

high price of the 21-gene RT-PCR assay is not cancelled out by the reduction of adjuvant chemotherapy.

Costs treating LN-, ER+, ESBC incidence with NCCN/St Gallen/RS-guided treatment are calculated by the year taking mortality into account, and incremental costs are also calculated by the year according to comparisons. Calculated with these costs, the budget impact of the diffusion of the assay in Japan is estimated as ¥2,638 million (US\$23 million) to ¥3,225 million (US\$28 million).

Discussion

We evaluate the cost-effectiveness of the 21-gene RT-PCR assay in Japan's health care system with two scenarios depicting status quo and one scenario of the routine use of the assay for LN-, ER+, ESBC. Our economic model indicates that the diffusion of the assay gains more in terms of outcome but costs more at the same time. The estimated ICERs, 2,997,495 ¥/QALY (26,065 US\$/QALY) and 1,239,055 ¥/QALY (10,774 US\$/QALY), comparing NCCN/St Gallen-guided treatment with RS-guided treatment, respectively, are not more than a suggested social willingness-to-pay for one life year gain from an innovative medical intervention in Japan, 6,000,000 ¥/QALY (52,174 US\$/QALY) [36]. Sensitivity analyses show that this result is plausibly robust, since ICERs do not exceed the threshold by various changes of assumptions made or values employed. In this sense, the assay has good value for money.

Incremental effects in terms of QALY are longer than those in terms of YOLS; and ICERs in terms of yen per QALY are smaller than those in terms of yen per YOLS in both comparisons. These imply that the assay is not only efficient in prolonging survival but also improving quality of life.

Our sensitivity analyses also reveal that the price of the assay is one of the major determinants of cost-effectiveness as expected. An intuitive comparison with the price of a conventional gene diagnosis test of malignant tumour in Japan, ¥450,000 (US\$3,913) vs. ¥20,000 (US\$174), seems to make a health manager feel it difficult to reimburse the cost of the assay by the social insurance, because there may be an incompatibility to an incremental manner of revising fee schedule. Our study, however, implies that the price offered by Japanese supplier of Oncotype DX[®] Breast Cancer Assay still makes ICER an acceptable level from the viewpoint of welfare economics.

We estimate the budget impact of the assay on the social health insurance system. The policy implication of the budget impact is not prescriptive [37]. Yet, the estimated impact, ¥2,638 million (US\$23 million) to ¥3,225 million (US\$28 million) per year for the coming 5 years, is

substantially less than the estimated budget impact of adjuvant trastuzumab, which is about to be included into social insurance benefit, ¥16,000 million (US\$139 million) to ¥32,000 million (US\$278 million) [38]. The characteristics of the assay of which application is limited to only once per case probably contribute to this difference, since the cost of trastuzumab amounts through its repeated administration. This implies that the diffusion of the assay through listing as an approved diagnostic test by the social health insurance could be justifiable.

The past economic evaluation of the assay reported from the U.S. considers a change from NCCN-guided treatment to RS-guided treatment [19], while our model allows a comparison between NCCN-guided treatment and St Gallen-guided treatment as an *ex ante* scenario. We find a notable difference in ICERs in this comparison. The ICER of the change from St Gallen-guided treatment is more favourable than that from NCCN-guided treatment. This is interesting because the reduction of use of adjuvant chemotherapy according to the reclassification from St Gallen criteria, 26%, is smaller than that from NCCN, 43%. The difference in ICER is due to more gain in the outcome. Although caution is needed in transferring the findings from economic models to any different context [39], our model might indicate that the assay has better value for money in countries where St Gallen-guided treatment is widely used.

However, this study has its own limitations. First, our outcome estimation depends on the validation studies carried out in the U.S. Although the evidences adopted are considered as the best available knowledge, it is needless to say that there are differences in population, as well as in cancer care practice between the U.S. and Japan. With this in regard, another validation study employing Japanese historical clinical trial data with the gene assay of preserved tumour tissue is launched [40]. A further economic evaluation incorporating new evidences is necessary to confirm the findings of this study. Second, utility weights adopted here are also derived from Western countries due to an unavailability of data from Japan. Third, our model does not include potentially costly clinical stages such as local recurrence or contralateral breast cancer due to the lack of data in validation studies. Regarding these shortcomings, reports and data that refines the model are awaited. Fourth, consensus guidelines are renewed continuously by incorporating newly available evidences [11, 41], so that the relative usefulness of the assay may be diminished in the near future, or the assay may be incorporated in the guidelines in a long run.

The use of the 21-gene RT-PCR assay has just begun to have an impact on clinical recommendations made by the U.S. oncologists and patients' choice [42]. It is easy to imagine that similar change in practice will occur in Japan

soon, because patients have strong preference to innovation such as tailor-made medicine [1]. As the prognostic usefulness of the 21-gene RT-PCR assay in guiding treatment for lymph-node-positive cases is recently reported [43], the indication of the assay will expand. Further economic evaluation that responds to this contextual change may become imperative.

Once the usefulness of the assay is confirmed by the Japanese validation study, Japanese health manager inevitably needs to decide how to fit the assay to the health care system. The results of this study imply the possibility of coverage by the social insurance. If health manager gives much importance to fiscal policy or cost containment, the selective indication of the assay for higher risk patients, which results to avoid additional use of adjuvant chemotherapy, might be a potential option. Further analysis incorporating such scenarios may be useful.

In conclusion, the routine use of the 21-gene RT-PCR assay for LN-, ER+, ESBC is indicated as cost-effective with a fundable level of budget impact in Japan. The results could inform health managers in developed countries where NCCN-guided treatment as well as St Gallen-guided treatment are practiced.

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Economic evaluation of chemoprevention of breast cancer with tamoxifen and raloxifene among high-risk women in Japan

M Kondo^{*1,2}, S-L Hoshi¹ and M Toi³

¹Department of Health Care Policy and Management, Graduate School of Comprehensive Human Sciences, University of Tsukuba, 1-1-1 Tennoudai, Tsukuba, Ibaraki 305-8577, Japan; ²Clinical Research Division, Tokyo Metropolitan Cancer and Infectious Disease Centre, Komagome Hospital, 3-18-22 Honkomagome, Bunkyo-ku, Tokyo 113-8677, Japan; ³Department of Surgery, Graduate School of Medicine, Kyoto University, 54 Kawaracho, Shogoin, Sakyo-ku, Kyoto 606-8507, Japan

Raloxifene was approved for chemoprevention against breast cancer among high-risk women in addition to tamoxifen by the US Food and Drug Administration. This study aims to evaluate cost-effectiveness of these agents under Japan's health system. A cost-effectiveness analysis with Markov model consisting of eight health states such as healthy, invasive breast cancer, and endometrial cancer is carried out. The model incorporated the findings of National Surgical Adjuvant Breast and Bowel Project P-1 and P-2 trial, and key costs obtained from health insurance claim reviews. Favourable results, that is cost saving or cost-effective, are found by both tamoxifen and raloxifene for the introduction of chemoprevention among extremely high-risk women such as having a history of atypical hyperplasia, a history of lobular carcinoma *in situ* or a 5-year predicted breast cancer risk of $\geq 5.01\%$ starting at younger age, whereas unfavourable results, that is 'cost more and gain less' or cost-ineffective, are found for women with a 5-year predicted breast cancer risk of $\leq 5.00\%$. Therapeutic policy switch from tamoxifen to raloxifene among postmenopausal women are implied cost-effective. Findings suggest that introduction of chemoprevention targeting extremely high-risk women in Japan can be justifiable as an efficient use of finite health-care resources, possibly contributing to cost containment.

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Several clinical trials have demonstrated the effectiveness of prophylactic administration of selective oestrogen receptor modulators (SERMs) such as tamoxifen (Fisher *et al*, 2005; Cuzick *et al*, 2007; Powles *et al*, 2007; Veronesi *et al*, 2007b) and raloxifene (Cauley *et al*, 2001; Martino *et al*, 2004; Vogel *et al*, 2006) in reducing incidence of breast cancer among women at high risk of developing the disease. Tamoxifen was approved for prophylaxis by the US Food and Drug Administration in 1998, and raloxifene was also approved for postmenopausal women in 2007.

Tamoxifen reduces the risk of breast cancer whereas increasing the risk of adverse events such as endometrial cancer and pulmonary embolism. Raloxifene is a second-generation SERM usually used for osteoporosis treatment, and it reduces the risk of invasive breast cancer with a lower risk of known adverse events associated with SERMs, compared to tamoxifen. This is because raloxifene does not induce the unwanted stimulation of endometrium (Delmas *et al*, 1997). Therefore, raloxifene is considered to have a better clinical property as prophylactic agent, although it is inferior to tamoxifen in preventing noninvasive breast cancer. More women at high risk of developing breast cancer are expected to take raloxifene as their breast cancer prevention drug in the United States (Bevers, 2007).

