generations (<70 and 70-79) (Supplementary Table 1) and in the urban community (Supplementary Table 2).

DISCUSSION

The present study revealed that the prevalence of radiographic lumbar spondylosis with KL≥2 and KL≥3 in the elderly (≥ 60 years) was 75.8 and 50.4% respectively, and that of low back pain was 28.8% in the overall population. Although the KL≥2 spondylosis was more prevalent in men (84.1%) than in women (70.7%), the KL≥2 spondylosis and low back pain were more prevalent in women. This study also showed that KL=2 spondylosis was not significantly associated with low back pain compared to KL=0 or 1, while KL≥3 spondylosis was related to the pain only in women.

Most previous epidemiologic studies on lumbar spondylosis were focused on the middle-aged or younger populations, reporting the prevalence to be 46.5-83.7% (4,6-8,10,11). Our previous small-scale study on a younger population has shown that to be 76.3 and 37.4% (9). Interestingly, the subjects were living in a mountainous area in Japan, which was shown to have a lower risk for spondylosis in the present study. The variability may therefore be due to the differences of age, community, the sample size and ethnic variation. In fact, a study on the elderly (≥ 65 years) showed that the prevalence of KL≥2 spondylosis was 84.8 and 70.6%, similar to the present results, although in a relatively small number of subjects (5). We have reported different prevalence of lumbar spondylosis in Japan and the United Kingdom by a small-scale comparative study (9), which may in part relate to ethnic variation. It should be noticed that this is the first population-based study that investigated the age-related prevalence of lumbar spondylosis in the elderly. Although KL≥2 and KL≥3 spondylosis tended to increase with age, significant difference was detected between the sixties and the seventies, but not thereafter. However, this cross-sectional analysis, of course, does not lead to the conclusion that individual lumbar spondylosis hardly progresses after 80 years. Since the ROAD study is a prospective cohort study for more than ten years, the follow-up data will clarify the progression with aging. Furthermore, there was a difference of the prevalence between urban and mountainous communities. Considering that lumbar spondylosis is a common disease whose progression is governed by environmental and genetic factors, the regional difference is inevitable, as previously reported (6). Although age and obesity are known to be representative risk factors for lumbar spondylosis (2), the difference between communities in the present study was significant even after adjustment for age and BMI, indicating the involvement of other factors. Here again, further longitudinal survey in the ROAD study that collects database including detailed environmental and genomic information will elucidate the underlying backgrounds.

Interestingly, KL≥2 spondylosis was more prevalent in men than in women, while KL≥3 spondylosis was more prevalent in women. We and others also have reported that osteophytosis of lumbar spine was more common in men than in women (8,9), while disc space narrowing was prevalent in women (9). Based on the definition of the KL grading (12), the discrepancy may be due to distinct etiologic mechanisms between osteophyte formation and disc space narrowing. A cross-sectional study which investigated the extent, prevalence and distribution of spinal spondylosis in women also showed that osteophytosis and disc space narrowing were significantly correlated, but each predicted only 19% of the variation in the other (11). A previous prospective study in knee joints using a famous cohort, the Chingford Study, has

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reported that there was no association between osteophyte formation and joint space narrowing (14). A recent study using quantitative magnetic resonance imaging (MRI) in knee joints has also shown that osteophyte formation was unrelated to cartilage loss (15). Furthermore, in an experimental mouse knee OA model, we have identified a cartilage specific molecule carminerin that induces only osteophyte formation without affecting cartilage degeneration during the OA progression (16,17). Further clinical and basic research will disclose the distinct backgrounds of these two representative OA features.

Symptomatic low back pain was solely associated with KL≥3 spondylosis in women, but not KL≥2 spondylosis in either gender or KL≥3 spondylosis in men. Considering the definition of KL grading again, this may suggest that disc space narrowing, but not osteophytosis, of the lumbar spine contributes to low back pain in consistency with previous reports (18). Difference of the association between genders might possibly be dependent on muscle strength to compensate the spinal instability due to disc space narrowing, since men are known to have greater muscle strength than women in all decades (19). However, approximately 30% of participants without definite radiographic lumbar spondylosis (KL=0 or 1) actually had low back pain, and the odds ratio of KL≥3 spondylosis for the pain was 1.44 and 1.80 in men and women, respectively, which is much lower than the previously reported odds ratio 8.5 of KL≥3 osteoarthritis in the knee joint for knee pain (20). This may be because low back pain arises from a number of disorders other than disc space narrowing, like nociceptive stimuli, inflammation, muscle weakness, and abnormal load on muscle, ligament or capsular tissues (21). Indeed, disc degeneration was reported to be detected by MRI in at least one lumbar level in all but one of asymptomatic volunteers in a 60-80 age group (22). Furthermore, pain is also influenced by psychological status like depression, since significant association between low back pain and depression has been confirmed in many longitudinal studies (23-25). A recent psychophysical study has shown that anxiety was linked to self-reported and induced low back pain for men, but not for women (26). This might be an alternative reason for lower association between radiographic spondylosis and low back pain in men.

There are several limitations in the present study. First, prevalence figures in this study using a large-scale population based sample of the elderly may be generalized to the Japanese population. However, this study investigated elderly participants who lived independently, not those who lived in institutional settings. Therefore, the calculated prevalence may be underestimated. Second, the definition of low back pain in the present study did not determine the severity. The association of lumbar spondylosis with the severity of low back pain could not be examined in this study. Third, analyses did not include facet joint osteoarthritis nor vertebral fracture, which would likely be associated with low back pain. This is the next task in the ROAD study to be investigated. Fourth, since the KL system emphasizes osteophytosis, it is unclear how to handle lumbar spondylosis with disc space narrowing but no osteophytosis. Since quantitative MRI is still too laborious and expensive to perform in general clinical practice, we are now developing a computer-aided diagnostic program which enables a fully automatic measurement of major features of lumbar spondylosis including disc space narrowing and osteophytosis on plain radiographs.

In conclusion, the present cross-sectional study using a large-scale population from the ROAD study revealed a high prevalence of radiographic lumbar spondylosis in the elderly. The prevalence differed to some extent by age, gender, and community. Gender seems to be distinctly associated with KL≥2 and KL≥3 lumbar spondylosis, and disc space narrowing with

or without osteophytosis in women may be a risk factor for low back pain. Further progress, along with continued longitudinal survey in the ROAD study, will elucidate the environmental and genetic backgrounds of lumbar spondylosis and its relation with low back pain.

Conflict of interest/disclosure
There are no conflicts of interest or disclosures regarding the present manuscript.

Figure Legends
Figure 1. The percentage of subjects with low back pain according to the Kellgren/Lawrence grade in the overall population, and urban, mountainous and seacoast communities.

REFERENCES

1. **Hadjipavlou AG**, Simmons JW, Pope MH, Necessary JT, Goel VK. Pathomechanics and clinical relevance of disc degeneration and annular tear: A point-of-view review. *Am J Orthop* 1999;28:561–71.

2. **Emery SE**, Ringus VM. Osteoarthritis of the spine. In: Moskowitz RW, Altman RD, Hochberg MC, Buckwalter JA, Goldberg VM, eds. *Osteoarthritis*: Diagnosis and Medical/Surgical Management. 4th ed. Philadelphia: Lippincott Williams & Wilkins; 2007:427-52.

3. **Waddell G**. *The back pain revolution*. Edinburgh, Scotland: Churchill Livingstone; 1998:119-34.

4. **Kellgren JH**, Lawrence JS. Osteo-arthrosis and disk degeneration in an urban population. *Ann Rheum Dis* 1958;17:388-97.

5. **Lawrence JS**. Disc degeneration. Its frequency and relationship to symptoms. *Ann Rheum Dis* 1969;28:121-38.

6. **van Saase JL**, van Romunde LK, Cats A, Vandenbroucke JP, Valkenburg HA. Epidemiology of osteoarthritis: Zoetermeer survey. Comparison of radiological osteoarthritis in a Dutch population with that in 10 other populations. *Ann Rheum Dis* 1989;48:271-80.

7. **Symmons DP**, van Hemert AM, Vandenbroucke JP, Valkenburg HA. A longitudinal study of back pain and radiological changes in the lumbar spines of middle aged women. II. Radiographic findings. *Ann Rheum Dis* 1991;50:162-6.

8. **O'Neill TW**, McCloskey EV, Kanis JA, Bhalla AK, Reeve J, Reid DM, et al. The distribution, determinants, and clinical correlates of vertebral osteophytosis: a population based survey. *J Rheumatol* 1999;26:842-8.

