteine,²² enhanced coagulability,²³ anemia, left ventricular hypertrophy, increased arterial calcification, endothelial dysfunction, and arterial stiffness.^{2,24} How these and other factors interact to increase the risk of adverse outcomes remains unclear but is the focus of ongoing investigations.²⁴

Subjects with GFR levels of 50 to 59 mL/min/1.73m² were observed to be at risk for stroke. It is desirable to prevent CVD in subjects with both high-risk (<50 mL/min/1.73m²) and less severe kidney disease (50 to 59 mL/min/1.73m²), although an accelerated decline in GFR occurred for the subjects whose initial GFR <50 mL/min/1.73m².²⁵

Hypertension is a strong risk factor for early decline in kidney function; hypertensive patients (BP \geq 160/95 mm Hg) have a 5-fold greater decline in GFR (2.7 mL/min/1.73m²/yr) compared with patients with BP <140/90 mm Hg.²⁶ Furthermore, in this study, the association between BP and the incidence of CVD were evident by CKD. The risk of CVD was higher in CKD subjects with normal and high-normal BP than in non-CKD subjects in the same BP categories. Using the combination of BP and CKD, it could be possible to screen more efficiently for higher risk of stroke and MI. This is compatible with the CKD clinical guidelines, which state that the preferable BP for subjects with CKD is 130/80 mm Hg.²⁷ For the prevention of CVD incidence for all hypertensive subjects in health check-ups, it might be desirable to measure serum creatinine levels and to intervene in lifestyle modification such as reducing salt intake, more frequent exercise, or quit smoking.

Our study has several limitations. The primary limitation is dilution bias,²⁸ in that the current study was based on single-day measurement of creatinine levels. The creatinine levels might have been misclassified, despite the fact that measurements of creatinine levels on a single day have been found to be accurate in other epidemiological studies. Second, we did not perform a creatinine clearance test or 2 measurements of serum creatinine at least 3 months apart. Although our definition of CKD is based on a single assessment of serum creatinine, the equation provides an accurate estimated GFR value.¹⁵ Third, even with the moderate sample size (n=5494) and 12-year duration, the numbers of end points were limited, especially when the data were stratified by 2 variables, such as sex and glomerular filtration rates. A study with more participants with the same protocol is required to validate to the association between BP category and CVD by CKD.

In conclusion, CKD was associated with an increased risk for stroke and MI in a general urban Japanese population. Furthermore, the association between BP and CVD may be evident by CKD. To prevent the incidence of stroke and MI, it is necessary for subjects with CKD to control their BP by lifestyle modification and proper clinical treatment.

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Disclosures

None.

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CARDIOVASCULAR DISEASES

Elevated serum creatine kinase predicts first-ever myocardial infarction: a 12-year population-based cohort study in Japan, the Suita study

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Background In myocardial infarction (MI), it is well known that serum creatine kinase (s-CK) increases after onset, but it is unclear whether s-CK elevates before MI onset. The present analysis examined whether elevated s-CK levels predicted first-ever MI or stroke.

Methods This study was a population-based cohort study in a Japanese urban area. Study subjects were comprised of 5026 initially healthy Japanese (2370 men and 2656 women, mean age: 54.5 years) without a history of MI or stroke. They were followed-up for 11.8 years on average, and 103 MIs (definite: 45; probable: 58) and 168 strokes (definite: 126; probable: 42) were observed. There was no subject who developed MI just at baseline (the follow-up period among those with definite MI was, at earliest, 0.20 years).

Results The adjusted hazard ratio for definite MI was 4.18 (95% confidence interval 1.66–10.53) with s-CK levels of ≥ 200 IU/l, compared with the reference category (s-CK levels of ≤ 99 IU/l), whereas no relationship was observed between s-CK levels and the risk for stroke. With regard to definite MI, an interaction between s-CK levels and dyslipidaemia was observed. Among subjects with hypercholesterolaemia, the hazard ratio linearly elevated with increased s-CK levels. On the other hand, no linear elevation was observed among subjects without hypercholesterolaemia (P for interaction = 0.011).

Conclusions The present study suggested that screening for elevated s-CK levels in initially healthy Japanese subjects was useful to predict first-ever MI in the future, especially in subjects with dyslipidaemia.

Keywords Creatine kinase, myocardial infarction, stroke, cohort studies, urban population

Introduction

Creatine kinase is included mainly in skeletal muscle, cardiac muscle and brain, and is involved in energy production. Serum creatine kinase (s-CK) levels in healthy individuals are influenced by age, sex, race, pregnancy, muscle mass and physical activity, and is well known to be elevated following any damage to or disease of the above-mentioned organs.¹

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In myocardial infarction (MI), s-CK levels usually increase within 4–8 h after MI onset and return to normal 48–72 h later.² It is unclear, however, whether s-CK elevation occurs before MI onset. If s-CK elevates before MI onset, s-CK elevation might predict MI incidence, and three possible hypotheses for background mechanism can be considered. First, small amounts of intramyocardial proteins such as creatine kinase might leak into serum by micro-infarction or ischaemia in myocardium.^{3,4} Secondly, s-CK might be a surrogate marker of hypothyroidism, as recent meta-analyses reported the relationship between hypothyroidism and a modest increased risk for coronary heart diseases.^{5,6} Thirdly, medication for hyperlipidaemia that can occasionally cause rhabdomyolysis with s-CK elevation might be involved.^{7,8}

To our knowledge, there has been no previous cohort study that elucidated the association between s-CK levels at baseline and a risk of cardiovascular diseases (CVDs). Therefore, as a first step of testing the above-stated hypotheses, the present analysis examined whether elevated s-CK levels predicted first-ever MIs or strokes in a 12-year population-based cohort study of a Japanese urban area.

Methods

Study design and samples

The Suita study is a population-based cohort study in an urban area located in the mid-west part of Japan.^{9–11} In 1989, the Suita study randomly sampled 12 200 Japanese men and women, aged 30–79 years, from the Suita city residents by sex and age class (10-year increments). From this sample, 6406 men and women participated in a baseline survey (participation rate: 52.5%) at the National Cardiovascular Center from September 1989 to March 1994, and were enrolled as study subjects. They have been examined every 2 years and followed-up to December 2005.

A total of 1380 subjects were excluded from this analysis for the following reasons: past or present history of MI or stroke at baseline ($n=208$); abnormal Q wave on standard 12-lead electrocardiogram (ECG) ($n=3$, corresponding to 1-1 in Minnesota code); missing data for s-CK or any other explanatory variables at baseline ($n=330$); ST depression on baseline electrocardiogram (corresponding to 4-1 in Minnesota code) or angina on treatment ($n=139$); and lost to follow-up ($n=700$). Consequently, 5026 initially healthy subjects (2370 men and 2656 women, mean age: 54.5 years) were included in this analysis. Informed consent was obtained from all subjects. This study was approved by the institutional review board at the National Cardiovascular Center.

Baseline data collection

A baseline survey included questionnaires, anthropometric measurements, a standard 12-lead ECG

and blood sample testing with overnight fasting (≥ 10 h). Height and weight were measured in light clothing, and body mass index was calculated as weight (kilograms) divided by height (square metres). Blood pressure of subjects was measured three times by well-trained physicians in a sitting position after ≥ 5 min of rest, using a standard mercury sphygmomanometer. Blood sample tests included serum total cholesterol, high-density lipoprotein (HDL)-cholesterol, triglycerides, creatinine, creatine kinase and fasting plasma glucose (FPG), and all samples were immediately centrifuged after blood sampling and analysed by an automatic analyser at the laboratory of the National Cardiovascular Center. Creatine kinase was measured in serum with the Japan Society of Clinical Chemistry (JSCC) method.