However, both of these agents have been neither approved nor made available for its use as breast cancer prevention in Japan, although experts have shown their expectations (Iwata and Saeki, 2006). It is said that there are five hurdles to overcome in addressing intervention in the diffusion process of new drug: quality, safety, efficacy, cost-effectiveness, and affordability (Trueman *et al*, 2001). This paper aims to present evidence to the fourth hurdle, cost-effectiveness of both agents, under Japan's health system. Although cost-effectiveness of prophylactic use of tamoxifen has been reported from the USA (Noe *et al*, 1999; Grann *et al*, 2000; Smith and Hillner, 2000; Hershman *et al*, 2002; Melnikow *et al*, 2006) and Australia (Eckermann *et al*, 2003), that of raloxifene has not been published to date except as a part of economic evaluation of osteoporosis management (Armstrong *et al*, 2001; Kanis *et al*, 2005). This paper also simulates a therapeutic policy switch from tamoxifen to raloxifene among postmenopausal women to illustrate the relative value of raloxifene. Consequently, it should have implications to the developed countries where chemoprevention with tamoxifen is already in practise.

METHODS

We conduct a cost-effectiveness analysis with Markov modelling based on the findings of the National Surgical Adjuvant Breast and Bowel Project (NSABP) P-1 trial (Fisher *et al*, 2005), the NSABP P-2 trial (Vogel *et al*, 2006), and the literature on costing under

*Correspondence: Dr M Kondo; E-mail: mkondo@md.tsukuba.ac.jp
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Japan's health system including sensitivity analyses from societal perspective. Although longer follow-up results for tamoxifen are reported from the first International Breast Cancer Intervention Study (IBIS-I; Cuzick *et al*, 2007) and the Royal Marsden trial (Powles *et al*, 2007), NSABP P-1 trial with a shorter follow-up period is chosen as clinical evidence for our modelling to make clear comparisons with NSABP P-2 trial of raloxifene. The long-term outcomes for tamoxifen (Veronesi *et al*, 2007a) are considered in our sensitivity analyses. We use TreeAge Pro 2008 (TreeAge Software Inc.) for our economic modelling.

High-risk women

We model high-risk women according to the risk classifications featured in the report of clinical trials: three levels (≥ 1.66 , 3.01–5.00%, $\geq 5.01\%$) of a 5-year predicted breast cancer risk, with a history of lobular carcinoma *in situ* (LCIS), and with a history of atypical hyperplasia (AH). A 5-year predicted breast cancer risk of an individual woman used in the trials is based on Gail *et al* model 2 (Gail and Costantino, 2001), which is validated for white women (Rockhill *et al*, 2001) and African American women (Gail *et al*, 2007), to date. We assume the same model is good for Japanese women.

We also model the ages of starting prophylaxis: 35, 50, 60 years old for tamoxifen, and 50, 60 years old for raloxifene taking the menopause into account.

Markov model

We construct a Markov model of courses followed by high-risk women, which is shown in Figure 1. Eight health states are modelled according to clinical events monitored and found significant in P-1 trial and P-2 trial: (1) healthy; (2) invasive breast cancer; (3) noninvasive breast cancer, (4) endometrial cancer; (5) pulmonary embolism; (6) cataract; (7) hip fracture; and (8) dead. Healthy women at high risk of the disease, women with invasive and noninvasive breast cancer are the target health states for chemoprevention. An increase in risk of endometrial cancer, pulmonary embolism, and cataract are known as adverse effects of SERMs, whereas a decrease in risk of hip fracture is known as a beneficial effect. Transitions between health states are indicated with arrows.

The time span of each stage is set at 1 year, since trials report annual incidence rates. Markov process is repeated until death or age 100, whichever comes first, since all events are expected to occur within this time horizon. Women who survive after the age

of 100 years are assumed to die regardless of breast cancer development.

Chemoprevention

Prophylaxis with SERMs is continued for 5 years, or discontinued in case of adverse events, which is similar to the regimen employed in clinical trials.

Comparisons

We compare outcomes and costs in terms of incremental cost-effectiveness ratios (ICERs) between *status quo* in Japan, without prophylaxis, and hypothetical practise, with prophylaxis, by the agent (tamoxifen and raloxifene), the risk classification, and the age of starting prophylaxis.

$$\text{ICER} = \frac{\text{Cost}_{\text{with prophylaxis}} - \text{Cost}_{\text{without prophylaxis}}}{\text{Effect}_{\text{with prophylaxis}} - \text{Effect}_{\text{without prophylaxis}}}$$

We also compare prophylaxis with tamoxifen and prophylaxis with raloxifene to estimate the relative value of raloxifene to tamoxifen, although this does not depict any marginal change in Japan.

Outcome estimation

Outcomes in terms of life years gained (LYGs) and quality adjusted life years (QALYs) are estimated by assigning transitional probabilities and utility weights to Markov model from the literature.

Transitional probabilities from healthy state to disease states in Markov model are shown in Table 1 according to the findings from the clinical trials. Risk reduction effect of SERMs is assumed to continue during the 5-year course of prophylaxis.

Table 2 summarises other assumptions such as transitional probabilities from disease states to dead state and utility weights used in Markov model. The share of clinical stages of invasive breast cancer at diagnosis are adopted from a nationwide survey on breast cancer screening (Japan Cancer Society, 2007), of which prognosis is calculated from corresponding follow-up cases at Tokyo Metropolitan Cancer and Infectious Disease Centre Komagome Hospital. The prognosis of endometrial cancer is also adopted from a nationwide cancer registry (Japanese Society of Obstetrics and Gynecology, 2000). The prognosis of pulmonary embolism and hip fracture are taken from Sakuma *et al* (2004); Kitamura *et al* (1998), respectively. Japanese female population

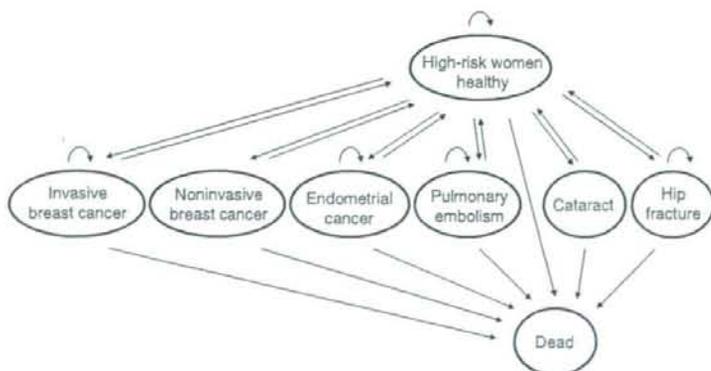


Figure 1 Markov model.

Table 1 Transitional probabilities from healthy state to disease states in Markov model

| | Placebo | | Tamoxifen | | | Raloxifene | | |
|---|-----------------|---------------------|-----------------|---|---------------------|-----------------|---|---|
| | Base-case value | Source | Base-case value | Range tested in sensitivity analysis ^a | Source | Base-case value | Range tested in sensitivity analysis ^a | Source |
| Invasive breast cancer | | | | | | | | |
| Five-year predicted breast cancer risk \geq 1.66% | | | | | | | | |
| Age of starting prophylaxis | | | | | | | | |
| 35 | 0.00632 | Fisher et al (2005) | 0.00404 | 0.00235–0.00641 | Fisher et al (2005) | | | |
| 50 | 0.00587 | Fisher et al (2005) | 0.00333 | 0.00168–0.00573 | Fisher et al (2005) | 0.00310 | 0.00184–0.00490 | Fisher et al (2005), Vogel et al (2006) |
| 60 | 0.00668 | Fisher et al (2005) | 0.00330 | 0.00165–0.00567 | Fisher et al (2005) | 0.00366 | 0.00213–0.00585 | Fisher et al (2005), Vogel et al (2006) |
| Five-year predicted breast cancer risk 3.01–5.00% | 0.00451 | Fisher et al (2005) | 0.00270 | 0.00108–0.00534 | Fisher et al (2005) | 0.00203 | 0.00101–0.00349 | Fisher et al (2005), Vogel et al (2006) |
| Five-year predicted breast cancer risk \geq 5.01% | 0.01198 | Fisher et al (2005) | 0.00515 | 0.00245–0.00893 | Fisher et al (2005) | 0.00561 | 0.00323–0.00894 | Fisher et al (2005), Vogel et al (2006) |
| History of lobular carcinoma in situ | 0.01170 | Fisher et al (2005) | 0.00627 | 0.00161–0.01476 | Fisher et al (2005) | 0.00614 | 0.00239–0.01226 | Fisher et al (2005), Vogel et al (2006) |
| History of atypical hyperplasia | 0.01042 | Fisher et al (2005) | 0.00255 | 0.00029–0.00686 | Fisher et al (2005) | 0.00286 | 0.00133–0.00523 | Fisher et al (2005), Vogel et al (2006) |
| Noninvasive breast cancer | 0.00012 | Fisher et al (2005) | 0.00004 | 0.00000–0.00652 | Fisher et al (2005) | 0.00006 | 0.00003–0.00009 | Fisher et al (2005), Vogel et al (2006) |
| Endometrial cancer | | | | | | | | |
| Age of starting prophylaxis | | | | | | | | |
| 35 | 0.00082 | Fisher et al (2005) | 0.00116 | 0.00010–0.00410 | Fisher et al (2005) | | | |
| 50 and 60 | 0.00058 | Fisher et al (2005) | 0.00308 | 0.00061–0.00992 | Fisher et al (2005) | 0.00194 | 0.00065–0.00403 | Fisher et al (2005), Vogel et al (2006) |
| Pulmonary embolism | | | | | | | | |
| Age of starting prophylaxis | | | | | | | | |
| 35 | 0.00013 | Fisher et al (2005) | 0.00025 | 0.00000–0.00420 | Fisher et al (2005) | | | |
| 50 and 60 | 0.00044 | Fisher et al (2005) | 0.00096 | 0.00020–0.00275 | Fisher et al (2005) | 0.00061 | 0.00028–0.00114 | Fisher et al (2005), Vogel et al (2006) |
| Cataract | 0.02285 | Fisher et al (2005) | 0.02775 | 0.02384–0.03206 | Fisher et al (2005) | 0.02192 | 0.01735–0.02734 | Fisher et al (2005), Vogel et al (2006) |
| Hip fracture | 0.00086 | Fisher et al (2005) | 0.00059 | 0.00022–0.00122 | Fisher et al (2005) | 0.00052 | 0.00016–0.00115 | Fisher et al (2005), Vogel et al (2006) |