9. **Yoshimura N**, Dennison E, Wilman C, Hashimoto T, Cooper C. Epidemiology of chronic disc degeneration and osteoarthritis of the lumbar spine in Britain and Japan: a comparative study. *J Rheumatol* 2000;27:429-33.

10. **Hassett G**, Hart DJ, Manek NJ, Doyle DV, Spector TD. Risk factors for progression of lumbar spine disc degeneration: the Chingford Study. *Arthritis Rheum* 2003;48:3112-7.

11. **Kramer PA**. Prevalence and distribution of spinal osteoarthritis in women. *Spine* 2006;31:2843-8.

12. **Kellgren JH**, Lawrence JS, eds. *The epidemiology of chronic rheumatism*: atlas of standard radiographs of arthritis. Oxford: Blackwell Scientific; 1963.

13. **Shimada H**, Lord SR, Yoshida H, Kim H, Suzuki T. Predictors of cessation of regular leisure-time physical activity in community-dwelling elderly people. *Gerontology* 2007;53:293-7.

14. **Hart DJ**, Doyle DV, Spector TD. Incidence and risk factors for radiographic knee osteoarthritis in middle-aged women: the Chingford Study. *Arthritis Rheum* 1999;42:17-24.

15. **Jones G**, Ding C, Scott F, Glisson M, Cicuttini F. Early radiographic osteoarthritis is associated with substantial changes in cartilage volume and tibial bone surface area in both males and females. *Osteoarthritis Cartilage* 2004;12:169-74.

16. **Yamada T**, Kawano H, Koshizuka Y, Fukuda T, Yoshimura K, Kamekura S, et al. Carminerin contributes to chondrocyte calcification during endochondral ossification. *Nat Med* 2006;12:665-70.

17. **Kamekura S**, Kawasaki Y, Hoshi K, Shimoaka T, Chikuda H, Maruyama Z, et al. Contribution of runt-related transcription factor 2 to the pathogenesis of osteoarthritis in mice after induction of knee joint instability. *Arthritis Rheum* 2006;54:2462-70.

18. **Frymoyer JW**, Pope MH, Clements JH, Wilder DG, MacPherson B, Ashikaga T. Risk factors in low-back pain. An epidemiological survey. *J Bone Joint Surg Am* 1983;**65**:213-8.
19. **Sinaki M**, Nwaogwugwu NC, Phillips BE, Mokri MP. Effect of gender, age, and anthropometry on axial and appendicular muscle strength. *Am J Phys Med Rehabil* 2001;**80**:330-8.
20. **Felson DT**, Naimark A, Anderson J, Kazis L, Castelli W, Meenan RF. The prevalence of knee osteoarthritis in the elderly. The Framingham Osteoarthritis Study. *Arthritis Rheum* 1987;**30**:914-8.
21. **Parkkola R**, Rytokoski U, Kormano M. Magnetic resonance imaging of the discs and trunk muscles in patients with chronic low back pain and healthy control subjects. *Spine* 1993;**18**:830-6.
22. **Boden SD**, Davis DO, Dina TS, Patronas NJ, Wiesel SW. Abnormal magnetic-resonance scans of the lumbar spine in asymptomatic subjects. A prospective investigation. *J Bone Joint Surg Am* 1990;**72**:403-8.
23. **Larson SL**, Clark MR, Eaton WW. Depressive disorder as a long-term antecedent risk factor for incident back pain: a 13-year follow-up study from the Baltimore Epidemiological Catchment Area sample. *Psychol Med* 2004;**34**:211-9.
24. **Sarzi-Puttini P**, Atzeni F, Fumagalli M, Capsoni F, Carrabba M. Osteoarthritis of the spine. *Semin Arthritis Rheum* 2005;**34**:38-43.
25. **Hicks GE**, Simonsick EM, Harris TB, Newman AB, Weiner DK, Nevitt MA, et al. Cross-sectional associations between trunk muscle composition, back pain, and physical function in the health, aging and body composition study. *J Gerontol A Biol Sci Med Sci* 2005;**60**:882-7.
26. **Robinson ME**, Dannecker EA, George SZ, Otis J, Atchison JW, Fillingim RB. Sex differences in the associations among psychological factors and pain report: a novel psychophysical study of patients with chronic low back pain. *J Pain* 2005;**6**:463-70.

Table 1. Characteristics of participants

	Men				Women			
	Overall	Urban	Mountainous	Seacoast	Overall	Urban	Mountainous	Seacoas
No. of subjects	818	397	266	155	1,470	742	434	294
Age, years	74.7 ± 6.1	77.3 ± 4.1	72.1 ± 6.2	72.7 ± 7.4	74.0 ± 6.4*	76.4 ± 4.8*	72.1 ± 7.1	70.9 ± 6.8
Height, cm	161.3 ± 6.3	161.2 ± 5.9	160.3 ± 6.6	163.0 ± 6.1	148.6 ± 6.2	148.6 ± 5.8	146.8 ± 6.4	151.2 ± 5
Weight, kg	60.1 ± 9.9	59.8 ± 8.3	59.3 ± 11.4	62.2 ± 10.6	50.9 ± 9.0	50.7 ± 8.4	49.8 ± 9.8	53.1 ± 8.
BMI, kg/m^2	23.0 ± 3.2	23.0 ± 2.7	23.0 ± 3.8	23.3 ± 3.3	23.0 ± 3.7	22.9 ± 3.4	23.1 ± 4.2	23.2 ± 3.
Current smoker, %	24.6	25.2	26.3	20.0	3.1*	3.1*	4.4*	1.0*
Current drinker, %	61.2	60.0	67.0	54.8	20.2*	21.0*	22.1*	15.3*

Data are means ± SD. *p<0.05 vs. men in the corresponding group by the non-paired t-test.
BMI = Body mass index

Table 2. Number (percentage) of participants with radiographic lumbar spondylosis and low back pain according to gender and age

		Radiographic lumbar spondylosis		Low back pain
		KL≥2	KL≥3	
Overall	2,288	1,728 (75.8)	1,149 (50.4)	659 (28.8)
Men	818	688 (84.1)	383 (46.8)	201 (24.6)
<70	154	114 (74.0)	51 (33.1)	35 (22.7)
70-79	491	419 (85.3)*	232 (47.3)*	119 (24.2)
80≤	173	155 (89.6)*	100 (57.8)*	47 (27.2)
Women	1,470	1,040 (70.7)†	766 (52.1)†	458 (31.2)†
<70	356	196 (55.1)	128 (36.0)	80 (22.5)
70-79	818	612 (74.8)*	456 (55.7)*	273 (33.4)*
80≤	296	232 (78.3)*	182 (61.5)*	105 (35.5)*

Radiographic spondylosis was determined at the severest level among L1/2-L5/S1.
*p<0.05 vs. subjects aged <70 by Scheffe's test after adjustment for BMI.
There was no significant difference between 70-79 and ≤80 in both genders.
†p<0.05 vs. men by logistic regression analysis after adjustment for age and BMI.

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Table 3. Association of gender and community with radiographic lumbar spondylosis and low back pain

| | Radiographic lumbar spondylosis | | | | Low back pain | |
| | KL≥2 | | KL≥3 | | | |
	OR	95%CI	OR	95%CI	OR	95%CI
Age, years	1.07	1.06-1.09†	1.05	1.04-1.07†	1.02	1.00-1.04*
BMI, kg/m^2	1.06	1.03-1.09†	1.04	1.01-1.06†	1.02	0.99-1.05
Women (vs. Men)	0.68	0.61-0.76†	1.13	1.03-1.23†	1.19	1.08-1.31†
Community (vs. Urban)						
Mountainous	0.82	0.65-1.04	0.56	0.45-0.69†	0.87	0.69-1.08
Seacoast	1.24	0.93-1.66	1.06	0.84-1.34	0.86	0.66-1.11

Radiographic spondylosis was determined at the severest level among L1/2-L5/S1.
The odds ratios were calculated by logistic regression analysis after adjustment for all other variables.
*p<0.05, †p<0.01
OR = odds ratio, CI = confidence interval

Table 4. Number (percentage) of subjects with radiographic lumbar spondylosis at each intervertebral level in all cohorts

| | KL≥2 | | KL≥3 | |
	Men	Women	Men	Women
L1/2	474 (57.9)	609 (41.4)	116 (14.2)	254 (17.3)
L2/3	541 (66.1)	749 (51.0)	164 (20.1)	355 (24.2)
L3/4	554 (67.7)	735 (50.0)	194 (23.7)	419 (28.5)
L4/5	523 (63.9)	736 (50.1)	306 (37.5)	605 (41.2)
L5/S	400 (48.9)	576 (39.2)	197 (24.2)	413 (28.1)