The standard 12-lead ECGs were double-checked and coded with Minnesota code by well-trained physicians.^{12,13} Physicians or nurses administered questionnaires for life styles and past medical history. Current cigarette use was defined as smoking at least one cigarette a day. Current alcohol consumption was defined as drinking alcohol at least once a week.

Ascertainment of MI or stroke events

Details for the ascertainment of MI or stroke events were previously described elsewhere.^{9–11} To detect MI or stroke events, each subject's health status was checked by physicians or nurses at clinical visits to the National Cardiovascular Center every 2 years. In addition, yearly questionnaires by mail or telephone were also completed for all participants. Those who reported to have developed MI or stroke were confirmed by a review of medical records after informed consent was obtained from each individual. Medical record review was performed by physicians at either the registered hospitals or the administration office of this cohort study. The percentage of informed consents obtained for medical record review was 86.2%.

With regard to MI, definite and probable MI was defined according to the criteria set out by the MONICA project, which requires evidence from ECGs, cardiac enzymes and/or autopsy.¹⁴ Strokes were defined according to the National Survey of Stroke criteria, which require the rapid onset of a series of neurological deficits lasting ≥ 24 h or until death. For each stroke subtype [i.e. cerebral infarction (thrombotic or embolic infarction), intracerebral haemorrhage and subarachnoid haemorrhage], a definite diagnosis was established, based on examination of computed tomographic scans, magnetic resonance images or autopsies.¹⁵

In order to complete the surveillance for fatal stroke and MI, we also conducted a systematic search of death certificates (all death certificates in Japan are forwarded to the Ministry of Health, Welfare and Labour and coded for National Vital Statistics) for the following subjects: (i) those without informed

consent for medical records survey to confirm MI or stroke events that were detected at regular visits to the National Cardiovascular Center or by annual questionnaires; and (ii) those with death certificates who had not been in our register of MI or stroke events. These subjects were defined as probable stroke or MI.

Statistical analysis

s-CK levels were divided into four categories with 100 IU/l increments (i.e. ≤ 99 , 100–199, 200–299, ≥ 300) to compare baseline characteristics and calculate overall and sex-specific crude incidence rates (per 1000 person-years) by the following subtypes of CVD: all MI (including probable and definite MI), definite MI, all strokes, cerebral infarction and haemorrhagic or unclassified strokes. Hypertension was defined as systolic blood pressure ≥ 140 mmHg, diastolic blood pressure ≥ 90 mmHg or use of anti-hypertensive medication. Impaired fasting glucose or diabetes was defined as FPG ≥ 6.1 mmol/l (110 mg/dl) or use of anti-diabetic medication. Hypercholesterolaemia was defined as total cholesterol ≥ 5.7 mmol/l (220 mg/dl) or use of anti-hypercholesterolemic medication. Hypertriglyceridaemia was defined as triglycerides ≥ 1.7 mmol/l (150 mg/dl) or use of anti-hypertriglyceridemic medication. Low HDL-cholesterolaemia was defined as HDL cholesterol ≤ 1.0 mmol/l (39 mg/dl).

s-CK levels of ≥ 200 IU/l were unified into one category (i.e. creating three categories with 100 IU/l increments: ≤ 99 , 100–199 and ≥ 200) to estimate age- and multiple-adjusted hazard ratios (HRs), using a Cox regression model with a reference category of s-CK of ≤ 99 IU/l. This is because the number of subjects and events in s-CK levels of ≥ 300 IU/l was substantially smaller than other categories and because crude incidence rates of s-CK levels of ≥ 300 IU/l were also similar with those of s-CK of 200–299 IU/l. Categories of s-CK levels were entered into models as dummy variables. Men and women were combined in estimating HRs with 95% confidence interval (95% CI) because there was no interaction between sex and s-CK levels. Each model adjusted for the following confounding factors: Model 1 adjusted for age; Model 2 adjusted for age, sex, body mass index, hypertension, impaired fasting glucose or diabetes, hypercholesterolaemia, hypertriglyceridaemia, low HDL-cholesterolaemia, creatinine, current cigarette use and current alcohol consumption. Three additional analyses were conducted in Model 2 for definite MI: one excluded 111 further subjects with 4-2, 4-3 in Minnesota code ($n=4915$, number of MI events=45), another further adjusted for regular physical activity habits (yes or no; 224 subjects were excluded because of missing data, $n=4802$, number of MI events=45) and the other excluded 51 subjects in the 99th percentile or higher of s-CK ($n=4975$, number of MI events=44).

The follow-up period of each participant was divided into four categories (0.20–3.30 years, 3.31–4.99 years, 5.00–8.40 years and 8.41–13.89 years), each of which included an equal number of subjects with definite MI. In each category, the median baseline s-CK levels were compared between subjects with and without definite MI in order to examine whether a specific period of follow-up was involved in the present results.

With regard to definite MI, interactions between s-CK levels and each of several confounding factors (i.e. hypercholesterolaemia, hypertriglyceridaemia, low HDL-cholesterolaemia, hypertension, impaired glucose tolerance or diabetes, current cigarette use and current alcohol consumption) were tested by adding the interaction term to Model 2. HRs by s-CK levels and confounding factors were estimated, adjusting for the same factors as Model 2 except for the stratifying factor of each interaction.

The *P*-value for trend was calculated to test for linearity of HRs. For blood pressure, a third measurement was adopted in all analyses. All *P*-values were two-tailed.

Results

The overall distribution of baseline s-CK levels (minimum, 25th percentile, median, 75th percentile, maximum) for men was 30, 101, 132, 176 and 3547 IU/l, respectively, and for women it was 25, 87, 110, 141 and 1820 IU/l, respectively. There were 1.0% of men and 0.3% of women with s-CK levels of ≥ 500 IU/l.

Baseline characteristics are shown by s-CK levels in Table 1. Regardless of sex, body mass index and creatinine tended to increase with elevation in s-CK levels, and the prevalence of current cigarette smoker decreased with elevation in s-CK levels. In women, higher s-CK levels were associated with an increased prevalence of hypertension, hypercholesterolaemia and hypertriglyceridaemia.

Study subjects were followed for 11.8 years on average, and 103 MIs (definite: 45; probable: 58) and 168 strokes (definite: 126; probable: 42) were observed. Strokes were comprised of 27 intracerebral haemorrhages, 106 cerebral infarctions, 20 subarachnoid haemorrhages and 15 unclassified strokes. The follow-up period among those with definite MI was from 0.20 to 13.89 years.

Median baseline s-CK levels were higher among subjects with definite MI than among subjects without definite MI, regardless of the follow-up period (Figure 1).

Overall and sex-specific crude incidence rates were presented by s-CK levels and CVD subtype (Table 2). Overall, incidence rates of MI or stroke were higher in men than in women, and rates of stroke were higher than MI. Overall or sex-specific crude incidence rates of definite MI tended to be higher in subjects with

Table 1 Baseline characteristics by s-CK levels in a cohort study of a Japanese urban area, 1989–2005