^a1.5 times of 95% confidence interval.

mortality rates from Vital Statistics (Ministry of Health, Labour and Welfare, 2005a) are applied for other transitions to dead state.

It is more preferable to adopt utility weights from a consistent study that assesses our six disease states in Japan, but there is no Japanese utility weight in the literature to date, which may be applied to any health states in our model. To illustrate the typical patient states, we adopt the weights assessed in developed countries considering them as the best available knowledge, and choosing them under the consensus of staff doctors at Tokyo Metropolitan Cancer and Infectious Disease Centre Komagome Hospital (de Koning et al, 1991; Hillner et al, 1993; Smith and Hillner, 1993; Grann et al, 1998; Earle et al, 2000; Armstrong et al, 2001; Chau et al, 2003; Cykert et al, 2004; Naeim and Keeler, 2005; Ruof et al, 2005).

Outcome is discounted at a rate of 3%.

Costing

From societal perspective, costing should cover the opportunity cost borne by various economic entities in the society. In the context of this study, costs borne by women or third party payers including the government and social insurers are considered, although there is no particular assumption about who bears the cost of chemoprevention. According to the national medical care

fee schedule, the amount of direct payments to health-care providers is estimated as cost, whereas costs to sectors other than health and productivity losses are left uncounted.

Health states are identified as cost items in Markov model. Table 3 summarises the cost of each health states. Being in healthy state, women with chemoprevention take 20 mg per day, ¥82.6 (£0.41; £1 = ¥200), of tamoxifen, or 60 mg per day, ¥148.5 (£0.74), of raloxifene, prescribed regularly for 5 years, and annual mammography checkup. Women without chemoprevention also undergo annual mammography checkup. Although the state is labelled as 'healthy', it includes all other diseases that are not modelled in Markov model. Annual treatment costs by the age stratum are approximated by annual health-care expenditure per woman adopted from National Health-Care Expenditure (Ministry of Health, Labour and Welfare, 2005b). As it is well known that the cost of health care in the last year of life tends to be large, these are shown separately after an adjustment based on Fukawa (1998).

Table 3 also summarises the treatment cost of invasive breast cancer by the age stratum. In the case of cancer care, the cost in the first year after diagnosis tends to be large as well as in the last year of life, so here again, the costs are shown separately. These figures are obtained from insurance claim reviews at Tokyo Metropolitan Cancer and Infectious Disease Centre Komagome Hospital. As to the cost of the first year, recent breast cancer cases of stage I and

Table 2 Assumptions used in Markov model

| Assumption | Range tested in sensitivity analysis | Source |
|---|---|--|
| Transitional probabilities from disease states to dead state | | |
| Invasive breast cancer | 0–9 years after diagnosis: prognosis of Japanese breast cancer patients by the stage Stage I: 0.0074, 0.0155, 0.0113, 0.0218, 0.0254, 0.0248, 0.0289, 0.0165, 0.01632 Stage II: 0.0054, 0.0474, 0.0570, 0.0334, 0.0398, 0.0321, 0.0275, 0.0295, 0.04672 (Proportions of stage at diagnosis are assumed stage I as 72% and stage II as 28%) | Change by $\pm 50\%$ Change by $\pm 50\%$ Japan Cancer Society (2007) |
| Noninvasive breast cancer | Thereafter: Japanese female population mortality rates | Change by $\pm 50\%$ |
| Endometrial cancer | Japanese female population mortality rates 0–4 years after diagnosis: prognosis of Japanese endometrial cancer patients 0.0660, 0.0546, 0.0328, 0.02813 Thereafter: Japanese female population mortality rates | Change by $\pm 50\%$ Change by $\pm 50\%$ Change by $\pm 50\%$ Ministry of Health, Labour and Welfare (2005a) Ministry of Health, Labour and Welfare (2005a) Japanese Society of Obstetrics and Gynecology (2000) |
| Pulmonary embolism | 0 year after diagnosis: 0.08 Thereafter: Japanese female population mortality rates | Change by $\pm 50\%$ Change by $\pm 50\%$ Ministry of Health, Labour and Welfare (2005a) |
| Cataracts | Japanese female population mortality rates | Change by $\pm 50\%$ Ministry of Health, Labour and Welfare (2005a) |
| Hip fracture | 0–1 years after diagnosis: 0.11 and 0.19, respectively Thereafter: Japanese female population mortality rates | Change by $\pm 50\%$ Change by $\pm 50\%$ Kitamura et al (1998) Ministry of Health, Labour and Welfare (2005a) |
| Utility weights | | |
| Healthy | 1.00 | Change by $\pm 20\%$ |
| Healthy under chemoprevention for 5 years | 0.99 | Change by $\pm 20\%$ Smith and Hillner (1993), Hillner et al (1993), Naeim and Keeler (2005) |
| Invasive breast cancer | 0 year after diagnosis: 0.87, thereafter: 0.89 | Change by $\pm 20\%$ de Koning et al (1991), Granin et al (1998) |
| Noninvasive breast cancer | 0.98 | Change by $\pm 20\%$ Earle et al (2000) |
| Endometrial cancer | 0 year after diagnosis: 0.83, thereafter: 0.88 | Change by $\pm 20\%$ Armstrong et al (2001), Cykert et al (2004) |
| Pulmonary embolism | 0.70 | Change by $\pm 20\%$ Chau et al (2003) |
| Cataract surgery | 0.96 | Change by $\pm 20\%$ Ruof et al (2005) |
| Hip fracture | 0–1 years after diagnosis: 0.61 and 0.92, respectively | Change by $\pm 20\%$ Armstrong et al (2001) |

stage II that have undergone initial treatment with a follow-up of 1 year are retrospectively selected so that each age strata has 40 cases. As to the yearly cost of the second year and thereafter, 40 cases for each age strata are randomly selected from follow-up cases initially diagnosed as stage I and stage II. As to the cost of the last year of life, recent 80 fatal cases are retrospectively selected, as the number of these is relatively limited. Insurance claims of these total of 400 cases for 1 year are reviewed to calculate average annual costs by the age strata. Then an adjustment is made to include the cost of prescription to be filled at external pharmacies, such as in the case of adjuvant hormonal therapy, which is based on the consensus among staff doctors.

Costs of disease states are summarised in Table 3 as well. Treatment costs of noninvasive breast cancer, endometrial cancer, cataract, and hip fractures are adopted from a background study for the development of Japanese prospective payment system to health-care providers, diagnosis procedure combination (Matsuda and Ishikawa, 2003), whereas treatment cost of pulmonary embolism is adopted from Fuji et al (2005).

Costs are also discounted at a rate of 3%.

Sensitivity analyses

To deal with the uncertainty of probabilities, utility weights, and costs used in our economic model, one-way sensitivity analyses are performed. Transitional probabilities from healthy state to disease states shown in Table 1 are varied in 1.5 times of 95% confidence intervals (CI) reported from the clinical trials. 95% CI is often used for similar exercises of sensitivity analyses, but we set wider range for the applicability of the clinical trial data to Japanese women. The other probabilities shown in Table 2 are changed by $\pm 50\%$. Utility weights are changed by $\pm 20\%$, and we think this could

cover the difference between the utility weights of Japanese women and those of the other developed nations. Costs shown in Table 3 are changed by $\pm 50\%$. Discount rate is also changed from 0 to 6%.