Table 5. Association of KL grade at each intervertebral level with low back pain

	L1/2		L2/3		L3/4		L4/5		L5/S		Severest	
	OR	95%CI	OR	95%CI	OR	95%CI	OR	95%CI	OR	95%CI	OR	95%CI
Men												
KL=2	1.30	0.92-1.84	0.94	0.65-1.36	1.43	0.98-2.11	1.24	0.82-1.89	1.12	0.75-1.65	1.15	0.70-1.9
KL≥3	1.30	0.79-2.11	1.25	0.80-1.94	1.49	0.96-2.32	1.42	0.97-2.08	1.22	0.82-1.81	1.44	0.89-2.3
Women												
KL=2	1.20	0.91-1.57	0.99	0.75-1.31	0.96	0.71-1.30	1.25	0.82-1.88	1.07	0.73-1.54	0.99	0.69-1.4
KL≥3	1.66	1.23-2.24*	1.74	1.32-2.30*	2.10	1.62-2.72*	1.88	1.48-2.38*	1.60	1.25-2.06*	1.80	1.38-2.3

The odds ratio was calculated by logistic regression analysis compared with subjects with KL grade 0 or 1 after adjustment for age and BMI. *p<0.01
OR = odds ratio, CI = confidence interval

Supplementary Table 1. Association of KL grade at the severest level with low back pain according to age

	Overall		<70		70-79		80≤	
	OR	95%CI	OR	95%CI	OR	95%CI	OR	95%CI
Men								
KL=2	1.15	0.70-1.92	0.94	0.49-1.86	2.78	0.96-9.38	0.70	0.23-2.37
KL≥3	1.44	0.89-2.38	1.37	0.74-2.62	2.50	0.82-8.71	0.80	0.28-2.53
Women								
KL=2	0.99	0.69-1.42	1.11	0.69-1.79	1.18	0.57-2.39	0.59	0.24-1.36
KL≥3	1.80	1.38-2.37†	1.93	1.34-2.80†	1.82	1.03-3.22*	1.39	0.77-2.57

The odds ratio was calculated by logistic regression analysis compared with subjects with KL grade 0 or 1 after adjustment for BMI. *$p<0.05$, †$p<0.01$
OR = odds ratio, CI = confidence interval

Supplementary Table 2. Association of KL grade at the severest level with low back pain according to community

	Overall		Urban		Mountainous		Seacoast	
	OR	95%CI	OR	95%CI	OR	95%CI	OR	95%CI
Men								
KL=2	1.15	0.70-1.92	1.96	0.87-4.88	0.54	0.25-1.16	1.75	0.48-8.46
KL≥3	1.44	0.89-2.38	2.49	1.16-6.01*	0.68	0.31-1.48	2.24	0.64-10.53
Women								
KL=2	0.99	0.69-1.42	1.15	0.68-1.93	0.72	0.39-1.31	1.21	0.49-2.88
KL≥3	1.80	1.38-2.37†	1.94	1.32-2.88†	1.52	0.93-2.51	1.80	0.94-3.56

The odds ratio was calculated by logistic regression analysis compared with subjects with KL grade 0 or 1 after adjustment for BMI. *$p<0.05$, †$p<0.01$
OR = odds ratio, CI = confidence interval

Figure 1

ORIGINAL ARTICLE
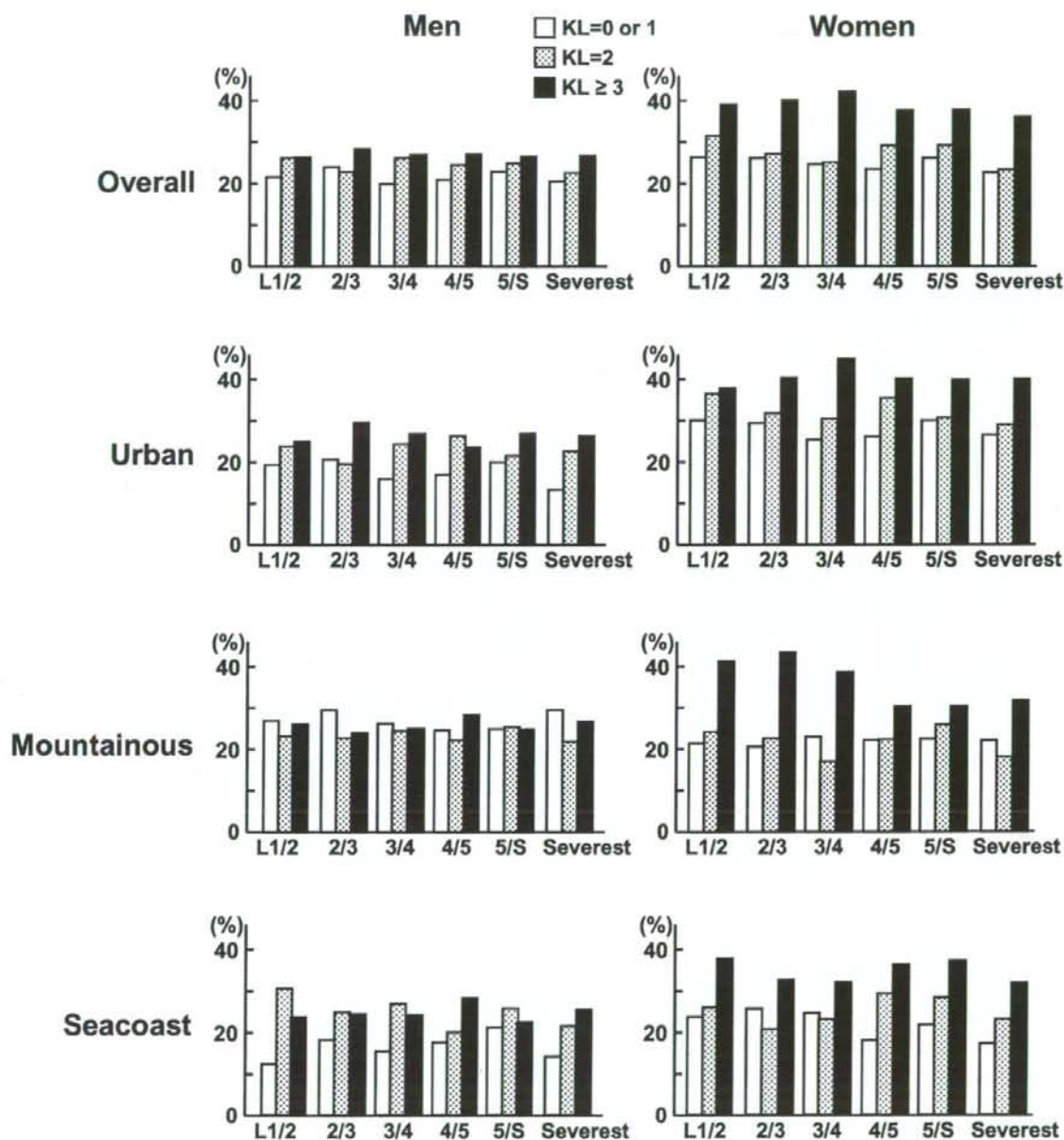
Epidemiology of lumbar osteoporosis and osteoarthritis and their causal relationship—is osteoarthritis a predictor for osteoporosis or vice versa?: The Miyama study

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Abstract

Summary In a 10-year follow-up of a population-based cohort of Japanese subjects, incidences of and causal relationships between osteoporosis (OP) and osteoarthritis (OA) at the lumbar spine were clarified. OP might reduce the risk of subsequent OA at the spine in women, but not in men.

Introduction The aim of this study is to clarify the contribution of osteoarthritis (OA) to osteoporosis (OP) and vice versa.

Methods A population-based, epidemiological study was conducted in a Japanese rural community. From 1,543 participants aged 40–79 years, 200 men and 200 women were selected and followed up for 10 years. Bone mineral density measurements were repeated after 3, 7, and 10 years, and X-rays were repeated after 10 years.

Results The incidence of lumbar OP per 10,000 person-years for persons in their 40s, 50s, 60s, and 70s was 0, 0, 109.5, and 151.1 for men and 124.2, 384.0, 227.3, and 239.5 for women, respectively. The cumulative incidence of lumbar OA over 10 years aged 40–79 years was 25.8% in men and 45.2% in women. Cox's proportional hazards model showed no significant relationship between the presence of lumbar OA at the baseline and incidence of lumbar and femoral neck OP in both genders. A significant relationship was demonstrated between the presence of lumbar OP, not femoral neck OP, at the baseline and cumulative incidence of lumbar OA in women (odds ratio, 0.20; 95% confidence interval, 0.05–0.80; $P=0.02$).