	Men				Women			
	Creatine kinase (IU/l)				Creatine kinase (IU/l)			
	≤99	100–199	200–299	≥300	≤99	100–199	200–299	≥300
Number of subjects	566	1396	306	102	1024	1420	176	36
Age (years)	58.1 (12.4)	54.2 (13.1)	55.4 (13.3)	57.4 (14.4)	51.1 (12.5)	54.8 (12.4)	57.6 (12.3)	58.8 (10.3)
Body mass index (kg/m ²)	22.5 (2.8)	22.9 (2.9)	23.2 (3.1)	22.9 (2.8)	21.8 (3.0)	22.4 (3.3)	22.8 (3.4)	23.2 (2.5)
Hypertension (%) ^a	32.9	32.3	34.6	36.3	21.8	27.6	34.7	50.0
IFG or diabetes (%) ^b	21.2	13.8	8.8	13.7	8.6	8.7	10.2	11.1
Hypercholesterolaemia (%) ^c	30.6	26.8	34.6	26.5	33.2	46.0	48.3	63.9
Hypertriglyceridaemia (%) ^d	38.2	30.7	26.8	28.4	16.4	16.7	19.9	30.6
Low HDL-cholesterolaemia (%) ^e	29.0	24.1	18.6	21.6	9.9	7.8	9.7	5.6
Creatinine (mmol/l)	78.1 (19.7)	80.4 (17.7)	80.9 (17.1)	85.7 (18.7)	58.7 (17.3)	62.2 (15.1)	62.6 (15.6)	63.6 (16.5)
Medication								
Hypertension (%)	11.1	10.0	9.2	12.8	7.8	10.2	11.4	27.8
Diabetes (%)	3.2	2.3	0.0	1.0	1.4	1.1	1.1	0.0
Hypercholesterolaemia (%)	1.6	1.2	1.6	2.9	1.1	2.6	4.6	8.3
Hypertriglyceridaemia (%)	1.2	0.4	0.3	2.0	0.2	0.3	1.1	2.8
Current cigarette use (%) ^f	53.0	51.0	46.1	36.3	15.3	9.7	8.5	5.6
Current alcohol consumption (%) ^g	73.1	75.7	75.2	77.5	35.3	32.0	29.6	30.6

Means in continuous variables are shown with standard deviation in parentheses.

^aHypertension was defined by systolic blood pressure ≥140 mm Hg or diastolic blood pressure ≥90 mm Hg or use of anti-hypertensive medication.

^bIFG or diabetes were defined by FPG ≥6.1 mmol/l (110 mg/dl) or use of anti-diabetic medication.

^cHypercholesterolaemia was defined by total cholesterol ≥5.7 mmol/l (220 mg/dl) or use of anti-hypercholesterolaemic medication.

^dHypertriglyceridaemia was defined by triglycerides ≥1.7 mmol/l (150 mg/dl) or use of anti-hypertriglyceridaemic medication.

^eLow HDL-cholesterolaemia was defined by HDL cholesterol ≤1.0 mmol/l (39 mg/dl).

^fCurrent cigarette use was defined as smoking at least one cigarette a day.

^gCurrent alcohol consumption was defined as drinking alcohol at least once a week.

IFG = impaired fasting glucose.

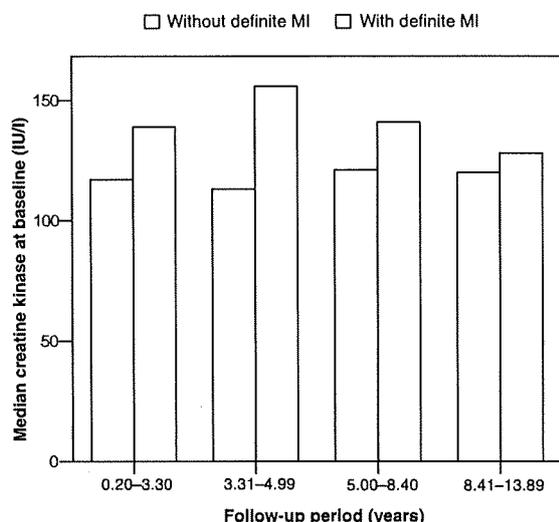


Figure 1 Comparison of median s-CK levels between subjects with and without definite MI by follow-up periods of 0.20–3.30 years, 3.31–4.99 years, 5.00–8.40 years and 8.41–13.89 years

s-CK levels of ≥ 200 IU/l, compared to subjects with levels of ≤ 199 IU/l. Regarding all strokes, this relationship was unclear, although the incidence rate of cerebral infarction in subjects with s-CK levels of ≥ 200 IU/l was slightly higher than in subjects with levels of ≤ 199 IU/l.

Age- and multiple-adjusted HRs are shown in Table 3. Compared with the reference group (≤ 99 IU/l), the adjusted HRs for definite and all MI in subjects with s-CK levels of ≥ 200 IU/l were 4.18 (95% CI 1.66–10.53) and 2.06 (95% CI 1.15–3.70), respectively, in Model 2, and were also linearly elevated with increased s-CK levels, regardless of the model.

In Model 2 for definite MI, excluding further subjects with ST depression on the standard 12-lead ECG, corresponding to 4.2, 4.3 and 4.1 in Minnesota code, or adjusting regular physical activity habits had little effect on the results. In addition, excluding 51 subjects in the 99th percentile or higher of s-CK from the analysis also hardly altered the results. No relationships were observed between s-CK levels and risk for all strokes or cerebral infarction.

For definite MI, interactions between lipid status (hypercholesterolaemia, hypertriglyceridaemia, low HDL-cholesterolaemia) and s-CK levels were shown in Table 4. Interactions in hypercholesterolaemia were observed (P for interaction = 0.011). Among those subjects with hypercholesterolaemia, HRs were linearly elevated, and the HR in s-CK levels of ≥ 200 IU/l was 10.05 (95% CI 2.74–36.81). Similar interaction and linear elevations of HRs were observed among subjects with hypertriglyceridaemia (P for interaction = 0.033) or low-HDL cholesterolaemia (P for interaction = 0.036). Interactions between

other confounding factors (hypertension, impaired glucose tolerance or diabetes, current cigarette use and current alcohol consumption) and s-CK levels were not observed: P for interaction in hypertension, 0.94; for impaired glucose tolerance or diabetes, 0.85; for current cigarette smoking, 0.87; and for current alcohol drinking, 0.96 (data not shown).

There were 45 subjects with definite MI during follow-up: 31 subjects with s-CK of ≤ 199 IU/l, and 14 subjects with s-CK of ≥ 200 IU/l (maximum s-CK was 421 IU/l). Among these 45 definite MI subjects, there was one subject with medication for hypercholesterolaemia and s-CK of ≤ 199 IU/l, one with medication for hypertriglyceridaemia and s-CK of ≥ 200 IU/l, nine with medication for hypertension (seven with s-CK of ≤ 199 IU/l, two with s-CK of ≥ 200 IU/l), and four with medication for diabetes and s-CK of ≤ 199 IU/l (data not shown).

Baseline s-CK distribution or other CVD risk factor profiles did not differ so much between 700 individuals (360 men and 340 women) excluded for loss to follow-up and 5026 study subjects as follows, respectively; s-CK distribution (IU/l) (25th percentile, median, 75th percentile): 88, 118, 157 vs 92, 119, 158, mean age (years): 55.0 vs 54.5, hypertension (%): 32.9 vs 29.3, hypercholesterolaemia (%): 33.6 vs 35.4.

Discussion

This is the first report of a prospective cohort study in which elevated s-CK levels at baseline predicted first-ever MI. Elevated s-CK levels were associated with a moderately increased risk for MI. On the other hand, there was no association observed between s-CK and a risk for stroke. Therefore, elevated s-CK was a specific predictor for MI in the present analysis. In addition, an interaction between s-CK levels and dyslipidaemia (hypercholesterolaemia, hypertriglyceridaemia, low-HDL cholesterolaemia) was observed. S-CK elevation might be a stronger predictor for MI among subjects with dyslipidaemia.