Acknowledging the long-term outcomes for tamoxifen in the IBIS-I trial (Cuzick et al, 2007) and the Royal Marsden trial (Powles et al, 2007), risk reduction effect of tamoxifen is prolonged from 5 to 10 and 15 years without any risk increase of adverse events after the completion of prophylaxis.

RESULTS

Outcomes

Table 4 shows the results of cost-effectiveness analysis comparing prophylaxis with no prophylaxis.

In the comparison between prophylaxis with tamoxifen vs no prophylaxis, most outcomes in terms of LYGs are increased by chemoprevention except for women with a 5-year predicted breast cancer risk of $\geq 1.66\%$ starting at age 50, and women with a 5-year predicted breast cancer risk of 3.01–5.00% starting at age 50 and 60. Outcomes in terms of QALYs are also increased except for women with a 5-year predicted breast cancer risk of $\geq 1.66\%$ starting at age 50 and 60, women with a 5-year predicted breast cancer risk of 3.01–5.00%, and women with a history of LCIS starting at age 60. The largest outcome gain in terms of QALYs, 0.105, is estimated among women with a history of AH starting at age 35.

Between prophylaxis with raloxifene vs no prophylaxis, all outcomes in terms of LYGs are increased by chemoprevention. Outcomes in terms of QALYs are increased except for women with a 5-year predicted breast cancer risk of $\geq 1.66\%$, and women with

Table 3 Costs (¥)

| | Healthy | | | Breast cancer | | |
|---------------------------------------|-----------------|--------------------------------------|---|-----------------|--------------------------------------|------------------------|
| | Base-case value | Range tested in sensitivity analysis | Source | Base-case value | Range tested in sensitivity analysis | Source |
| Chemoprevention | | | | | | |
| Tamoxifen | 30 149 | Change by \pm 50% | Drug price list, etc | | | |
| Raloxifene | 54 203 | Change by \pm 50% | | | | |
| Prescription+annual mammography | 44 980 | Change by \pm 50% | | | | |
| Annual mammography | 15 520 | Change by \pm 50% | | | | |
| Ages 35–49 | | | | | | |
| First year after diagnosis | | | | 1978 064 | Change by \pm 50% | |
| Yearly cost | | | | 383 743 | Change by \pm 50% | |
| Ages 35–39 | 81 937 | Change by \pm 50% | Ministry of Health, Labour and Welfare (2005b), Fukawa (1998) | | | Insurance claim review |
| Ages 40–44 | 94 529 | Change by \pm 50% | | | | |
| Ages 45–49 | 110 604 | Change by \pm 50% | | | | |
| Terminal care cost, last year of life | | | | 5495 224 | Change by \pm 50% | |
| Ages 35–39 | 352 331 | Change by \pm 50% | | | | |
| Ages 40–44 | 406 474 | Change by \pm 50% | | | | |
| Ages 45–49 | 475 599 | Change by \pm 50% | | | Change by \pm 50% | |
| Ages 50–64 | | | | | | |
| First year after diagnosis | | | | 2211 083 | Change by \pm 50% | |
| Yearly cost | | | | 542 857 | Change by \pm 50% | |
| Ages 50–54 | 151 625 | Change by \pm 50% | Ministry of Health, Labour and Welfare (2005b), Fukawa (1998) | | | Insurance claim review |
| Ages 55–59 | 195 085 | Change by \pm 50% | | | | |
| Ages 60–64 | 258 723 | Change by \pm 50% | | | | |
| Terminal care cost, last year of life | | | | 4106 271 | Change by \pm 50% | |
| Ages 50–54 | 651 986 | Change by \pm 50% | | | | |
| Ages 55–59 | 838 866 | Change by \pm 50% | | | | |
| Ages 60–64 | 1112 510 | Change by \pm 50% | | | | |
| Ages 65–79 | | | | | | |
| First year after diagnosis | | | | 1530 259 | Change by \pm 50% | |
| Yearly cost | | | | 441 458 | Change by \pm 50% | |
| Ages 65–69 | 324 347 | Change by \pm 50% | Ministry of Health, Labour and Welfare (2005b), Fukawa (1998) | | | Insurance claim review |
| Ages 70–74 | 460 617 | Change by \pm 50% | | | | |
| Ages 75–79 | 549 284 | Change by \pm 50% | | | | |
| Terminal care cost, last year of life | | | | 3252 302 | Change by \pm 50% | |
| Ages 65–69 | 1394 690 | Change by \pm 50% | | | | |
| Ages 70–74 | 1980 653 | Change by \pm 50% | | | | |
| Ages 75–79 | 2361 923 | Change by \pm 50% | | | | |
| Ages 80+ | | | | | | |
| First year after diagnosis | | | Ministry of Health, Labour and Welfare (2005b), Fukawa (1998) | 961 181 | Change by \pm 50% | Insurance claim review |
| Yearly cost | | | | 185 151 | Change by \pm 50% | |
| Ages 80–84 | 576 290 | Change by \pm 50% | | | | |
| Ages 85–89 | 647 941 | Change by \pm 50% | | | | |
| Ages 90–94 | 557 429 | Change by \pm 50% | | | | |
| Ages 95–100 | 465 059 | Change by \pm 50% | | | | |
| Terminal care cost, last year of life | | | | 427 042 | Change by \pm 50% | |
| Ages 80–84 | 2478 049 | Change by \pm 50% | | | | |
| Ages 85–89 | 2786 147 | Change by \pm 50% | | | | |
| Ages 90–94 | 2396 943 | Change by \pm 50% | | | | |
| Ages 95–100 | 1999 754 | Change by \pm 50% | | | | |

Table 3 (Continued)

| | Diseases | | |
|---|----------------------------------|--------------------------------------|--------------------------------|
| | Base-case value | Range tested in sensitivity analysis | Source |
| Noninvasive breast cancer surgery, etc (DPC0900103x020xxx+ reimbursements by FFS) | 847 928 | Change by \pm 50% | Matsuda and Ishikawa (2003) |
| Endometrial cancer Total hysterectomy, etc (DPC 1200203x01x0xx+ reimbursements by FFS) | 1 183 839 | Change by \pm 50% | Matsuda and Ishikawa (2003) |
| Pulmonary embolism Total (Diagnosis) (Treatment) | 469 890 (52 350) (417 540) | Change by \pm 50% | Fuji et al (2005) |
| Cataract Surgery, etc (DPC 0201103x01x 000+reimbursements by FFS) | 309 120 | Change by \pm 50% | Matsuda and Ishikawa (2003) |
| Hip fracture Surgery, etc (DPC 1608003x02x0x+ reimbursements by FFS) | 1 553 195 | Change by \pm 50% | Matsuda and Ishikawa (2003) |

DPC: diagnosis procedure combination; FFS: fee for service.

a 5-year predicted breast cancer risk of 3.01–5.00%. The largest outcome gain in terms of QALYs, 0.058, is estimated among women with a history of AH starting at age 50.

Table 5 shows the results of cost-effectiveness analysis of therapeutic policy switch from tamoxifen to raloxifene.

Raloxifene is consistently superior to tamoxifen across presented risk classifications and starting ages of prophylaxis.

Costs

In the comparison between prophylaxis with tamoxifen vs no prophylaxis (Table 4), cost savings are estimated in higher risk classifications, among women with a history of LCIS or AH, starting at younger age. The largest saving, ¥367 901 (£1840), is estimated among women with a history of AH starting at age 35.

Between prophylaxis with raloxifene vs no prophylaxis, prophylaxes are found more costly. A cost saving of ¥10 387 (£52) is estimated among women with a history of AH starting at age 50.

When considering the therapeutic policy switch (Table 5), the use of raloxifene is consistently more costly than tamoxifen, as anticipated by the difference in price of agents.

Cost-effectiveness

There is a suggested criterion for cost-effectiveness in Japan (Ohkusa, 2003) to be ¥6000 000 (£30 000) for one QALY gain, and both Tables 4 and 5 report judgements with this criterion.

In the comparison between prophylaxis with tamoxifen vs no prophylaxis, favourable results, that is 'cost less and gain more' or cost-effective, are obtained in higher risk classifications starting at younger age. Those are: women with a history of AH regardless of starting age, women with a history of LCIS starting at age 35 and 50, and women with a 5-year predicted breast cancer risk of \geq 5.01% starting at age 35 and 50.

Similar results are found between prophylaxis with raloxifene vs no prophylaxis. Favourable results are: women with a history of

AH regardless of starting age, women with a history of LCIS starting at age 50, and women with a 5-year predicted breast cancer risk of \geq 5.01% starting at age 50.