Conclusion OP in women appears to reduce the future incidence of OA at the lumbar spine.

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Introduction

As the proportion of aging population rapidly increases, the strategy for disease prevention is changing from simply extending life expectancy to extending healthy life expectancy in Japan. Thus, there is an urgent need for the development of methods for preventing musculoskeletal

disorders that impair activities of daily life (ADL) and quality of life (QOL) in the elderly. Osteoporosis (OP) and osteoarthritis (OA) are two major bone and joint health problems among the elderly that cause impairment of ADL and QOL, leading to increased morbidity and mortality. The estimated number of patients with OP in Japan is about 11 million [1], and the prevalence of this disease is the highest among bone metabolic diseases. Hip fracture is the most severe complication of OP, and is ranked third among diseases responsible for bedridden status, according to the National Livelihood Survey of the Ministry of Health, Labour and Welfare in Japan [2]. OP also increases mortality rate [3, 4]. The number of patients with OA has rapidly increased, and OA is now ranked second among the causes of disabilities requiring support for ADL in Japan [2].

Some studies have reported an inverse relationship between OP and OA [5–7]. A higher bone mineral density (BMD) in lumbar OA is well documented [8–11]. A decrease in the amount of bone in OP and the formation of bone spurs and increased amounts of bone in OA are evident from BMD measurements; radiography also reveals the opposing features of these two diseases. According to epidemiological studies, risk factors for the two diseases are in opposition. For example, low body weight is a risk factor for OP [12, 13], whereas high body weight represents a risk factor for OA [14, 15].

In contrast to previous opinions, however, recent studies have indicated the association of osteoporotic fractures with lumbar OA. Thus, narrowing of the intervertebral disc space was suggested to increase the risk of osteoporotic vertebral fractures [16, 17]. Although these results imply that lumbar OA should cause osteoporotic fractures, causal relationships between OP itself (not only osteoporotic fractures) and OA at the same site remain obscure. It is uncertain if OA causes OP, OP causes OA, the conditions only coexist, or OP and OA represent concomitant modifications of each other.

To clarify the contribution of OA to OP and vice versa in the general population, a 10-year follow-up study was performed on a cohort established in Miyama village, a rural Japanese community.

Materials and methods

Establishment of baseline cohort

This population-based, epidemiological study was initiated in 1990 in Miyama, a mountain village in Wakayama Prefecture, Japan. As the Miyama cohort has been profiled in detail elsewhere [18, 19], characteristics of the participants are briefly summarized here. A list of all inhabitants born in this village from 1910 to 1949, and therefore aged 40 to 79 years, was compiled from the register of residents

as of the end of 1989. A total cohort of 1,543 inhabitants (716 men, 827 women) was identified, and all members of the cohort completed a self-administered, 125-item questionnaire addressing topics such as dietary habits, smoking habits, alcohol consumption, and physical exercise.

A baseline BMD cohort was recruited from the total cohort, consisting of 400 participants divided into four groups each of 50 men and 50 women and stratified into age decades by year of birth (1910–1919, 1920–1929, 1930–1939, and 1940–1949). An interviewer administered a second questionnaire to these 400 participants, covering items of past medical history including questions related to osteoporotic fractures and falls, family history, calcium intake, dietary habits, physical exercise, occupational activities, sun exposure, and, for women, additional questions about reproductive variables. In addition to the baseline questionnaire survey, physical measurements were performed for participants including height (centimeter), body weight (kilogram), arm span (centimeter), bilateral grip strengths (kilogram) and circumstances of both wrists (centimeter), and body mass index (kilogram per square meter). These questionnaire surveys and measurements were repeated on the same 400 participants after 3, 7, and 10 years (1993, 1997, and 2000, respectively).

BMD measurements

The baseline BMD was measured in 1990 by dual energy X-ray absorptiometry (DXA; Lunar DPX, GE Medical Systems, Madison WI, USA), which provided anteroposterior images of lumbar vertebrae (L2–4) and the proximal femur (femoral neck, Ward's triangle, trochanter). These measurements were repeated on the same participants after 3, 7, and 10 years.

To control the precision of DXA, the equipment was checked at every examination in 1990, 1993, 1997, and 2000 using the same phantom. The BMD of the phantom was regulated to 1.270 ± 0.025 g/cm^2 (2%) during all examinations. In addition, the same physician (N.Y.) examined all participants in order to control observer variability. Intraobserver variability of DXA using the Lunar DPX in vitro and in vivo had been measured by the same physician for another study [20], and the coefficient of variance (CV) for L2–4 in vitro was 0.35%. The CV for L2–4, the proximal femur, Ward's triangle, and the trochanter examined in vivo in five male volunteers was 0.61–0.90%, 1.02–2.57%, 1.97–5.45%, and 1.77–4.17%, respectively.

OP was defined based on World Health Organization (WHO) criteria, in which OP was diagnosed mainly by that T-scores of BMD were lower than peak bone mass −2.5 standard deviations (SD) [21]. Mean L2–4 BMD for young adult men and women measured by Lunar DXA in Japan is 1.192 g/cm^2 while the SD is 0.146 g/cm^2 [22]. The present study therefore defined OP at the lumbar spine as L2–4

BMD <0.827 g/cm². Mean femoral neck BMD for young adult women measured by Lunar DXA in Japan is reportedly 0.914 g/cm² and the SD is 0.119 g/cm² [22]. OP at the femoral neck in women was defined as femoral neck BMD <0.617 g/cm². We could not define OP at the femoral neck in men because there was no reported mean femoral neck BMD for young adult men measured by Lunar DXA in Japan.

Radiography

The spine of each participant was examined by radiography in 1990. Diagnoses were based on anteroposterior and lateral images of thoracolumbar vertebrae Th5–L5 (initial X-ray survey). Radiography was repeated for individuals who provided consent after 10 years. Lateral images of thoracolumbar vertebrae Th5–L5 were again used for diagnosis (second X-ray survey).

Anteroposterior and lateral radiographs were scored for OA of the lumbar spine in L1–L5 using the Kellgren–Laurence (KL) grade as follows: KL0, normal; KL1, slight osteophytes; KL2, definite osteophytes; KL3, disc space narrowing with large osteophytes; KL4, bone sclerosis, disc space narrowing, and large osteophytes [23]. In the present study, we defined the lumbar spine with disc space narrowing with and without osteophytes as KL3. KL grade was determined at intervertebral spaces from L1/2 to L5/S1, and the highest score among all intervertebral spaces was then identified as the KL grade for that individual. KL scores of all radiographs were determined by a well-experienced orthopedist (S.M.).

Lateral radiographs of the spine were also utilized for the diagnosis of morphometric vertebral fracture (VFx) between Th5 and L5 using the criteria defined by the Japan Bone and Mineral Society as follows: wedged VFx, anterior height/posterior height ≤0.75; biconcave VFx, central height/anterior height or posterior height ≤0.80; compound VFx, anterior/anterior, central/central, and posterior/posterior height of sequential lower or upper vertebra ≤0.80 [24]. Diagnosis of VFx on all radiographs was performed by the same orthopedist (H.K.).

Detection of incidence of OP and OA

Incidence of OP over 10 years was calculated utilizing the results of BMD measurements at the baseline and follow-up studies after 3, 7, and 10 years. It was obtained by the following formula: the total number of incident cases with new OP divided by totaling the person-years of 'population at risk' at baseline. Population at risk refers to a group of participants having the potential of developing OP. Therefore, individuals with OP at the lumbar spine and femoral neck in the initial survey (lumbar spine, 13 men, 63

women; femoral neck, 46 women) were excluded from the numerators and denominators. To calculate the person-years, information on the drop-out (death or movement from the town) of participants was collected every year.

The cumulative incidence of OA over 10 years was calculated utilizing the diagnosis results. Cumulative incidence is simply defined as the ratio of incident cases to the population at risk at the beginning of the observation period. In the present study, we defined incident OA at the lumbar spine as KL grade ≥3 over 10 years in an individual whose KL grade ≤2 at the baseline.

The cumulative incidence of lumbar OA was determined by the following formula: individuals who developed new lumbar OA over 10 years/population at risk at the baseline. Individuals with existing lumbar OA with KL grade ≥3 at the baseline (69 men, 70 women) were excluded from both numerators and denominators.

Statistical analysis

All statistical analyses were performed using STATA statistical software (STATA Corp., College Station, TX, USA). Differences were tested for significance using ANOVA for comparison among multiple groups and using Scheffe's LSD test for pairs of groups. Significant items were selected, and multiple regression analysis was performed with adjustment of suitable variables.