On the basis of present results, following three possible hypotheses could be considered. One hypothesis is the involvement of hypothyroidism that often accompanies both s-CK elevation and dyslipidaemia including hypercholesterolaemia.¹⁶ On the basis of this hypothesis, s-CK elevation is supposed to be a surrogate marker of hypothyroidism. Although the association between hypothyroidism and coronary heart disease is still controversial,^{17–20} recent meta-analyses suggested an association between hypothyroidism and a modest increased risk for coronary heart disease.^{5,6} On the other hand, there were few studies demonstrating causal relationship between hypothyroidism and stroke.^{19,21,22} Present results also demonstrated no relationship between stroke and s-CK elevation, and were consistent with that. Influence of hypothyroidism on stroke might not be as strong as coronary heart disease. It might be partly

Table 2 Crude incidence rates of CVD by s-CK levels in a cohort study of a Japanese urban area, 1989–2005

	Creatine kinase (IU/l)			
	≤99	100–199	200–299	≥300
Men and Women				
<i>N</i>	1590	2816	482	138
Person-years	18 684	33 392	5514	1490
Number of events (crude incidence rate per 1000 person-years)				
All MIs	23 (1.2)	57 (1.7)	18 (3.3)	5 (3.4)
Definite MI	7 (0.4)	24 (0.7)	10 (1.8)	4 (2.7)
All strokes	52 (2.8)	89 (2.7)	21 (3.8)	6 (4.0)
Cerebral infarction	33 (1.8)	55 (1.6)	13 (2.4)	5 (3.4)
Haemorrhagic or unclassified stroke	19 (1.0)	34 (1.0)	8 (1.5)	1 (0.7)
Men				
<i>N</i>	566	1396	306	102
Person-years	6403	16 184	3424	1123
Number of events (crude incidence rate per 1000 person-years)				
All MIs	16 (2.5)	39 (2.4)	12 (3.5)	4 (3.6)
Definite MI	4 (0.6)	18 (1.1)	6 (1.8)	3 (2.7)
All strokes	28 (4.4)	47 (2.9)	17 (5.0)	5 (4.5)
Cerebral infarction	19 (3.0)	33 (2.0)	11 (3.2)	4 (3.6)
Haemorrhagic or unclassified stroke	9 (1.4)	14 (0.9)	6 (1.8)	1 (0.9)
Women				
<i>N</i>	1024	1420	176	36
Person-years	12 281	17 208	2090	366
Number of events (crude incidence rate per 1000 person-years)				
All MIs	7 (0.6)	18 (1.0)	6 (2.9)	1 (2.7)
Definite MI	3 (0.2)	6 (0.3)	4 (1.9)	1 (2.7)
All strokes	24 (2.0)	42 (2.4)	4 (1.9)	1 (2.7)
Cerebral infarction	14 (1.1)	22 (1.3)	2 (1.0)	1 (2.7)
Haemorrhagic or unclassified stroke	10 (0.8)	20 (1.2)	2 (1.0)	0 (0.0)

because of weak relation between dyslipidaemia and stroke in the Japanese population.²³

A second possible hypothesis is that very small infarction or vulnerability of myocardium might precede clinical MI onset.^{3,4} That is, a small amount of intramyocardial protein leakage might occur after micro-MI, or it may follow reversible disturbance of myocardium because permeability of the myocardial cell membrane is metabolically controlled. s-CK elevation might be an indicator that can detect micro-MI or vulnerability before any symptoms or ECG change appears. In this hypothesis, it might be reasonable that s-CK elevation does not predict stroke because s-CK elevation is due to leakage from myocardium.

A third hypothesis is that medication for hyperlipidaemia can occasionally cause rhabdomyolysis of various degrees with s-CK elevation,^{7,8} and is prescribed to individuals with hyperlipidaemia who are possibly at high risk for MI. Therefore, it could become a

confounder for an association between s-CK and MI. However, we thought that this hypothesis had little influence on the present results because there was just one subject medicated for hyperlipidaemia at baseline among those with both definite MI and s-CK of ≥200 IU/l.

We have no information for creatine kinase isozyme that is important to elucidate the background mechanism. Since s-CK in the present study was just one of screening measurements for CVD in the health check-up to healthy individuals that was conducted almost 20 years before, we did not measure creatine kinase isozyme at baseline. Therefore, we can neither infer a further mechanism nor positively support any of above-mentioned hypotheses.

A cross-sectional study demonstrated a positive association between s-CK and blood pressure, and hypothesized that elevated s-CK levels reflected high tissue creatine kinase activity, which might cause

Table 3 Adjusted HRs (95% CIs) for CVDs by s-CK levels in a cohort study of a Japanese urban area, 1989–2005

	N	Number of events	Creatine kinase (IU/l)					P for trend
			≤99 HR	100–199		≥200		
				HR	95% CI	HR	95% CI	
All MIs								
Model 1	5026	103	1 (reference)	1.29	(0.79–2.09)	2.10	(1.17–3.74)	0.015
Model 2	5026	103	1 (reference)	1.35	(0.83–2.20)	2.06	(1.15–3.70)	0.018
Definite MI								
Model 1	5026	45	1 (reference)	1.83	(0.79–4.24)	4.56	(1.84–11.31)	0.001
Model 2	5026	45	1 (reference)	1.75	(0.75–4.09)	4.18	(1.66–10.53)	0.002
All strokes								
Model 1	5026	168	1 (reference)	0.89	(0.63–1.25)	1.10	(0.69–1.75)	0.89
Model 2	5026	168	1 (reference)	0.91	(0.64–1.28)	1.09	(0.68–1.75)	0.90
Cerebral infarction								
Model 1	5026	106	1 (reference)	0.87	(0.56–1.33)	1.13	(0.63–2.00)	0.87
Model 2	5026	106	1 (reference)	0.91	(0.59–1.41)	1.13	(0.63–2.02)	0.83

Model 1: adjusting for age.

Model 2: adjusting for age, sex, body mass index, creatinine, hypertension, impaired fasting glucose or diabetes, hypercholesterolaemia, low HDL-cholesterolaemia, hypertriglyceridaemia, current cigarette use and current alcohol consumption.

Men and women were combined.

Table 4 Adjusted HRs (95% CIs) for definite acute MI stratified by lipid status and s-CK levels in a cohort study of a Japanese urban area, 1989–2005

N	Number of events	Creatine kinase (IU/l)					P for trend	P for interaction
		≤99 HR	100–199		≥200			
			HR	95% CI	HR	95% CI		
Total cholesterol <5.7 mmol/l (220 mg/dl)								
3245	22	1 (reference)	1.90	(0.63–5.76)	0.87	(0.16–4.89)	0.85	0.011
Total cholesterol ≥5.7 mmol/l (220 mg/dl) or use of anti-hypercholesterolaemic medication								
1781	23	1 (reference)	1.39	(0.36–5.31)	10.05	(2.74–36.81)	<0.0001	
Triglycerides <1.7 mmol/l (150 mg/dl)								
3819	28	1 (reference)	1.21	(0.47–3.11)	1.69	(0.53–5.35)	0.39	0.033
Triglycerides ≥1.7 mmol/l (150 mg/dl) or use of anti-hypertriglyceridaemic medication								
1207	17	1 (reference)	5.32	(0.61–46.15)	24.02	(2.79–207.05)	0.0002	
HDL-cholesterol ≤1 mmol/l (39 mg/dl)								
809	17	1 (reference)	4.08	(0.49–33.86)	23.37	(2.78–196.53)	0.0002	0.036
HDL-cholesterol >1 mmol/l (39 mg/dl)								
4217	28	1 (reference)	1.24	(0.48–3.19)	1.88	(0.59–5.98)	0.30	

All HRs adjusting for age, sex, body mass index, creatinine, hypertension, impaired fasting glucose or diabetes, hypercholesterolaemia, low HDL-cholesterolaemia, hypertriglyceridaemia, current cigarette use and current alcohol consumption except for each stratifying factor.

higher blood pressure levels.²⁴ The present analysis of baseline characteristics also demonstrated the possibility of a positive association between hypertension and s-CK levels in women. If this hypothesis is correct, s-CK levels might also be considered a risk factor for stroke. However, results from the present study were not consistent with this hypothesis.

To assure validity of our analysis, it was very important to exclude MI with a past or present history

of MI. First, this analysis excluded the subjects with a medical history of MI, abnormal Q-wave or ST depression on ECG at the baseline survey. Additionally, no subject who had developed MI just at baseline was included (the follow-up period among those with definite MI was, at earliest, 0.20 years). Secondly, regardless of the follow-up period duration, median baseline s-CK levels among subjects with definite MI were higher than among those without

definite MI. As seen in Figure 1, even with a follow-up period of ≥ 8.41 years, median s-CK level among subjects with definite MI was higher than among those without definite MI. This suggested that an association between s-CK levels and MI in a specific period did not influence the present results. Therefore, the possibility that any obviously biased data collection for MI events affected the present results seemed to be low.