As shown in Table 5, ICERs for the therapeutic policy switch of prophylactic agent from tamoxifen to raloxifene varies from ¥1839 670 per QALY (£9198 per QALY) to ¥6771 100 per QALY (£33 856 per QALY). The larger ICER is yet still close to the suggested criterion of ¥6000 000 per QALY (£30 000 per QALY).

Stability of cost-effectiveness

One-way sensitivity analyses produce similar results across the agents, the risk classifications and the ages of starting prophylaxis. Therefore, we draw a cost-effectiveness plane to show the comparison between prophylaxis with raloxifene vs no prophylaxis among three risk classifications as an example: women with a 5-year predicted breast cancer risk of \geq 5.01%, women with a history of LCIS, and women with a history of AH.

Figure 2 plots three base-case values and 306 results (102 changes of variables \times three different risk classifications). Line OA indicates the threshold of favourable ICER compared to the suggested criterion of ¥6000 000 (£30 000) for one QALY gain. Most results are plotted close to base-case value, which suggest the stability of our model. Results for women with a history of AH remain constantly favourable being cost saving or cost-effective by the change of variables except for one plot shown as in area B. However, several results for women with a 5-year predicted breast cancer risk of \geq 5.01% and for women with a history of LCIS cross the threshold line, the vertical axis or the horizontal axis from the base-case values. Three plots in area B and seven plots in area C indicate that results turn unfavourably, that is cost-ineffective or 'gain less', whereas plots in area D show that results become cost saving.

Our model is most sensitive to the utility weight for healthy state under chemoprevention, of which plots are drawn in area B. Its change to 0.79 turns incremental effectiveness into

Table 4 Results of cost-effectiveness analysis (1)

| | CoCost (¥) | | | Effectiveness (LYGs) | | | Effectiveness (QALYs) | | | Incremental cost-effectiveness ratio | |
|--|----------------|------------|-------------|----------------------|-----------|-------------|-----------------------|-----------|-------------|--------------------------------------|-----------------------|
| | No prophylaxis | Tamoxifen | Incremental | No prophylaxis | Tamoxifen | Incremental | No prophylaxis | Tamoxifen | Incremental | (¥/LYG) | (¥/QALY) |
| No prophylaxis vs prophylaxis with tamoxifen | | | | | | | | | | | |
| Five-year predicted breast cancer risk $\geq 1.66\%$ | | | | | | | | | | | |
| Starting at age 35 | 13 958 679 | 13 983 626 | 24 947 | 25 916 | 25 953 | 0.037 | 25 757 | 25 759 | 0.002 | 678 210 | 14247447 |
| Starting at age 50 | 17 630 814 | 17 751 353 | 120 538 | 22 168 | 22 167 | -0.001 | 22 040 | 22 000 | -0.040 | Cost more, gain less | Cost more, gain less |
| Starting at age 60 | 20 160 906 | 20 324 294 | 163 388 | 18 806 | 18 807 | 0.001 | 18 688 | 18 654 | -0.034 | 120849 008 | Cost more, gain less |
| Five-year predicted breast cancer risk 3.01–5.00% | | | | | | | | | | | |
| Starting at age 35 | 13 627 472 | 13 685 368 | 57 896 | 26 005 | 26 035 | 0.030 | 25 879 | 25 872 | -0.007 | 1 946 092 | Cost more, gain less |
| Starting at age 50 | 17 579 407 | 17 732 900 | 153 493 | 22 195 | 22 185 | -0.010 | 22 088 | 22 037 | -0.051 | Cost more, gain less | Cost more, gain less |
| Starting at age 60 | 20 251 937 | 20 444 141 | 192 203 | 18 808 | 18 797 | -0.011 | 18 718 | 18 666 | -0.052 | Cost more, gain less | Cost more, gain less |
| Five-year predicted breast cancer risk $\geq 5.01\%$ | | | | | | | | | | | |
| Starting at age 35 | 14 956 349 | 14 667 969 | -288 380 | 25 651 | 25 755 | 0.105 | 25 396 | 25 480 | 0.084 | Cost less, gain more | Cost less, gain more |
| Starting at age 50 | 17 867 146 | 17 800 766 | -66 379 | 22 049 | 22 096 | 0.047 | 21 832 | 21 854 | 0.022 | Cost less, gain more | Cost less, gain more |
| Starting at age 60 | 19 958 433 | 20 058 020 | 99 857 | 18 797 | 18 825 | 0.028 | 18 614 | 18 618 | 0.004 | 3548 049 | 26 648 821 |
| History of lobular carcinoma in situ | | | | | | | | | | | |
| Starting at age 35 | 14 908 314 | 14 717 649 | -190 665 | 25 663 | 25 747 | 0.083 | 25 414 | 25 472 | 0.058 | Cost less, gain more | Cost less, gain more |
| Starting at age 50 | 17 856 158 | 17 850 722 | -5 386 | 22 054 | 22 085 | 0.031 | 21 841 | 21 843 | 0.002 | Cost less, gain more | Cost less, gain more |
| Starting at age 60 | 19 968 466 | 20 093 211 | 124 745 | 18 798 | 18 815 | 0.017 | 18 618 | 18 606 | -0.011 | 7262 700 | Cost more, gain less |
| History of atypical hyperplasia | | | | | | | | | | | |
| Starting at age 35 | 14 687 003 | 14 319 102 | -367 901 | 25 722 | 25 844 | 0.122 | 25 493 | 25 598 | 0.105 | Cost less, gain more | Cost less, gain more |
| Starting at age 50 | 17 806 095 | 17 692 020 | -114 075 | 22 079 | 22 139 | 0.060 | 21 884 | 21 922 | 0.038 | Cost less, gain more | Cost less, gain more |
| Starting at age 60 | 20 015 243 | 20 096 731 | 81 488 | 18 800 | 18 837 | 0.037 | 18 635 | 18 651 | 0.016 | 2226 684 | 5234 647 ^a |
| No prophylaxis vs prophylaxis with raloxifene | | | | | | | | | | | |
| Five-year predicted breast cancer risk $\geq 1.66\%$ | | | | | | | | | | | |
| Starting at age 50 | 17 630 814 | 17 833 020 | 202 206 | 22 168 | 22 190 | 0.022 | 22 040 | 22 027 | -0.013 | 9256 382 | Cost more, gain less |
| Starting at age 60 | 20 160 906 | 20 427 386 | 266 480 | 18 806 | 18 822 | 0.016 | 18 688 | 18 670 | -0.018 | 16 806 286 | Cost more, gain less |
| Five-year predicted breast cancer risk 3.01–5.00% | | | | | | | | | | | |
| Starting at age 50 | 17 579 407 | 17 794 890 | 215 482 | 22 195 | 22 214 | 0.019 | 22 088 | 22 071 | -0.017 | 11 599 422 | Cost more, gain less |
| Starting at age 60 | 20 251 937 | 20 529 452 | 277 515 | 18 808 | 18 820 | 0.012 | 18 718 | 18 694 | -0.024 | 23 845 594 | Cost more, gain less |
| Five-year predicted breast cancer risk $\geq 5.01\%$ | | | | | | | | | | | |
| Starting at age 50 | 17 867 146 | 17 911 198 | 44 053 | 22 049 | 22 111 | 0.062 | 21 832 | 21 871 | 0.039 | 705 126 | 1123 880 ^a |
| Starting at age 60 | 19 958 433 | 20 161 888 | 203 455 | 18 797 | 18 839 | 0.042 | 18 614 | 18 633 | 0.019 | 4848 677 | 10 664 954 |
| History of lobular carcinoma in situ | | | | | | | | | | | |
| Starting at age 50 | 17 856 158 | 17 935 497 | 79 540 | 22 054 | 22 107 | 0.053 | 21 841 | 21 869 | 0.027 | 1496 425 | 2904 386 ^a |
| Starting at age 60 | 19 968 466 | 20 186 549 | 218 083 | 18 798 | 18 833 | 0.036 | 18 618 | 18 628 | 0.010 | 6133 167 | 21462 765 |
| History of atypical hyperplasia | | | | | | | | | | | |
| Starting at age 50 | 17 806 095 | 17 795 708 | -10 387 | 22 079 | 22 156 | 0.077 | 21 884 | 21 942 | 0.058 | Cost less, gain more | Cost less, gain more |
| Starting at age 60 | 20 015 243 | 20 198 328 | 183 085 | 18 800 | 18 852 | 0.052 | 18 635 | 18 668 | 0.033 | 3527 453 | 5570 154 ^a |

^aCost-effective when compared to a suggested criterion in Japan (Ohkusa, 2003) of ¥6000 000 for one QALY gain.

negative. Critical values to change the judgement are 0.98, which makes the ICERs of women with a 5-year predicted breast cancer risk of $\geq 5.01\%$ and woman with a history of LCIS cost-ineffective, and the value of 0.96 makes women with a

history of AH 'gain less'. The model is also sensitive to the discount rate, of which plot is drawn in area C. Its raise of 5.9 and 4.3% makes the ICERs of women with a 5-year predicted breast cancer risk of $\geq 5.01\%$ and women with a history of