To clarify the causal relationship of lumbar OA with OP, we applied Cox's proportional hazards model and calculated hazard ratio, in which the incidence of OP was used as an objective factor and lumbar OA at the baseline (1, yes vs. 0, no) was used as an explanatory factor. Next, to clarify the causal relationship of lumbar OA with osteoporotic fractures, we used logistic regression analysis using the cumulative incidence of morphometric VFx over 10 years (1, yes vs. 0, no) as an objective factor and lumbar OA at the baseline (1, yes vs. 0, no) as an explanatory factor, and obtained odds ratio (OR).

Furthermore, logistic regression analysis was used to assess causal relationships of: (a) OP at the lumbar spine and femoral neck with OA; (b) BMD at the lumbar spine L2–4 and femoral neck with OA; and (c) VFx with OA. In the analysis of OP and OA, we calculated the OR using the cumulative incidence of lumbar OA over 10 years (1, yes vs. 0, no) as an objective factor and OP at the baseline (1, yes vs. 0, no) as an explanatory factor. In the analysis of L2–4 and femoral neck BMD and OA, we calculated the OR using the cumulative incidence of lumbar OA over 10 years (1, yes vs. 0, no) as an objective factor and crude BMD values of the L2–4 and femoral neck at the baseline (vs. +1 SD) as an explanatory factor. Finally, in the analysis of VFx and OA, we obtained the OR using the cumulative incidence of lumbar OA over 10 years (1, yes vs. 0, no) as

an objective factor and the presence of VFx at the baseline (1, yes vs. 0, no) as an explanatory factor.

All data were analyzed in each gender group after adjustment for age and weight at the baseline.

Results

Eligible participants

A baseline BMD cohort comprising 400 participants was selected from the total cohort of 1,543 inhabitants. Characteristics of this baseline BMD cohort including anthropometric factors and BMD are shown in Table 1. Height, weight, and the body mass index (BMI; weight $(kg)/(height\ (m))^2$) for persons in their 70s were smaller than those for persons in their 40s and 50s for both men and women. BMD at the lumbar spine was significantly lower in men in their 60s and 70s than in their 40s. BMD at the lumbar spine in women tended to be lower with an increase in age and was significantly lower for women in their 50s, 60s, and 70s than in their 40s.

Of the 400 participants in the initial BMD examination, 390 provided written informed consent to participate in the initial X-ray survey (194 men, 196 women; 97.5%). Figure 1 shows the distribution of KL grades at the baseline for participants according to gender. The prevalence of KL grade ≥2 was 81.3% in men and 62.2% in women, and that of KL grade ≥3 was 35.8% in men and 35.7% in women.

Radiographic surveys after 10 years were performed for 299 (137 men, 162 women; 74.8%) of the 400 inhabitants. Data from 101 participants (63 men, 38 women) were unavailable due to the following reasons: 55 participants

died (37 men, 18 women); 16 moved (eight men, eight women); 13 were ill (four men, nine women); eight were busy (eight men); five declined to participate any further (five men); and four were absent from the area during the follow-up study (one man, three women).

A comparison of physical characteristics between completers and non-completers of the study has been described elsewhere [25] and is briefly summarized here. The height, weight, and BMI classified in terms of age group and gender were identical between completers and non-completers. In addition, the mean age of female completers in their 70s was significantly lower than that of female non-completers (mean (SD) of completers vs. mean (SD) of non-completers, 71.7 (1.8) years vs. 75.1 (2.8) years; $P<0.001$).

Prevalence of lumbar OP and OA and changes over 10 years

Table 2 shows the prevalence of lumbar OP and OA at the time of baseline measurements. Prevalence of lumbar OP in 1990 (baseline) and 2000 (over 10 years) were both significantly higher in women than men ($P<0.001$), while no significant difference was seen in the prevalence of lumbar OA in 1990 and 2000 between men and women. Prevalence of lumbar OP gradually increased with age in both men and women ($P<0.01$). However, age was not associated with the prevalence of lumbar OA in either men or women except female prevalence of lumbar OA in 2000 ($P<0.01$).

We then examined the prevalence of lumbar OP in the same age group of men and women in 2000, which was compared with that in 1990. Prevalence of lumbar OP in 1990 in the age group of 50–79 years was 8.7% in men

Table 1 Characteristics of the participants at the baseline measurement

Birth cohort	Age strata	N	Age (years)	Height (cm)	Weight (kg)	BMI (kg/m²)	BMD (g/cm²)
Men							
Total	40–79	200	58.9 (3.1)	160.9 (6.9)	57.6 (9.4)	22.1 (2.7)	1.11 (0.21)
1940–1949	40–49	50	44.2 (3.1)	165.6 (6.8)	63.6 (9.3)	23.1 (2.5)	1.19 (0.17)
1930–1939	50–59	50	54.1 (2.7)[a]	161.4 (5.7)[a]	59.5 (8.4)	22.8 (2.5)	1.15 (0.19)
1920–1929	60–69	50	63.4 (2.7)[a,b]	159.9 (5.5)[a]	56.1 (7.6)[a]	21.9 (2.4)	1.03 (0.18)[a,b]
1910–1919	70–79	50	73.9 (3.0)[a,b,c]	156.9 (6.8)[a,b]	51.0 (7.6)[a,b,c]	20.7 (2.7)[a,b]	1.06 (0.25)[a]
Women							
Total	40–79	200	59.3 (11.0)	148.3 (6.0)	48.8 (8.3)	22.1 (2.9)	0.95 (0.23)
1940–1949	40–49	50	44.7 (3.0)	152.4 (4.7)	53.2 (8.4)	22.8 (2.8)	1.18 (0.16)
1930–1939	50–59	50	54.8 (2.5)[a]	149.8 (5.3)	50.6 (7.4)	22.5 (2.7)	0.99 (0.18)[a]
1920–1929	60–69	50	64.3 (2.7)[a,b]	147.2 (5.0)[a]	47.1 (7.2)[a]	21.7 (3.1)	0.84 (0.19)[a,b]
1910–1919	70–79	50	73.3 (2.9)[a,b,c]	143.9 (5.7)[a,b,c]	44.5 (7.5)[a,b]	21.4 (2.9)[a,b]	0.78 (0.17)[a,b]

Data are means±SD

BMI body mass index, *BMD* bone mineral density

[a] Significantly different from values of the birth cohort group born in 1940–1949

[b] Significantly different from values of the birth cohort group born in 1930–1939

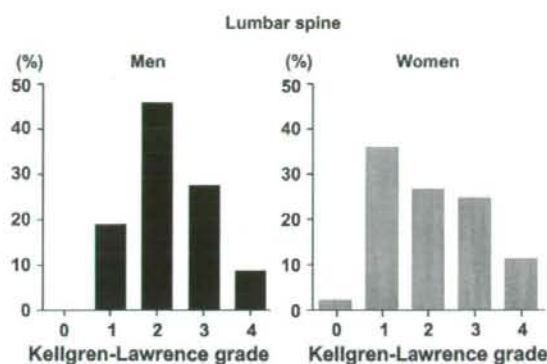[c] Significantly different from values of the birth cohort group born in 1920–1929

Fig. 1 Distribution of Kellgren–Lawrence grades at the lumbar spine by gender at the baseline in the Miyama population

and 42.0% in women and that in 2000 was 7.8% in men and 37.0% in women. Prevalence of lumbar OP in 2000 in the age group of 50–79 years tended to decrease compared with that in 1990 in both men and women, but no significant differences were identified (men $P=0.81$, women $P=0.39$).

Similarly, the prevalence of lumbar OA between the same age group of men and women in 2000 was compared with that in 1990. Prevalence in the age group of 50–79 years was 34.0% in men and 38.5% in women in 1990 and that in the same age group was 51.0% in men and 48.9% in women in 2000. Prevalence of lumbar OA in 2000 in the age group of 50–79 years increased in men and women compared with that in 1990, with significant differences in men (men $P<0.01$, women $P=0.08$).

Incidence of OP and cumulative incidence of OA at the lumbar spine

Figure 2 shows the incidence of lumbar OP in male and female participants of the cohort over 10 years. Incidence in men and women aged 40–79 years was 55.6 and 231.7 per 10,000 person-years, respectively. This means the annual incidence of lumbar OP among women is more than four times that of men.

The incidence of lumbar OP in men in their 40s, 50s, 60s, and 70s was 0, 0, 109.5, and 151.1 per 10,000 person-years, respectively, with the highest peak in the oldest group. In contrast, the incidence of lumbar OP in women in their 40s, 50s, 60s, and 70s was 124.2, 384.0, 227.3, and 239.5 per 10,000 person-years, respectively, with the highest peak for women in their 50s, the peri- and early postmenopausal periods, and another mild peak in the oldest group (Fig. 2). Incidence of OP at the femoral neck in women in their 40s, 50s, 60s, and 70s was 80.5, 221.9, 205.8, and 338.2 per 10,000 person-years, respectively, with the highest peak in the oldest age group and the second peak in their 50s.