The present study had several limitations except for the above-mentioned creatine kinase isozyme. First, since we did not measure thyroid function including thyroid-stimulating hormone at baseline, we could not estimate the involvement of hypothyroidism with the present results. Secondly, we did not collect medical history, such as muscle diseases, thyroid diseases, medications causing rhabdomyolysis or recent history of severe physical training, all of which might affect s-CK levels. Therefore, we could not completely rule out the existence of unknown confounding factors. Thirdly, the single s-CK measurement at baseline may have underestimated the relationship between s-CK levels and CVD due to regression dilution bias.²⁵ Fourthly, the participation rate at baseline was not so high, but those of other population-based cohort studies in Japan with baseline clinical laboratory data were similar,^{26–29} except for large cohort studies with baseline survey by questionnaires.^{30,31} Especially in urban areas, the participation rates in cohort studies tend to be lower because the urban residents generally have more opportunities for health check-up than the rural population.²⁹ That is, the present participation rate was not always low. Moreover, previous reports in the present cohort showed reasonable causal relationship among blood pressure, low-density lipoprotein cholesterol or metabolic syndrome and

CVD,^{9–11} which were similar with previous findings from the Japanese cohort studies.²³ This indirectly suggested that the results from the present cohort and analysis were not extremely biased by participation rates or loss to follow-up.

In conclusion, the present study suggested that screening for elevated s-CK levels in initially healthy Japanese subjects was useful to predict first-ever MI in the future, especially in subjects with dyslipidaemia. However, since the present study was conducted in a limited Japanese population with a relatively low MI incidence rate and had several limitations, these findings should be interpreted carefully and examined by further studies in various populations, races and areas.

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Conflict of interest: None declared.

KEY MESSAGES

- Slight elevation in s-CK is associated with increased risk of MI, not stroke, in population-based cohort study of a Japanese urban area.
- Especially in individuals with dyslipidaemia, measurement of s-CK might be more useful to predict MI.
- However, since present findings seemed to be the first report, they should be examined by further studies in various populations, races and areas.

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Risk of Smoking and Metabolic Syndrome for Incidence of Cardiovascular Disease

— Comparison of Relative Contribution in Urban Japanese Population: The Suita Study —

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Background: Risk factor clustering, the so-called metabolic syndrome (MetS), is an important risk factor for cardiovascular disease (CVD). Smoking is also an important CVD risk factor with still a high prevalence. However, few previous studies have compared the risk for CVD or the population-attributable fraction (PAF) of smoking, MetS, and both.

Methods and Results: The present study was an 11.9-year cohort study of 1,822 men and 2,089 women, aged 40–74 years, selected randomly from an urban general population in Japan. MetS was defined according to the National Cholesterol Education Program on Adult Treatment Panel III (NCEP-ATPIII) guideline modified by the Asian criteria for waist circumference. The prevalence of smoking was 49.5% in men and 11.1% in women, and that of MetS was 19.8% and 23.5%, respectively. In men, the multivariate-adjusted hazard ratio for CVD incidence, compared with non-smoking participants without MetS, was 2.07 (1.26–3.40) in those who smoked, 2.09 (1.08–4.04) in those with MetS, and 3.56 (1.89–6.72) in those with both. In men the PAF for CVD incidence was 21.8% because of smoking, 7.5% because of MetS, and 11.9% because of both.

Conclusions: Although countermeasures for MetS are important, smoking should continue to be considered an important public health problem and antismoking campaigns should be promoted, especially for men, to prevent CVD. (Circ J 2009; 73: 2258–2263)

Key Words: Cohort; Hazard ratio; Metabolic syndrome; Smoking

Risk factor clustering, the so-called metabolic syndrome (MetS), is an important risk factor for cardiovascular disease (CVD), and previous studies have shown the risk of MetS for CVD in the Japanese population.^{1–4} In addition, health guidance for people aged 40–74 years who fulfill the Japanese MetS criteria⁵ began in April 2008 and countermeasures for MetS has become a national project.⁶

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However, cigarette smoking is a widely accepted risk factor for CVD,^{7–9} and the prevalence of smoking is still high in Japan compared with Western developed countries.¹⁰ Accordingly, in Japan, countermeasures for MetS are being applied with a still high prevalence of smoking, which might be different from the situation in Western developed countries with a lower prevalence of smoking.¹⁰ To improve this situation, it is important to examine and show the combined risk of MetS and smoking, and compare the impact of each risk factor and both for CVD from the viewpoint of the impact not only on the individual but also

on the population using indicators such as population-attributable fraction (PAF). In addition, such an assessment could be useful for motivating individuals with MetS, smoking, or both because both MetS and smoking are targets of life-style modification. However, few studies have compared the risk of smoking, MetS, and both for CVD.

Our a priori hypothesis was that the coexistence of smoking and MetS worsens the CVD risk, and that the PAF of smoking in Japanese men is larger than that of MetS because of their high prevalence of smoking. To examine this hypothesis, we performed a 11.9 year (mean length) cohort study in an urban general Japanese population to compare the effects of smoking, MetS and both on CVD risk.

Methods

Population

The Suita study,^{2,11–14} a cohort study of CVD, was established in 1989 in Suita City, Osaka. In that study, 6,485 participants who were randomly selected from the municipal population registry participated in a baseline survey at the National Cardiovascular Center (NCVC) between

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Table 1. Baseline Characteristics of the Participants According to the Combination of Smoking and MetS

	MetS (-)		MetS (+)	
	Non-smoker	Smoker	Non-smoker	Smoker
Men				
n	732	730	189	171
Age (years)	58.9±9.9	56.1±9.4	59.3±8.4	57.4±9.1
Waist (m)	0.82±0.07	0.81±0.07	0.89±0.07	0.89±0.07
BMI (kg/m ²)	22.7±2.7	22.1±2.5	24.9±2.5	24.8±2.5
Total cholesterol (mmol/L)	5.26±0.88	5.08±0.85	5.49±0.88	5.44±0.98
Non-HDL-cholesterol (mmol/L)	3.90±0.88	3.79±0.88	4.41±0.85	4.41±0.98
High blood pressure (%)	48.6	39.3	86.8	84.8
High triglycerides (%)	19.3	22.5	83.1	80.7
Low HDL-cholesterol (%)	13.5	19.6	63.0	69.6
High blood glucose (%)	9.6	10.0	47.1	42.1
Abdominal obesity (%)	13.9	7.5	56.1	59.1
Medication				
For hypertension (%)	32.7	33.7	36.8	39.3
For hypercholesterolemia (%)	1.0	0.5	4.8	4.1
For hypertriglyceridemia (%)	0.5	0.4	2.1	1.2
For diabetes (%)	14.9	12.9	26.9	14.3
Smoking				
Never (%)	37.8	0.0	32.3	0.0
Ex (%)	62.2	0.0	67.7	0.0
Current (%)	0.0	100.0	0.0	100.0
Alcohol drinking				
Never (%)	20.9	19.6	20.6	22.8
Ex (%)	4.2	2.5	5.8	3.5
Current (%)	74.9	77.9	73.5	73.7
Women				
n	1,424	174	433	58
Age (years)	55.3±9.4	52.6±9.1	60.3±8.7	59.3±8.6
Waist (m)	0.77±0.09	0.75±0.09	0.88±0.09	0.87±0.09
BMI (kg/m ²)	21.8±2.8	21.4±3.0	24.8±3.3	24.7±3.2
Total cholesterol (mmol/L)	5.57±0.90	5.39±0.98	5.93±1.00	5.83±0.98
Non-HDL-cholesterol (mmol/L)	4.02±0.90	3.97±1.03	4.75±1.01	4.77±0.95
High blood pressure (%)	35.1	20.1	85.2	70.7
High triglycerides (%)	6.6	6.3	58.0	81.0
Low HDL-cholesterol (%)	18.3	34.5	82.0	87.9
High blood glucose (%)	4.3	1.7	30.5	24.1
Abdominal obesity (%)	30.1	27.6	86.6	79.3
Medication				
For hypertension (%)	33.7	17.4	43.6	44.4
For hypercholesterolemia (%)	1.6	0.0	6.5	3.4
For hypertriglyceridemia (%)	0.1	0.0	1.4	1.7
For diabetes (%)	16.7	0.0	17.5	30.0
Smoking				
Never (%)	97.1	0.0	94.2	0.0
Ex (%)	2.9	0.0	5.8	0.0
Current (%)	0.0	100.0	0.0	100.0
Alcohol drinking				
Never (%)	67.4	50.6	75.5	65.5
Ex (%)	1.0	5.7	1.6	0.0
Current (%)	31.6	43.7	22.9	34.5