Table 5 Results of cost-effectiveness analysis (2)

| | Cost (¥) | | Effectiveness (LYGs) | | | Effectiveness (QALYs) | | | Incremental cost-effectiveness ratio | | |
|--|------------|------------|----------------------|------------|-------------|-----------------------|------------|-------------|--------------------------------------|----------|-----------|
| | Tamoxifen | Raloxifene | Tamoxifen | Raloxifene | Incremental | Tamoxifen | Raloxifene | Incremental | (¥/LYG) | (¥/QALY) | |
| Prophylaxis with tamoxifen vs prophylaxis with raloxifene | | | | | | | | | | | |
| Five-year predicted breast cancer risk $\geq 1.66\%$ | | | | | | | | | | | |
| Starting at age 50 | 17 751 353 | 17 833 020 | 81 667 | 22 167 | 22 190 | 0.023 | 22 000 | 22 027 | 0.027 | 3501 723 | 3035 955* |
| Starting at age 60 | 20 324 294 | 20 427 386 | 103 093 | 18 807 | 18 822 | 0.015 | 18 654 | 18 670 | 0.016 | 7107 875 | 6364 920 |
| Five-year predicted breast cancer risk 3.01–5.00% | | | | | | | | | | | |
| Starting at age 50 | 17 732 900 | 17 794 890 | 61 990 | 22 185 | 22 214 | 0.029 | 22 037 | 22 071 | 0.034 | 2163 079 | 1839 670* |
| Starting at age 60 | 20 444 141 | 20 529 452 | 85 312 | 18 797 | 18 820 | 0.023 | 18 666 | 18 694 | 0.028 | 3741 906 | 3063 477* |
| Five-year predicted breast cancer risk $\geq 5.01\%$ | | | | | | | | | | | |
| Starting at age 50 | 17 800 766 | 17 911 198 | 110 432 | 22 096 | 22 111 | 0.015 | 21 854 | 21 871 | 0.017 | 7150 490 | 6542 190 |
| Starting at age 60 | 20 058 020 | 20 161 888 | 103 869 | 18 825 | 18 839 | 0.014 | 18 618 | 18 633 | 0.015 | 7476 332 | 6771 100 |
| History of lobular carcinoma in situ | | | | | | | | | | | |
| Starting at age 50 | 17 850 772 | 17 935 697 | 84 925 | 22 085 | 22 107 | 0.022 | 21 843 | 21 869 | 0.025 | 3846 426 | 3359 650* |
| Starting at age 60 | 20 093 211 | 20 186 549 | 93 338 | 18 815 | 18 833 | 0.018 | 18 606 | 18 628 | 0.022 | 5064 724 | 4311 015* |
| History of atypical hyperplasia | | | | | | | | | | | |
| Starting at age 50 | 17 692 020 | 17 795 708 | 103 688 | 22 139 | 22 156 | 0.018 | 21 922 | 21 942 | 0.019 | 5922 294 | 5320 037* |
| Starting at age 60 | 20 096 731 | 20 198 328 | 101 598 | 18 837 | 18 852 | 0.015 | 18 651 | 18 668 | 0.017 | 6637 332 | 5872 017* |

*Cost-effective when compared to a suggested criterion in Japan (Ohkusa, 2003) of ¥6000 000 for one QALY gain.

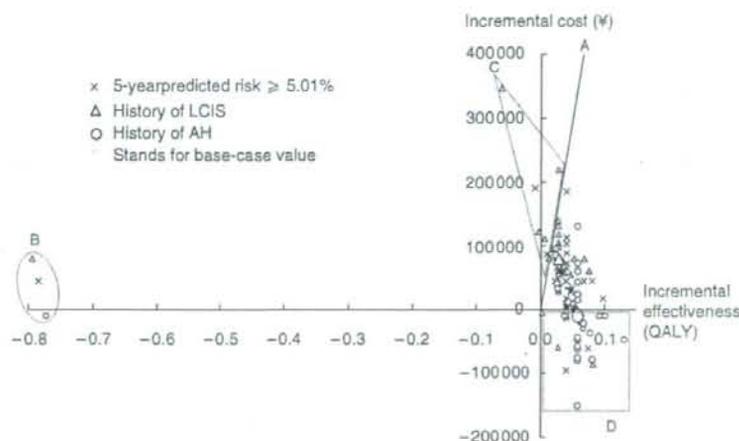


Figure 2 Illustration of key results of sensitivity analyses: prophylaxis with raloxifene vs no prophylaxis starting at age 50.

LCIS cost-ineffective, respectively. The cost of chemoprevention is also influential to the results, of which results are shown in areas C and D. A price increase of more than 30% for raloxifene makes the ICER of women with a history of LCIS cost-ineffective, whereas a price decrease of more than 16 or 29% make the results for women with a 5-year predicted breast cancer risk of $\geq 5.01\%$ and women with a history of LCIS cost saving, respectively. Changes of the probabilities of transition to invasive breast cancer, endometrial cancer, and hip fracture are also plotted in areas C and D. Raising the probability of invasive breast cancer beyond 0.00710 and 0.00683 makes the ICERs of women with a 5-year predicted breast cancer risk of $\geq 5.01\%$ and women with a history of LCIS cost-ineffective, whereas lowering to less than 0.00456 or 0.00436 make the results for women with a 5-year predicted breast cancer risk of $\geq 5.01\%$ and women with a history of LCIS cost saving, respectively. Raising the probability of endo-

metrial cancer beyond 0.00369 and 0.00271 makes the ICERs of women with a 5-year predicted breast cancer risk of $\geq 5.01\%$ and women with a history of LCIS cost-ineffective, respectively. Raising probability of hip fracture beyond 0.00098 makes the results for women with a history of LCIS cost saving. The other plots in area C reflect a raise of utility weight for invasive breast cancer after the second year.

Prolonging risk reduction effect of tamoxifen from 5 to 10 and 15 years without any risk increase of adverse events after the completion of prophylaxis brings more favourable results. For example, the effect of 10 years results in 'cost less and gain more' for every risk classification starting at age 35, whereas the effect of 15 years makes no change in the results of 'cost more and gain less' among women with a 5-year predicted breast cancer risk of $\geq 1.66\%$ starting at age 50 and 60.

DISCUSSION

We conduct a cost-effectiveness analysis of SERMs as prophylactic agents against breast cancer among high-risk women by making comparisons between *status quo* in Japan, without prophylaxis, and hypothetical practise, with prophylaxis, by the agent (tamoxifen and raloxifene), the risk classification, and the age of starting prophylaxis.

We find that prophylaxis with tamoxifen results in 'cost less and gain more' among extremely high-risk women such as those with a 5-year predicted breast cancer risk of $\geq 5.01\%$, those with a history of LCIS, and those with a history of AH starting at age 35 and 50. Prophylaxis with raloxifene is also found 'cost less and gain more' for women with a history of AH starting at age 50. The younger the age of starting prophylaxis, the more the cost saving and outcome gain. We also find that prophylaxis with tamoxifen for women with a history of AH starting at age 60 results in favourable ICER compared to the suggested criterion of ¥6000 000 (£30 000) for one QALY gain. Prophylaxis with raloxifene is also found cost-effective for women with a 5-year predicted breast cancer risk of $\geq 5.01\%$ starting at age 50, those with a history of LCIS starting at age 50 and those with a history of AH starting at age 60. The younger the age of starting prophylaxis, the more favourable the ICER. Within the same risk classification and starting age, raloxifene tends to gain more and cost more compared to tamoxifen. On the contrary, we also find that prophylaxes with tamoxifen or raloxifene for women with a 5-year predicted breast cancer risk of $\leq 5.00\%$ tend to result in 'cost more and gain less'.

These findings are similar to the previous economic evaluations of chemoprevention of breast cancer with tamoxifen including analyses of risk level differences such as Noe *et al* (1999); Grann *et al* (2000); Hershman *et al* (2002); Melnikow *et al* (2006), although these studies are carried out under the US health system.

Our findings suggest that introduction of chemoprevention with SERMs targeting extremely high-risk women in Japan can be justifiable as an efficient use of finite health-care resources, possibly contributing to cost containment. The cost saving results suggest chemoprevention not only cost-effective but also affordable. Taking the superiority of raloxifene in outcome gain and the difference in indication into account, it is recommendable to administer tamoxifen for premenopausal women and raloxifene for postmenopausal women.

Our economic model is found sensitive to the utility weight for healthy state under chemoprevention, the discount rate and the cost of chemoprevention, in addition to the probabilities of transition to invasive breast cancer, endometrial cancer, or hip fracture. This is anticipated because these variables are supposed to influence the cost-effectiveness of preventive services. We think that our economic model succeeds in explaining the context under consideration.