The cumulative incidence of lumbar OA over 10 years aged 40–79 years was 25.8% in men and 45.2% in women. That for persons in their 40s, 50s, 60s, and 70s was 18.5%, 20.0%, 27.6%, and 37.9% for men and 37.1%, 53.6%, 48.4%, and 43.8% for women, respectively (Fig. 3). The cumulative incidence of lumbar OA tended to increase with age in men but not in women. The peak of the cumulative incidence of lumbar OA as well as that of lumbar OP in women was shown in the perimenopausal stratum. The cumulative incidence of lumbar OA was significantly higher in women than in men ($P<0.05$).

Table 2 Change of prevalence of osteoporosis and osteoarthritis at the lumbar spine over 10 years

Birth cohort	Baseline study					Follow-up study over 10 years			
	Age strata (years)	Number of participants (BMD)	Number of participants (X-ray)	Prevalence (%)		Age strata (years)	Number of participants	Prevalence (%)	
				Osteoporosis	Osteoarthritis[a]			Osteoporosis	Osteoarthritis[a]
Men									
Total	40–79	200	194	6.5	35.8	50–89	137	11.7	55.4
1940–1949	40–49	50	47	0.0	41.3	50–59	36	0.0	51.4
1930–1939	50–59	50	48	0.0	23.9	60–69	41	0.0	43.3
1920–1929	60–69	50	50	12.0	39.6	70–79	38	23.7	57.6
1910–1919	70–79	50	49	14.0	38.3	80–89	22	31.8	68.8
Women									
Total	40–79	200	196	31.5	35.7	50–89	162	42.6	54.1
1940–1949	40–49	50	48	0.0	27.1	50–59	49	12.2	35.4
1930–1939	50–59	50	49	18.0	42.9	60–69	46	45.7	50.0
1920–1929	60–69	50	50	48.0	38.0	70–79	40	57.5	64.1
1910–1919	70–79	50	49	60.0	34.7	80–89	27	70.4	83.3

[a] Osteoarthritis at the lumbar spine was defined as the KL grade ≥ 3
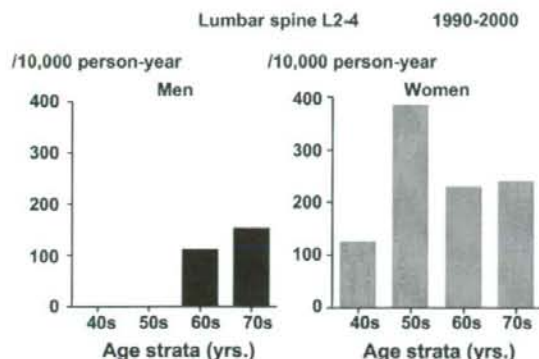
Lumbar spine L2-4   **1990-2000**

Fig. 2 Incidence of osteoporosis at the lumbar spine over 10 years by age group and gender

Causal relationship between OP and OA

The causal relationships between lumbar OA and OP, BMD, and VFx are summarized in Table 3.

First, the contribution of OA to OP was assessed. Cox's proportional hazard model showed no significant relationship between the presence of lumbar OA at the baseline and incidence of lumbar and femoral neck OP (lumbar OP, men $P=0.71$, women $P=0.79$; femoral neck OP, women $P=0.52$). Then, the association between lumbar OA and the cumulative incidence of VFx was determined by logistic regression analysis. As reported elsewhere, the cumulative incidence of VFx including subjects with previous VFx in their 40s, 50s, 60s, and 70s was 2.1%, 8.3%, 10.0%, and 12.2% for men and 2.1%, 6.1%, 18.0%, and 22.0% for women, respectively [26]. There was no significant relationship between the presence of lumbar OA at the baseline and incidence of VFx in men and women (men $P=0.21$, women $P=0.64$).

Secondly, the contribution of OP to OA was examined (Table 3). A significant relationship existed between the presence of lumbar OP at the baseline and cumulative incidence of lumbar OA in women ($P<0.05$) but not in men ($P=0.07$). Similarly, there was significant association between lumbar BMD at the baseline and the cumulative incidence of lumbar OA in women (vs. +1 SD, $P<0.05$) but not in men ($P=0.25$). No significant association was identified between femoral neck OP and BMD at the baseline and cumulative incidence of lumbar OA in men and women (OP at femoral neck, women $P=0.32$; BMD at femoral neck, vs. +1 SD, men $P=0.23$, women $P=0.77$). These results indicate that the presence of lumbar OP at the baseline would prevent the occurrence of lumbar OA, and conversely, high lumbar BMD would accelerate the progression of lumbar OA in women.

Finally, the association between the presence of VFx at the baseline and cumulative incidence of lumbar OA was assessed. As shown elsewhere, the prevalence of VFx in the present cohort among men in their 40s, 50s, 60s, and 70s was 4.3%, 14.6%, 22.0%, and 24.5% and that among women was 2.1%, 10.2%, 14.0%, and 44.9%, respectively [27]. Logistic regression analysis showed that there was no significant relationship between the presence of previous VFx and the incidence of lumbar OA in men and women (men $P=0.72$, women $P=0.91$; Table 3).

Discussion

The present study is a 10-year follow-up study of a population-based cohort of Japanese middle-aged people and elderly who were assessed for lumbar OP and OA. We clarified the prevalence of lumbar OP and OA and its trend of changes as well as the incidence of lumbar OP and cumulative incidence of lumbar OA. As for causal relationship, the presence of lumbar OA did not increase the risk of lumbar OP in both genders. However, the presence of lumbar OP significantly reduced the risk of lumbar OA, and high lumbar BMD values would accelerate the occurrence of lumbar OA over 10 years in women, while the presence of OP and BMD at the femoral neck did not influence the occurrence of lumbar OA.

The prevalence of lumbar OP in both 1990 and 2000 was significantly higher in women than in men ($P<0.001$) and gradually increased with age. Regarding the trend of changes in the prevalence of lumbar OP between 1990 and 2000 in same-age groups, no significant difference was shown in both men and women. We previously reported that both men and women in later birth cohorts showed higher BMDs in their middle age in this cohort [25]. However, we failed to clarify any significant decrease in the prevalence of lumbar OP in same-age groups of younger birth cohorts in the present study, although the prevalence
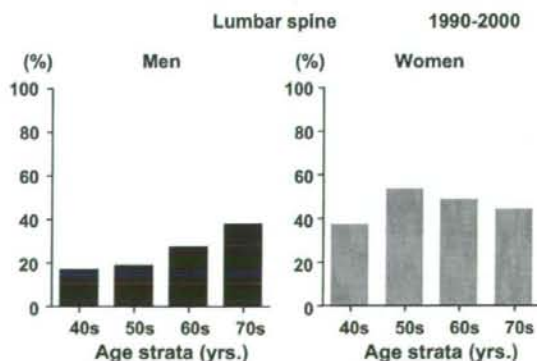

Lumbar spine   **1990-2000**

Fig. 3 Cumulative incidence of osteoarthritis at the lumbar spine over 10 years by age group and gender

Table 3 Causal relationship between osteoporosis (OP) and osteoarthritis (OA)

Baseline	Outcome	Reference	Gender	Risk ratio	95% CI	P value
Contribution of OA to OP						
OA at lumbar spine	Incidence of OP at lumbar spine	Yes/No	Men	HR 0.76	0.19–3.15	0.71
			Women	HR 0.90	0.40–1.99	0.79
OA at lumbar spine	Incidence of OP at femoral neck	Yes/No	Women	HR 0.74	0.30–1.84	0.52
OA at lumbar spine	Cumulative incidence of VFx	Yes/No	Men	OR 0.41	0.10–1.64	0.21
			Women	OR 1.27	0.46–3.47	0.64
Contribution of OP to OA						
OP at lumbar spine	Cumulative Incidence of OA at lumbar spine	Yes/No	Men	OR 8.68	0.82–92.3	0.07
			Women	OR 0.20	0.05–0.80	0.02
OP at femoral neck	Cumulative Incidence of OA at lumbar spine	Yes/No	Women	OR 0.52	0.14–1.89	0.32
BMD at lumbar spine	Cumulative incidence of OA at lumbar spine	+1 SD	Men	OR 0.80	0.54–1.17	0.25
			Women	OR 1.87	1.16–2.99	0.01
BMD at femoral neck	Cumulative incidence of OA at lumbar spine	+1 SD	Men	OR 0.80	0.56–1.15	0.23
			Women	OR 0.92	0.53–1.60	0.77
VFx	Cumulative incidence of OA at lumbar spine	Yes/No	Men	OR 0.79	0.21–2.95	0.72
			Women	OR 0.91	0.19–4.36	0.91

All analyses were adjusted for age and weight at the baseline

OA at lumbar spine was defined as the KL grade ≥ 3

BMD bone mineral density, *VFx* vertebral fracture, *SD* standard deviation, *HR* hazard ratio, *OR* odds ratio, *CI* confidence interval

of lumbar OP in 2000 tended to be lower than that in 1990 for all identical age groups in women. This might be explained by the effect of the time gap between the decrease in BMD and occurrence of lumbar OP. Although higher BMD was observed in the middle-aged group, this might not influence epidemiological indices of lumbar OP such as prevalence within only a 10-year span. As participants become old enough to be expected to have lumbar OP, its prevalence is expected to decrease.