Data are value± indicate standard deviation.

MetS = presence of 3 or more of the following: (1) abdominal obesity defined as a waist circumference ≥ 90 cm in men and ≥ 80 cm in women; (2) high blood pressure defined as average systolic/diastolic blood pressures of $\geq 130/85$ mmHg and/or current medication for hypertension; (3) high triglycerides defined as serum level ≥ 1.68 mmol/L; (4) low HDL-cholesterol defined as serum level < 1.03 mmol/L in men and < 1.29 mmol/L in women; (5) high blood glucose defined as fasting blood glucose ≥ 6.10 mmol/L and/or current use of insulin or oral medication for diabetes.

MetS, metabolic syndrome; BMI, body mass index; HDL, high-density lipoprotein.

September 1989 and February 1994. Of the 4,285 participants who were aged 40–74 years at baseline, a total of 374 were excluded for the following reasons: past history of CVD (ischemic heart disease and stroke: $n=127$), non-fasting visit ($n=155$), and missing information at the time of the baseline survey or lost to follow-up ($n=92$). The data for the remaining 3,911 participants (1,822 men and 2,089 women) were then analyzed. Informed consent was given by all participants. The present cohort study was approved by the

Institutional Review Board of the NCVC.

Baseline Examination

Well-trained nurses obtained information on smoking (never, ex-, or current smoker), alcohol drinking (never, ex-, or current drinker), and the medical history of each participant. If the participant answered yes to "current smoker", information was obtained for how many cigarettes per day were smoked.

Table 2. HRs and 95% CIs of Smoking for Incidence of CVD (Stroke + MI), Stroke, Ischemic Stroke, and MI

	Never-smoker	Ex-smoker	Current-smoker	
			≤20 cigarettes/day	>20 cigarettes/day
Men (n)	338	583	524	373
Person-years	4,147	6,837	5,965	4,343
CVD (stroke + MI)				
Cases (n)	11	29	40	16
Incidence (/1,000 person-years)	2.65	4.24	6.71	3.68
Multivariate-adjusted HR (95%CI)	1.00	1.34(0.67–2.69)	2.65(1.35–5.21)	2.31(1.06–5.05)
Stroke				
Cases (n)	8	18	30	12
Incidence (/1,000 person-years)	1.93	2.63	5.03	2.76
Multivariate-adjusted HR (95%CI)	1.00	1.07(0.46–2.48)	2.47(1.12–5.45)	2.48(1.00–6.20)
Ischemic stroke				
Cases (n)	4	16	24	8
Incidence (/1,000 person-years)	0.96	2.34	4.02	1.84
Multivariate-adjusted HR (95%CI)	1.00	1.94(0.64–5.86)	4.06(1.40–11.83)	3.37(1.00–11.41)
MI				
Cases (n)	3	11	10	4
Incidence (/1,000 person-years)	0.72	1.61	1.68	0.92
Multivariate-adjusted HR (95%CI)	1.00	2.21(0.61–8.00)	2.74(0.80–10.90)	1.89(0.41–8.70)
Women (n)	1,790	67	209	23
Person-years	21,881	727	2,363	240
CVD (stroke + MI)				
Cases (n)	45	0	10	1
Incidence (/1,000 person-years)	0.21	–	4.23	4.17
Multivariate-adjusted HR (95%CI)	1.00	–	2.70(1.34–5.45)	2.80(0.36–21.55)
Stroke				
Cases (n)	37	0	5	1
Incidence (/1,000 person-years)	1.69	–	2.12	4.17
Multivariate-adjusted HR (95%CI)	1.00	–	1.60(0.62–4.16)	2.70(0.34–21.68)
Ischemic stroke				
Cases (n)	19	0	4	1
Incidence (/1,000 person-years)	0.87	–	1.69	4.17
Multivariate-adjusted HR (95%CI)	1.00	–	3.00(1.00–8.97)	7.15(0.84–60.64)
MI				
Cases (n)	8	0	5	0
Incidence (/1,000 person-years)	0.37	–	2.12	–
Multivariate-adjusted HR (95%CI)	1.00	–	8.35(2.64–26.48)	–

Multivariate-adjusted HR (95%CI): age, BMI, systolic blood pressure, blood glucose, non-HDL-cholesterol, glomerular filtration rate, and alcohol drinking were adjusted.

HRs, hazard ratios; CIs, confidence intervals; CVD, cardiovascular disease; MI, myocardial infarction. Other abbreviations see in Table 1.

Well-trained physicians measured blood pressure (BP) 3 times in the right arm using a standard mercury sphygmomanometer while the participant was seated after a 5-min rest. The average of the 2nd and 3rd measurements was used in the analyses. Height in stockings and weight in light clothing were measured. Trained public health nurses or technicians measured waist circumference at the umbilical level while the participant was standing.

Blood samples were collected at the NCVS after the participants had fasted for at least 12 h. The samples were centrifuged immediately, and a routine blood examination, which included serum total cholesterol (TC), high-density lipoprotein cholesterol (HDL-C), triglycerides and glucose levels, was then carried out. Non-HDL-C was calculated by subtracting the HDL from the TC. Serum creatinine (Cre) was measured by the non-compensated kinetic Jaffe method. The glomerular filtration rate (GFR: ml·min⁻¹·1.73 m⁻²) was calculated using the MDRD equation modified by the Japanese coefficient (0.881): $186 \times (\text{Cre (mg/dl)})^{-1.154} \times (\text{age (years)})^{-0.203} \times 0.881 \times (0.742 \text{ if female})$.^{15,16}

Definition of MetS

In the present study, MetS was defined using the criteria recommended in the National Cholesterol Education Program

on Adult Treatment Panel III guideline with a modification (modified NCEP-ATP III criteria).^{17,18} Specifically, abdominal obesity was defined as a waist circumference ≥90 cm in men and ≥80 cm in women according to the International Obesity Task Force central obesity criteria for Asia.¹⁷ High BP was defined as average systolic/diastolic BPs ≥130/85 mmHg and/or current medication for hypertension. High triglyceride was defined as a serum level ≥1.68 mmol/L. Low HDLC was defined as a serum level <1.03 mmol/L in men and <1.29 mmol/L in women. High blood glucose was defined as fasting blood glucose (FBG) ≥6.10 mmol/L and/or current use of insulin or oral medication for diabetes. MetS was defined as the presence of 3 or more of these components.