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We also analysed the cost-effectiveness of therapeutic policy switch of agent, tamoxifen to raloxifene among postmenopausal women, although this does not depict any marginal change in Japan. All simulated ICERs by risk classifications starting at age 50 and 60 fall in a favourable level. Due caution is needed in transferring these findings from our Japanese model to other health system (Drummond and Pang, 2001), but it implies that the administration of raloxifene instead of tamoxifen for postmenopausal high-risk women could be economically acceptable in developed countries where chemoprevention with tamoxifen is already in practise.

There are a couple of points to consider when interpreting our results. Our model depends on clinical evidence established in the United States by P-1 and P-2 trial. Composition of ethnicity and life styles of participating women are different from those of Japanese women. This also relates to another point, that is the validity of the 5-year risk prediction model defining high-risk women. As already mentioned in Methods section, it is based on Gail *et al* model 2 (Gail and Costantino, 2001), which has been validated for white women (Rockhill *et al*, 2001) and African American women (Gail *et al*, 2007) only. Our approach is acceptable as to these points, as the results of P-1 and P-2 trial are the best available evidence to date for the objectives of this study, and similar risk factors to Gail *et al* model 2 are identified in a model of individualised probability of developing breast cancer for Japanese women (Ueda *et al*, 2003), and the function of ethnic difference in developing breast cancer is reported as small (Chen *et al*, 2004). Our model also depends on utility weights reported from Western countries, as none of those from Japan are available. However, our findings of consistent outcomes in terms of LYGs offer reasonable conclusions.

In summary, this study suggests that chemoprevention of breast cancer with SERMs targeting high-risk women such as a 5-year predicted breast cancer risk of $\geq 5.01\%$, women with a history of LCIS, and women with a history of AH, clears the hurdles of introducing new intervention by means of cost-effectiveness and affordability, with best available evidence. Although further studies and policy formulations are necessary about breast cancer chemoprevention in Japan, this study also implies that the administration of raloxifene instead of tamoxifen may be cost-effective under the context of developed countries where chemoprevention with tamoxifen has already been adopted.

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■ 原著 ■

乳頭分泌液中 CEA 測定におけるイムノクロマトグラフィー
(ICGA) 法と酵素免疫測定 (EIA) 法の比較検討

多根井 智紀 増田 慎三 石飛 真人 徳田 由紀子
吉田 謙 真能 正幸 竹田 雅司 辻 伸利 政

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原著

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乳頭分泌液中CEA測定におけるイムノクロマトグラフィー
(ICGA) 法と酵素免疫測定 (EIA) 法の比較検討多根井 智紀^{*1,6} 増田 慎三^{*1} 石飛 真人^{*1,4} 徳田 由紀子^{*2}
吉田 謙^{*2} 真能 正幸^{*3} 竹田 雅司^{*3,5} 辻 伸利 政^{*1}

Comparison of "Immunochromatographic Assay (ICGA)" with "Enzyme Immunoassay (EIA)" For Measurement of CEA in Abnormal Nipple Discharge : Tanei T^{*1,6}, Masuda N^{*1}, Ishitobi M^{*1,4}, Tokuda Y^{*2}, Yoshida K^{*2}, Mano M^{*3}, Takeda M^{*3,5} and Tsujinaka T^{*1} (*1:Department of Surgery, *2:Department of Radiology, *3:Department of Pathology, Osaka National Hospital, *4:Department of Surgery Osaka Medical Center for Cancer and Cardiovascular Diseases, *5:Department of Pathology, Yao City Hospital, *6:Department of Surgical Oncology, Osaka University Graduate School of Medicine)

For the detection of early breast cancer, such as ductal carcinoma in situ (DCIS), it is important to measure CEA levels in abnormal nipple discharge. We compared Immunochromatographic assay "Lana Mammo Card CEA[®]" with enzyme immunoassay "MAMMOTEC[®]" for the measurement of CEA levels in abnormal nipple discharge from 20 patients (13 malignant cases, 7 benign cases). When the cut off value of CEA was set at 400ng/ml for each kit, the sensitivity, specificity and accuracy were 100%, 71.4% and 90% using the "Immunochromatographic assay (ICGA)". They were 92.3%, 71.4% and 85.0% using the "enzyme immunoassay (EIA)". The value obtained by radio immuno assay (RIA) or related with those obtained by "Immunochromatographic assay (ICGA)" and "enzyme immunoassay (EIA)". We should recommend a carefully considered approach combined with non-invasive tools. Using the "Immunochromatographic assay (ICGA)", the CEA level can be obtained by the easier single-step manipulation within 15 minutes. Therefore, a follow-up investigations can be planned during the initial medical examination.

Key words : Abnormal nipple discharge, Carcinoembryonic antigen (CEA), Immunochromatographic assay (ICGA), Enzyme immunoassay (EIA)

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はじめに

乳頭異常分泌は早期乳癌 (DCIS) や乳管内乳頭腫等の増殖性病変の発見動機として重要な症状であり、的確な診断と治療を行うことが求められる。乳頭異常分泌のスクリーニング法として、潜血反

応、乳頭分泌細胞診があるが、いずれも診断の精度が低い為、稲治らは診断の一助として乳頭分泌液中の癌胎児性抗原 (Carcinoembryonic antigen : CEA) を用いる測定法を考案した¹⁾。これは乳癌より乳管中に高濃度のCEAが産出されることを利用したものであり、測定法としては、酵素免疫測定法 (enzyme immunoassay ; EIA) を用いたマンモテック[®] が以前より知られている。また、最近、イムノクロマトグラフィー法 (immuno-chromatographic assay ; ICGA) を用いた簡易かつ迅速な測定キットであるラナマンモカードCEA[®] が販売され厚生省にて認可された。今

*1 国立病院機構大阪医療センター-外科

*2 国立病院機構大阪医療センター-放射線科

*3 国立病院機構大阪医療センター-臨床検査科

*4 大阪府立成人病センター-乳腺内分泌外科

*5 八尾市立病院病理診断科

*6 現：大阪大学大学院乳腺内分泌外科

回、われわれはICGA法(ラナマンモカードCEA[®])とEIA法(マンモテック[®])の2つの測定法について比較検討したので報告する。また、RIA法によるCEA定量の測定を行い、実測値との相関も検討してみた。

1. 対象と方法

2004年7月～2005年10月までの間に、乳頭異常分泌を主訴に受診した患者のうち、根治手術にて病理診断を得た20症例について検討した。全例女性で平均年齢は58歳、28歳から77歳までであり、20症例の内訳は乳癌13例(浸潤癌6例、非浸潤癌7例)、良性疾患7例(乳頭腫3例、乳腺症4例)であった。全例に対して乳頭分泌液細胞診とEIA法・ICGA法による分泌液中CEAの測定を行った。また、採取した乳頭分泌液の検体は保存し、一括してSRL社にてRIA法によるCEA定量の測定を行い、この測定値を対照にして2つの測定法(EIA法とICGA法)の正確性を比較した。この研究は当院の受託研究審査委員会で承認の上、患者から同意を取得して実施した。

1) EIA法による分泌液中CEA測定(マンモテック[®])の原理と方法

マンモテック[®]は、プラスチックフィルムに不溶化した抗CEA抗体と酵素標識抗CEA抗体を用いたmicrodot-limmunobinding assay(サンドイッチ型酵素免疫測定法)である。検体と標識抗体を添加した後、30分間乾燥させ、洗浄が必要であり、基質添加後さらに60分乾燥させる²⁾。判定は目視による着色強度により半定量的に行い、測定には2時間以上を要するキットの推奨どおりの方法にて行った。

2) ICGA法による分泌液中CEA測定(ラナマンモカードCEA[®])の原理と方法

ラナマンモカードCEA[®]は、イムノクロマトグラフィを用いて、簡便で短時間に測定できるように考案されたキットであり、抗CEAモノクローナル抗体結合金コロイド粒子が含まれ、判定部位にも抗CEA抗体が固相化されている。分泌中のCEAがコロイド標識されたモノクローナル抗

CEA抗体と反応して免疫複合体を形成し、免疫複合体は展開液により移動した後、判定部位の固相化抗CEA抗体に捉えられ、15分以上静置している間にラインを形成し目視にて確認が可能になる³⁾。ラナマンモカードCEA[®]はこれを利用して半定量的に測定したものであり、分泌液(検体)中CEAが存在しない場合には、コロイド粒子が移動するのみにてラインは形成されない(図1)。

2. 結果

1) EIA法とICGA法との測定値および正診率

乳頭異常分泌にて根治手術を施行した20例(乳癌13例、良性疾患7例)の乳頭分泌液に対してICGA法とEIA法を測定し、2つのCEA測定値と摘出標本の病理組織診断を比較した。診断精度は、乳頭分泌CEA研究会に準じてカット・オフ値を400ng/ml以上を陽性としたところ、病理結果にて乳癌と診断された13例すべてにおいてICGA法にて高値を認めた(図2)。また、EIA法は乳癌13例のうち12例に高値を認めたが、浸潤癌(硬癌)1例において偽陰性(400ng/ml未満)であった。また、この偽陰性1例については、RIA法によるCEA定量での実測値も280ng/mlと低値であり、元来、乳管内成分からのCEA産生が少ない乳癌症例であったと推測される。