Contrary to lumbar OP, the prevalence of lumbar OA was not significantly different between men and women in 1990 and 2000, and age was not associated with the prevalence of lumbar OA except for women in 2000 ($P<$ 0.01). Regarding the trend of changes in the prevalence of lumbar OA between 1990 and 2000 in same-age groups, the prevalence of lumbar OA in 2000 was higher than that in 1990 in both men and women, with significance in men (men $P<0.01$, women $P=0.08$). Concerning the association between age and lumbar OA, Lawrence found that the radiological prevalence of disc degeneration in the lumbar spine in the age group of 35–45 years increased with age [28]. O'Neill et al. reported that the frequency of vertebral osteophytes increased with age [29]. We previously compared the prevalence of lumbar OA determined by KL grade ≥ 3 in British and Japanese populations and reported that prevalence was higher in Britain than in Japan [15]. The difference may be partly explained by ethnic variation.

To the best of our knowledge, the present study represents the first report on the incidence of lumbar OP in Japan. If the incidence obtained in this study is generalized to the current

Japanese population in the age group of 40–79 years, 970,000 new cases of lumbar OP (160,000 men, 810,000 women) are estimated to occur annually. When classified by age, the incidence of lumbar OP in women was the highest in their 50s, followed by those in their 70s. We previously reported that the rate of change in lumbar spine BMD in women in the present population was the highest in their 50s [12, 25] and is related to the decrease in female hormones [30]. The present finding that the incidence of lumbar OP was the highest among women in their 50s suggests that the incidence of lumbar OP is closely related to the menstrual status, particularly menopause, and rate of change in lumbar spine BMD. Since more than 2.2% of women are estimated to develop lumbar OP annually in their 60s and 70s (ages at which the effects of menopause are thought to be attenuated), measures for preventing lumbar OP among the elderly as well as women during perimenopause are urgently required. The annual incidence of lumbar OP among men in their 60s and 70s was more than 1.0%. Although this incidence is lower than that among women, it is estimated that 160,000 male cases occur annually as previously mentioned, which nevertheless should not be ignored. Predictors for finding early and/or potential lumbar OP in both women and elderly men need to be established immediately.

In addition, we determined the cumulative incidence of lumbar OA with disc space narrowing for the first time in Japan. The 10-year cumulative incidence of lumbar OA with KL grade ≥ 3 tended to increase with age in men, but not in women, and it was higher in women than in men. Few reports have described the incidence of lumbar OA in

population-based cohorts. Hassett et al. showed that the progression rates for anterior osteophytes and disc space narrowing were 4% and 3% per year, respectively, among female participants in the Chingford study [31], which was approximately similar to the results of the present study. However, since epidemiological indices such as prevalence and incidence are highly dependent on the definition of OA, we cannot compare our results directly with those of other studies. For example, we defined lumbar OA as KL grade ≥3, which shows disc space narrowing with or without osteophytes, while the Chingford study determined lumbar OA based on the grading system of osteophytes and disc space narrowing reported by Lane et al. [32]. Since few reports have investigated the incidence of lumbar OA in the general population, further studies are needed to verify ethnic and geographical differences in the incidence of lumbar OA. When classified by age, the cumulative incidence of lumbar OA and OP was highest in women in their 50s during the early postmenopausal period. Therefore, it might be suggested that endogenous sex steroids play a role in the occurrence or progression of lumbar OA in women.

In some population-based prospective studies, OA of extremities was reported to increase the risk of osteoporotic fractures. In the Rotterdam study, knee OA increased the risk of vertebral and non-vertebral fractures [33]. Arden et al. reported that patients with knee OA and knee pain have an increased risk of hip and other non-vertebral fractures, which was not explained by the increased risk of falls [34]. Intervertebral disc space narrowing was found to increase the risk of VFx in the OFELY study [16, 17]. These findings suggest that OA is involved in the onset of fractures resulting from OP. Conversely, Roux et al. reported that intervertebral disc space narrowing and osteophytes decreased the prevalence of VFx in postmenopausal women with OP [35]. In the present study, there was no significant association between the presence of lumbar OA and future occurrence of lumbar OP and VFx. Lumbar OP is diagnosed by lumbar BMD (the value of which is easily affected by osteophytes and sclerosis of vertebrae and facets and the calcification of abdominal aorta [36]), which can artifactually increase BMD. Therefore, lumbar BMD might not be a good surrogate index of OP. As this is the first report about the causal relationship of lumbar OA and OP in the Japanese population, the difference might be partly due to the ethnic variation between Western and Oriental populations. Further studies are necessary to confirm the causal relationship of OA and OP in Japan and other countries.

Regarding the contribution of OP to OA, we elucidated that OP at the lumbar spine reduced the risk for the progression of lumbar OA in women while high BMD at the lumbar spine accelerated this progression.

Zhang et al. found that higher BMD at the hip was associated with prevalent and incident knee OA in older women in the Framingham study [37]. They also found that increased BMD over the follow-up period indicated a high risk of incident knee OA [37]. Hart et al. confirmed that, for women that developed incident knee OA, BMD was higher in the Chingford study [38]. Although these studies reported findings on the BMD and OA at extremities, not the spinal OP and OA, our results were almost similar to those of the above-mentioned cohort studies. Further prospective cohort studies with a larger sample size and longer observational periods are required to conclude the causal relationship of OP and OA.

Contrary to lumbar OP, no causal relationship was observed between OP or BMD at the femoral neck and cumulative incidence of lumbar OA. This might be because OP was diagnosed at different sites, which might have diluted the influence of OA occurrence. This hypothesis will be clarified in a study of the association between OP at the femoral neck and hip OA.

The presence of VFx at baseline showed no association with occurrence of lumbar OA. The prevalence of VFx includes various causes, and not all VFx were caused by OP. The geographic area in which the present cohort was established is mountainous, and a significant number of male subjects worked in the forestry industry and had experienced falls from trees or down slopes accidentally. In addition, most participants with previous VFx at the baseline were old and did not complete the 10-year follow-up. This survival bias might have influenced the evaluation of the influences of VFx on occurrence of OA.

The inverse causal relationship between lumbar OP and OA was only observed in women, not in men. These gender differences might be explained partly by differences in the incidence of lumbar OP. The incidence in men in the present study might be insufficient to detect the causal relationship. Alternatively, differences in gender-dependent factors such as endogenous sex steroids could influence the association of OP and OA.

There are several limitations in this study. The primary limitation is that the cohort comprised a relatively small number of participants. We were able to follow male and female residents with confirmed regional representativeness for 10 years with a high participation rate of 74.8%. However, 101 participants were lost in the follow-up study during the 10 years. The main reason for them dropping out of the study was death. The mean age of women completers of the age group 70-79 was significantly younger than that of drop-outs. Therefore, the prevalence of lumbar OP and cumulative incidence of lumbar OA in this age group might be underestimated due to the effects of survival bias. A secondary limitation is related to the definition of lumbar OA. Cumulative incidence as used in the present study was

detected by dividing the number of individuals who developed new lumbar OA by the number of participants in the follow-up study. Individuals with previous lumbar OA were excluded from both the numerators and denominators. In this formula, we excluded 69 male and 70 female participants with lumbar OA at the baseline to obtain the incidence of the first lumbar OA, which might reduce the total number of population at risk and cause a decrease in statistical power. Our result regarding lumbar OA incidence in the present study might need to be confirmed in larger population-based cohorts.

With the goal of elucidating the environmental and genetic background of bone and joint diseases represented by OA and OP, we established larger scale cohorts based on the present cohort, called Research on Osteoarthritis/Osteoporosis Against Disability (ROAD), and have already started the follow-up study [39]. This enlarged population-based cohort study may confirm the consistency of epidemiological trends for OP and OA and clarify the causal relationship between these two major bone and joint diseases.