Follow-up and Endpoints

The method of follow-up has been described elsewhere.^{2,11–14} Briefly, the participants were followed until December 31, 2005. The first step in the survey involved checking the health status of all participants by repeat visits to NCVS every 2 years and yearly questionnaires conducted by mail or telephone interview. The in-hospital medical records of the participants who were suspected of having had a myocardial infarction (MI) or stroke were reviewed

Table 3. Risk of Smoking and MetS for CVD (Stroke+MI)

	MetS (-)		MetS (+)	
	Non-smoker	Smoker	Non-smoker	Smoker
Men				
n	732	730	189	171
Person-years	8,721	8,506	2,263	1,835
CVD (stroke+MI) cases (n)	26	41	14	16
CVD incidence (/1,000 person-years)	2.98	4.82	6.19	8.72
Multivariate-adjusted HR (95%CI) [†]	Reference	2.03 (1.24–3.33)	2.11 (1.10–4.04)	3.39 (1.81–6.33)
Multivariate-adjusted HR (95%CI) [‡]	Reference	2.07 (1.26–3.40)	2.09 (1.08–4.04)	3.56 (1.89–6.72)
PAF		21.8	7.5	11.9
Women				
n	1,424	174	433	58
Person-years	17,684	2,027	4,925	577
CVD (stroke+MI) cases (n)	23	6	22	5
CVD incidence (/1,000 person-years)	1.30	2.96	4.47	8.67
Multivariate-adjusted HR (95%CI) [†]	Reference	2.64 (1.07–6.51)	2.58 (1.42–4.69)	5.40 (2.04–14.25)
Multivariate-adjusted HR (95%CI) [‡]	Reference	2.67 (1.07–6.65)	2.33 (1.25–4.34)	4.84 (1.81–12.97)
PAF		6.7	22.4	7.1

Multivariate-adjusted HR (95%CI): [†]adjusted for age.

Multivariate-adjusted HR (95%CI): [‡]adjusted for age, alcohol drinking (never-, ex-, current-), glomerular filtration rate and non-HDL-cholesterol.

PAF, population attributable fraction. Other abbreviations see in Tables 1, 2.

by registered hospital physicians or research physicians who were unaware of the baseline information.

The criteria for definite and probable MI were defined according to the criteria of the Monitoring Trends and Determinants of Cardiovascular Disease (MONICA) project,¹⁹ which requires evidence from an ECG, cardiac enzymes, and/or autopsy. Stroke was defined according to the National Survey of Stroke criteria,²⁰ which require rapid onset of a constellation of neurological deficits lasting at least 24 h or until death. Strokes were classified as ischemic stroke (thrombotic or embolic), intracerebral hemorrhage, subarachnoid hemorrhage, or undetermined type. A definite stroke was defined by autopsy or diagnostic imaging, such as computed tomography or magnetic resonance imaging. In the present study, cases of definite MI or stroke were used in the analysis.

Statistical Analysis

To compare baseline risk characteristics among the 4 groups classified by the combination of MetS and smoking status, analysis of variance was used for continuous variables, and the chi-squared test was used for dichotomous variables. In this analysis, ex-smoker and never-smoker were classified as non-smokers.

Sex-specific analyses were performed. First, the Cox proportional hazards model was used to estimate the hazard ratios (HR) of smoking status for the incidence of CVD (stroke+MI) and its subtypes. Smoking status was classified as never-, ex-, or current smoker (≤ 20 cigarettes/day and > 20 cigarettes/day). In this analysis, age, body mass index (BMI), systolic BP, FBG, non-HDL-C,¹¹ GFR, and alcohol drinking (never-, ex-, and current drinker) were included as confounding factors.

Second, the source population was divided into 4 groups according to the combination of smoking and the presence of MetS. In this analysis, ex-smoker and never-smoker were also classified as non-smokers. The 2 models were used for estimating the HRs of the combinations for CVD incidence. To adjust for the confounding factors, only age was included in model 1, and alcohol drinking (never, ex-, and current drinker), GFR and non-HDL-C were also included

in model 2. To express the impact of smoking on CVD incidence in the participants, the PAF (%) was estimated as $Pe \times (HR - 1) / HR$, in which Pe is the proportion of incident cases in each category.²¹

All statistical analyses were performed using SPSS statistical software, version 15.0 J (SPSS, Tokyo, Japan). $P < 0.05$ (2-tailed) was considered statistically significant.

Results

Baseline Characteristics

Among the participants, 901 of the 1,822 men and 232 of 2,089 women were current smokers (smoking rate: men, 49.5%; women, 11.1%). Similarly, 360 men and 491 women had MetS (prevalence: men, 19.8%; women, 23.5%). Table 1 summarizes the baseline characteristics of the participants classified into 4 groups according to the combination of current smoking and MetS by sex. All variables, except for alcohol drinking in men, were significantly different among the 4 groups.

Risk of Smoking for CVD Incidence

In the present study, the mean follow-up period was 11.9 years, and 42 definite cases of MI and 111 of definite stroke occurred.

Table 2 shows the multivariate-adjusted HRs and 95% confidence intervals (CI) of smoking status for the incidence of CVD and its subtypes. In men, the HR of current smokers who were smoking ≤ 20 cigarettes/day compared with never smokers was 2.65 (95%CI 1.35–5.21) for CVD, 2.47 (95%CI 1.12–5.45) for stroke, 4.06 (95%CI 1.40–11.83) for ischemic stroke, and 2.74 (95%CI 0.80–10.90) for MI. Similarly in women, the HR was 2.70 (95%CI 1.34–5.45) for CVD, 1.60 (95%CI 0.62–4.16) for stroke, 3.00 (95%CI 1.00–8.97) for ischemic stroke, and 8.35 (95%CI 2.64–26.48) for MI. Among the participants who were smoking > 20 cigarettes/day, the HRs for CVD incidence were similar to those who were smoking ≤ 20 cigarettes/day, although in both men and women most of them did not reach to statistical significance because of the small sample size.

Among the ex-smokers, the HR was 1.34 (95%CI 0.67–

2.69) for CVD incidence, 1.07 (95%CI 0.46–2.48) for stroke, 1.94 (95%CI 0.64–5.86) for ischemic stroke, and 2.21 (95%CI 0.61–8.00) for MI in men. In women, there was no case of CVD among ex-smokers.

Risk of Smoking and MetS for CVD Incidence

Table 3 shows the multivariate-adjusted HRs of the combination of smoking and MetS for CVD incidence.

In men, the multivariate-adjusted HRs were 2.07 (95%CI 1.26–3.40) for participants with smoking without MetS, 2.09 (95%CI 1.08–4.04) for those with MetS without smoking, and 3.56 (95%CI 1.89–6.72) for those with both, compared with those both smoking and MetS. In women, the multivariate-adjusted HRs were 2.67 (95%CI 1.07–6.65) for participants with smoking without MetS, 2.33 (95%CI 1.25–4.34) for those with MetS without smoking, and 4.84 (95%CI 1.81–12.97) for those with both, compared with those without both smoking and MetS. When we excluded the ex-smokers among women in this analysis, the HRs were almost similar to the results shown in **Table 3**. And these results were not substantially affected when TC instead of non-HDL-C was included as a confounding factor in the Cox proportional hazard models.

In men the PAF for CVD incidence was 21.8% because of smoking, 7.5% because of MetS, and 11.9% because of both. In women, the respective PAFs were 6.7%, 22.4%, and 7.1%.

Discussion

To our knowledge, this is the first report of a comparison of the CVD risk of smoking, MetS, and both. The magnitude of the HR of smoking or MetS was almost equal. As expected, the risk for the participants with both was the highest. The PAF for CVD incidence among men with smoking alone was much higher than that among those with MetS alone. In women, the PAF among those with MetS was higher than that among those with smoking.