偽陽性については、乳腺症2例がEIA法とICGA法ともに陽性であった。診断精度は、ICGA法の診断精度は感度100%(13/13)、特異度71.4%(5/7)、正診率90%(18/20)、陽性適中率86.7%(13/15)であった。また、EIA法は感度92.3%(12/13)、特異度71.4%(5/7)、正診率85.0%(17/20)、陽性適中率85.7%(12/14)であり、ICGA法とEIA法はともに同等の診断精度であった(表1)。

2) RIA法によるCEA定量値との比較

今回、われわれはEIA法、及びICGA法の測定値に対する正確性を判定する為、20例の乳頭分泌液に対して、RIA法によるCEA定量を測定して(SRL社に委託)、この定量値を対照にして両者の測定値との比較を行った。2つの測定法は図3に示すようにRIA法によるCEA定量値との間に相関

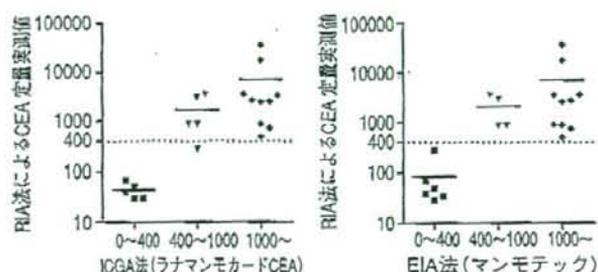


図3 RIA法によるCEA定量実測値との比較

表1 イムノクロマトグラフィー法 (ラナマンモカードCEA®) と酵素免疫測定法 (マンモテック®) の診断精度

| | 感度 | 特異度 | 正診率 | 陽性過中率 |
|--------------|---------------|-------------|-------------|---------------|
| ラナマンモカードCEA® | 100% (13/13) | 71.4% (5/7) | 90% (18/20) | 86.7% (13/15) |
| マンモテック® | 92.3% (12/13) | 71.4% (5/7) | 85% (17/20) | 85.7% (12/14) |

表2 イムノクロマトグラフィー法 (ラナマンモカードCEA®) と酵素免疫測定法 (マンモテック®) の相関

| | | マンモテックCEA® | | |
|------------------|-----------|------------|-----------|--------|
| | | 0~400 | ≤400~1000 | ≤1000~ |
| ラママンモ カードCEA® | 0~400 | 5 | | |
| | ≤400~1000 | 1 | 2 | 2 |
| | ≤1000~ | | 2 | 8 |

表3 乳頭異常分泌液中CEA測定法による乳癌診断精度

| 年 | CEA測定法 | cut off (ng/ml) | 症例数 | 感度 | 特異度 |
|--------------|-------------------|-----------------|-----|-----|------|
| Inaji H | 1987 エルモテックCEA | 100 | 30 | 86% | 89% |
| Mori T | 1992 マンモテック® | 400 | 54 | 76% | 79% |
| Nishiguchi T | 1992 マンモテック® | 400 | 16 | 60% | 75% |
| Yayoi E | 1994 エルモテックCEA | 600 | 60 | 77% | 100% |
| Nishi T | 2003 マンモテック® | 400 | 47 | 79% | 93% |
| Nishi T | 2003 ラナマンモカードCEA® | 400 | 47 | 68% | 96% |

が可能であるとされている。乳頭異常分泌に対するスクリーニング法としては、潜血反応、乳頭分泌細胞診があるが、感度において十分な成績ではなく⁹⁾、乳管内視鏡や乳管造影は検査に時間と労力を伴う為、多くの症例に行うことは不可能であり、分泌物の細胞診は精度が低い。それらを補う為、無侵襲にて測定可能な乳頭分泌液中の腫瘍マーカーとしてCEAが注目され⁹⁾、1987年、稲治らは乳頭分泌液中のCEA量の測定(エルモテック法)を提案し、非浸潤性乳管癌と良性疾患(乳管内乳頭腫)との鑑別に有用であることが報告した¹⁷⁾。現在、酵素免疫測定法(EIA法)を用いたマンモテ

ック®とイムノクロマトグラフィー法(ICGA法)を用いたラナマンモカードCEA®が、共に市販されている。表3のように、森らは、血性乳頭異常分泌52例に対してマンモテック®の測定を行い(カット・オフ値:400ng/ml)、マンモテック®により感度76%、特異度79%と、良好な測定結果を報告している⁹⁾。また、海外においてはimmunoenzymometric assay kit, AIA-PACK CEAを用いて優れた診断精度が報告されている⁹⁾。

ICGA法は、弥生らの考案により製品化された、外来精査中の短時間で測定結果が得られる簡便な方法で、同施設の西らが乳頭異常分泌115例(うち

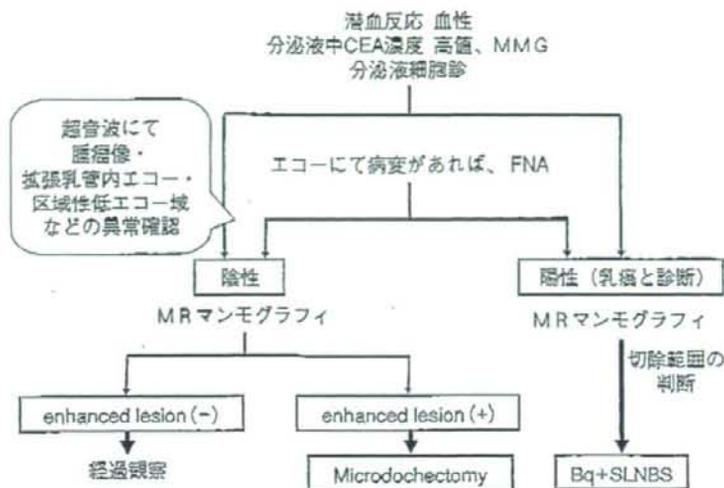


図4 当院における腫瘍非触知の乳頭異常分泌症の診断・治療方針

確定診断のついた47例)について感度68%, 特異度96%, 正診率85%であり, EIA法と同等の診断精度を有すると報告している⁹⁾。しかし, 同じ症例においてEIA法とICGA法との比較をした検討は今回が初めてである。ICGA法の診断能は, 感度100%, 特異度71.4%, 正診率90%であり, EIA法については, 感度92.3%, 特異度71.4%, 正診率85.0%であり, 両者ともほぼ同等の結果が得られた。2つの測定法は, 乳頭分泌液の性状(膿性, 粘性)や乳管内乳頭腫などの多寡な乳管内成分が一因となって偽陽性を生ずるとされているが, 今回の偽陽性を認めた2症例は, 共に乳腺症を病理診断とする同じ症例であり(図2), 性状も明らかな特徴を認めなかった。また, 西らは, ICGA法について, RIA法により測定されたCEA実量値とのよく相関すると報告されている¹⁰⁾。

当院での乳頭異常分泌の診断・治療シエマを示すが(図4), 乳頭異常分泌がある場合, MMGや乳腺超音波検査など画像検査に加えて, 分泌液の潜血反応検査, 分泌液細胞診, 分泌液中CEA濃度の測定を施行している。また, 乳腺超音波にて腫瘍像や低エコー域などの異常を確認した場合は体表より穿刺吸引細胞診(FNA)を行う。そして, 分泌液細胞診にて陽性を認めた場合や, FNAにて陽性を認めた場合は, 乳癌に準じた治療をしている。現在, 乳頭異常分泌症に対するマーカーとして, CEAが臨床応用されているが, その他のマ-

ーカーとしてAngiogenic growth factor (FGF-2, IGFBP-3)やKallikreins (PSA), Her2/neu protein, Urokinase-type plasminogen activator (uPA)などが挙げられ¹²⁾, 今後, さらに診断精度の高いマーカーや測定方法として期待される。

今回, ICGA法による乳頭分泌液中のCEA測定は従来のEIA法による測定と同等の良好な診断精度をもつことが明らかになった。また, EIA法(マンモテック[®])は測定時間が2時間以上と長く, 手技もやや煩雑であるのに対し, ICGA法(ラナマンモカードCEA[®])は約15分で測定でき, 手技も単純(single-step manipulation)である。したがって外来診療中に施行することが可能であり, 乳頭異常分泌症例の診断, 治療方針決定に有効な方法であると考えられる。

まとめ

イムノクロマトグラフィー(ICGA)法は従来の酵素免疫測定(EIA)法と同等の診断精度であった。ICGA法はsingle step manipulationにて簡易かつ迅速に測定することが可能であり, 乳頭異常分泌症例の診断, 治療方針決定に有効な方法である。

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