Furthermore, this is also the first report to show the risk of smoking for CVD in an urban area of Japan. In the present study, the prevalence of smoking was 49.5% in men and 11.1% in women. Compared with the data from the National Health and Nutrition Survey conducted in 1989 (men aged 40–69 years in 1989, 50.4–59.5%; women aged 40–69 years in 1988, 6.8–10.6%)²² and several large collaborative cohort studies in Japan,^{8,9,23,24} the prevalence of smoking in the present study was lower in men and higher in women, but is most consistent with the current Japanese prevalence of smoking (men: 39.9%; women: 10.0%). The present study might reflect the prevalence of smoking in urban Japanese communities around the 1990s. In addition, the high smoking prevalence in women and low prevalence in men in the present study is consistent with that in most of the Asia-Pacific region.

Our study showed that smoking is a prominent risk factor for CVD in an urban Japanese cohort, as shown in previous studies in Japanese rural populations.^{9,23,24} Similarly, as previously reported,^{1,25–27} MetS was a risk factor for CVD in our cohort.² The association between MetS and CVD has been reported in several Japanese cohort studies; however, the number of participants was fewer than in the present study,¹ or non-fasting blood samples and BMI were used instead of waist circumference for the analysis.²⁵ These points are another important strength of our study.

MetS has been reported as associated with high percent

plaque volume and abnormal plaque quality in coronary arteries,²⁸ and chronic subclinical inflammation.²⁹ As for smoking, Howard et al reported that smoking is associated with progression of an index of atherosclerosis expressed as the intima–medial thickness of the carotid artery.³⁰ Antoniadou et al also stated that smoking induces both functional and structural abnormalities in the vascular wall, by mechanisms involving endothelial dysfunction and impairment of vascular smooth muscle cells in the human arterial tree.³¹ They also stated that smoking must be approached within the context of the overall lifestyle: smoking coexists with a pro-atherogenic metabolic profile.³¹ The reason for the elevated CVD risk among the present participants with both MetS and smoking is unclear, but the concurrent effect on plaque formation by MetS and smoking, and the additional abnormality in function of vascular smooth muscle cells because of smoking, might be associated with the highest CVD risk among the participants with both risk factors in the present study. Individuals with both smoking and MetS are inevitably in the highest risk group for CVD and should be targeted for intervention.

We compared the HRs of these important CVD risk factors, and the HRs of smoking or MetS for CVD incidence were almost consistent. Accordingly, we calculated the PAF, which shows the impact on CVD incidence. As the result, the PAF of smoking was higher than that of MetS in men, and that of MetS was higher than that of smoking in women, a result that may reflect the higher smoking rate in men. Our study results offer a simple key to solving the problem of “which risk factor should we intervene on first for the population to improve their health outcome”. Recently, the smoking rate has been decreasing in Japanese men; however, compared with the United States for example,¹⁰ it remains still high. As well as countermeasures against MetS, we need to continue considering smoking as an important public health problem and promoting antismoking campaigns in Japan.

In Western developed countries such as the United States, evaluating the risk of MetS under a high prevalence of smoking is difficult because the prevalence of smoking is much lower¹⁰ than in Japan. Although the data of the present study are limited to 1 city in Japan, it might offer evidence of the risk of MetS under a high prevalence of smoking.

There has been controversy about defining the optimal diagnostic criteria for MetS. We have already compared the predictive value between the Japanese criteria and the modified NCEP-ATPIII criteria.² The results suggested that the modified NCEP-ATPIII criteria are suitable for predicting CVD in the Japanese community setting, as well as in the Hisayama study.¹ Accordingly, in the present study MetS was defined using the modified NCEP-ATPIII criteria.^{17,18} Some investigators consider that MetS is an adipose tissue disease different from obesity. If it is an adipose tissue disease, it would be characterized by inflammation detected through high-sensitivity C-reactive protein (hs-CRP) and insulin resistance, reflecting histological changes in adipose tissue.³² Thus, inflammation-related factors such as hs-CRP might be a candidate for 1 of the components of MetS.³³ Furthermore, according to the Japanese MetS criteria, the prevalence of MetS tends to be very low in women because obesity is a required component and the definition of obesity is waist circumference ≥ 90 cm. In addition, because some previous studies showed that the prevalence of non-obese individuals with several metabolic risk factors is high

and their CVD risk is also high, the simple exclusion of non-obese participants from the diagnosis of MetS may overlook their potential risk for CVD.^{25–27} We might misclassify participants with a high risk for CVD if we adopt the Japanese MetS criteria.

Study Limitations

First, we could not assess the risk of smoking on the incidence of hemorrhagic stroke because of the small number of cases. Second, the measurement of single MetS components and the questionnaire for smoking in the baseline survey may have underestimated the relationship between these risk factors and CVD because of a regression dilution bias.

In conclusion, smoking is still an important risk factor for CVD in urban areas of Japan, and the combination of smoking and MetS worsens the risk for CVD. Lifestyle modification for not only MetS but also smoking continues to be important in populations with a high PAF for CVD because of a high prevalence of smoking.

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Disclosure

None.

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Triglycerides and non-high-density lipoprotein cholesterol and the incidence of cardiovascular disease in an urban Japanese cohort: The Suita study

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ABSTRACT

Objective: The impact of elevated triglycerides (TG) and non-high density lipoprotein cholesterol (non-HDLc) on the incidence of stroke and myocardial infarction (MI) has not been well evaluated in Asian populations such as in Japan, which have a lower incidence of myocardial infarction, but a higher risk of stroke than Western populations.

Methods: The authors conducted an 11.7-year prospective study ending in 2005 of 5098 Japanese aged 30–79 living in an urban population, initially free of stroke or MI. The relationship between serum lipids and the risk for stroke and MI was determined by dividing the participants into four groups stratified by the combination of serum levels of TG and non-HDLc. The cut-off value was 1.7 mmol/L for TG and 4.9 mmol/L for non-HDLc.

Results and conclusion: The total person-years were 59,774 (27,461 for men and 32,313 for women). During the follow-up period, there were 113 cases of MI and 180 of stroke (with 116 cerebral infarctions). Compared with the low TG/low non-HDLc group, the hazard ratio (95% confidence interval) for MI in the high TG/high non-HDLc group was 2.55 (1.53–4.24) after adjustment for other cardiovascular risk factors. The hazard ratio for cerebral infarction in the high TG alone group was 1.63 (1.03–2.56); however, the risk of cerebral infarction was not significantly increased in the other groups. High serum levels of TG and non-HDLc are both important targets for the prevention of cardiovascular disease in Japan.

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1. Introduction

Previous studies suggested that high levels of serum total cholesterol (TC) and low density lipoprotein cholesterol (LDL-C) are causal risk factors for coronary artery disease (CAD) [1–4] and possibly for ischemic stroke [5]. However, less attention has been paid to high serum levels of triglycerides (TG) [6–8]. Furthermore, although the US National Cholesterol Education Program Adult Treatment Panel guideline III (NCEP-ATP III) has set goals for non-high-density lipoprotein cholesterol (non-HDLc) after the achievement of LDL-C goals in patients with elevated TG [9], the impact of TG and non-HDLc on the incidence of cardiovascular disease (CVD) has not been evaluated in the Japanese population, which has a lower incidence of CAD but a higher risk of stroke than Western populations [10].

Therefore, our a priori hypothesis was that the coexistence of high serum TG and non-HDLc increases the risk of CAD and stroke in the Japanese population. To investigate this hypothesis, we performed a long-term prospective study in an urban, community-dwelling Japanese population.

2. Methods

2.1. Populations

The Suita study, a cohort study for CVD of urban residents was established in 1989. The details of this study have been described elsewhere [4,11–14]. Briefly, 6485 men and women aged 30–79 years had a baseline survey at the National Cardiovascular Center between September 1989 and March 1994. Of these, a total of 1387 were excluded for the following reasons: past history of coronary heart disease or stroke ($n=210$), lack of participation in the baseline survey ($n=79$), non-fasting visit ($n=166$), use of lipid-lowering agents ($n=125$), missing data ($n=109$), and lost to follow-up ($n=698$). Data from the remaining 5098 participants (2404 men and 2694 women) were included in the analysis. This

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