Conclusion

Based on observations from a population-based cohort over a 10-year period, the estimated incidence of OP at the L2–4 level of the lumbar spine per 10,000 person-years for men in their 40s, 50s, 60s, and 70s was 0, 0, 109.5, and 151.1 and that for women was 124.2, 384.0, 227.3, and 239.5, respectively. The cumulative incidence of lumbar OA over 10 years for men in their 40s, 50s, 60s, and 70s was 18.5%, 20.0%, 27.6%, and 37.9% for men and 37.1%, 53.6%, 48.4%, and 43.8% for women, respectively. Cox's proportional hazards model showed no significant relationship between the presence of lumbar OA at the baseline and future incidence of lumbar and femoral neck OP. A significant relationship existed between the presence of lumbar OP at the baseline and future incidence of lumbar OA in women (odds ratio 0.20, 95% confidence interval 0.05–0.80, $P<0.05$). It may be suggested that the presence of OA does not increase the risk of incident OP in both genders and that the presence of OP reduces the risk of incident OA at the spine in women.

References

1. Yamamoto I (1999) Estimation for the number of patients of osteoporosis in Japan. Osteoporosis Jpn 7:10–11 (in Japanese)
2. Ministry of Health, Labour and Welfare. Outline of the results of National Livelihood Survey 2004. http://www.mhlw.go.jp/toukei/saikin/hw/k-tyosa/k-tyosa04/4-2.html
3. Muraki S, Yamamoto S, Ishibashi H, Nakamura K (2006) Factors associated with mortality following hip fracture in Japan. J Bone Miner Metab 24:100–104
4. Jornell O, Kanis JA, Oden A, Sembo I, Redlund-Johnell I, Petterson C, De Laet C, Jonsson B (2004) Mortality after osteoporotic fractures. Osteoporosis Int 15:38–42
5. Sambrook P, Naganathan V (1997) What is the relationship between osteoarthritis and osteoporosis? Baillieres Clin Rheumatol 11:695–710
6. Dequecker J, Boonen S, Aerssens J, Westhovens R (1996) Inverse relationship osteoarthritis–osteoporosis: what is the evidence? What are the consequences? Br J Rheumatol 35:813–818
7. Dequecker J, Aerssens J, Luyten FP (2003) Osteoarthritis and osteoporosis: clinical and research evidence of inverse relationship. Aging Clin Exp Res 15:426–439
8. Jones G, Nguyen T, Sambrook PN, Lord SR, Kelly PJ, Eisman JA (1995) A longitudinal study of the effect of spinal degenerative disease on bone density in the elderly. J Rheumatol 22:932–936
9. Liu G, Peacock M, Eilam O, Dorulla G, Braunstein E, Johnston CC (1997) Effect of osteoarthritis in the lumbar spine and hip bone mineral density and diagnosis of osteoporosis in elderly men and women. Osteoporos Int 7:564–569
10. Hart DJ, Mootoosamy I, Doyle DV, Spector TD (1994) The relationship between osteoarthritis and osteoporosis in the general population: the Chingford study. Ann Rheum Dis 53:158–162
11. Belmonte-Serrano MA, Bloch DA, Lane NE, Michel BE, Fries JF (1993) The relationship between spinal and peripheral osteoarthritis and bone density measurements. J Rheum 20:1005–1013
12. Yoshimura N, Hashimoto T, Morioka S, Sakata K, Kasamatsu T, Cooper C (1998) Determinants of bone loss in a rural Japanese community. The Taiji study. Osteoporos Int 8:604–610
13. De Laet C, Kanis JA, Oden A, Johanson H, Johnell O, Delmas P, Eisman JA, Kroger H, Fujiwara S, Garnero P, McCloskey EV, Mellstrom D, Melton LJ 3rd, Meunier PJ, Pols HA, Reeve J, Silman A, Tenenhouse A (2005) Body mass index as a predictor of fracture risk: a meta-analysis. Osteoporos Int 16:1330–1338
14. Hartz AJ, Fischer ME, Bril G, Kelber S, Rupley D Jr, Oken B, Rimm AA (1986) The association of obesity with joint pain and osteoarthritis in the HANES data. J Chronic Dis 39:311–319
15. Yoshimura N, Dennison E, Wilman C, Hashimoto T, Cooper C (2000) Epidemiology of chronic disc degeneration and osteoarthritis of the lumbar spine in Britain and Japan: a comparative study. J Rheumatol 27:429–433
16. Sornay-Rendu E, Munoz F, Duboeuf F, Delmas PD (2004) Disc space narrowing is associated with an increased vertebral fracture

risk in postmenopausal women: the OFELY Study. J Bone Miner Res 19:1994–1999

17. Somay-Rendu E, Allard C, Munoz F, Duboeuf F, Delmas PD (2006) Disc space narrowing as a new risk factor for vertebral fracture: the OFELY study. Arthritis Rheum 54:1262–1269

18. Kasamatsu T, Morioka S, Hashimoto T, Kinoshita H, Yamada H, Tamaki T (1991) Epidemiological study on bone mineral density of inhabitants in Miyama Village, Wakayama Prefecture (Part 1). Background of study population and sampling method. J Bone Miner Metab 9(suppl):50–55

19. Kinoshita H, Danjoh S, Yamada H, Tamaki T, Kasamatsu T, Ueda A, Hashimoto T (1991) Epidemiological study on the bone mineral density of inhabitants in Miyama Village, Wakayama Prefecture (part II). Bone mineral density of the spine and proximal femur. J Bone Miner Metab 9(suppl):56–60

20. Yoshimura N, Kakimoto T, Nishioka M, Kishi T, Iwasaki H, Niwa T, Morioka S, Sakata T, Hashimoto T (1997) Evaluation of reproducibility of bone mineral density measured by dual energy X-ray absorptiometry (Lunar DPX-L). J Wakayama Medical Society 48:461–466

21. World Health Organization (1994) Assessment of fracture risk and its application to screening for postmenopausal osteoporosis. WHO technical report series 843. WHO, Geneva

22. Orimo H, Hayashi Y, Fukunaga M, Sone T, Fujiwara S, Shiraki M, Kushida K, Miyamoto S, Soen S, Nishimura J, Oh-Hashi Y, Hosoi T, Gorai I, Tanaka H, Igai T, Kishimoto H (2001) Osteoporosis Diagnostic Criteria Review Committee: Japanese society for bone and mineral research. Diagnostic criteria for primary osteoporosis: year 2000 revision. J Bone Miner Metab 19:331–337

23. Kellgren JH, Lawrence LS (1957) Radiological assessment of osteo-arthrosis. Ann Rheum Dis 16:494–502

24. Inoue T (1990) Clinical features and findings, osteoporosis (in Japanese). Bone 4:39–47

25. Yoshimura N, Kinoshita H, Danjoh S, Takijiri T, Morioka S, Kasamatsu T, Sakata K, Hashimoto T (2002) Bone loss at the lumbar spine and the proximal femur in a rural Japanese community, 1990–2000: the Miyama study. Osteoporos Int 13:803–808

26. Yoshimura N, Kinoshita H, Oka H, Muraki S, Mabuchi A, Kawaguchi H, Nakamura K (2006) Cumulative incidence and changes in prevalence of vertebral fractures in a rural Japanese community: a 10-year follow-up of the Miyama cohort. Arch Osteoporos 1:43–49 doi:10.1007/s11657-006-0007-0

27. Yoshimura N, Kinoshita H, Danjoh S, Yamada H, Tamaki T, Morioka S, Kasamatsu T, Hashimoto T, Inoue T (1995) Prevalence of vertebral fractures in a rural Japanese population. J Epidemiol 5:171–175

28. Lawrence JS (1969) Disc degeneration. Its frequency and relationship to symptoms. Ann Rheum Dis 28:121–138

29. O' Neill TW, McCloskey EV, Kanis JA, Bhalla AK, Reeve J, Reid DM, Todd C, Woolf AD, Silman AJ (1999) The distribution, determinants, and clinical correlates of vertebral osteophytosis: a population based survey. J Rheumatol 26:842–848

30. Yoshimura N, Kasamatsu T, Sakata K, Hashimoto T, Cooper C (2002) The relationship between endogenous estrogen, sex hormone binding globulin and bone loss in female residents of a rural Japanese community: the Taiji study. J Bone Miner Metab 20:303–310

31. Hassett G, Hart DJ, Manek NJ, Doyle DV, Spector TD (2003) Risk factors for progression of lumbar spine disc degeneration, the Chingford study. Arthritis Rheum 48:3112–3117

32. Lane N, Nevitt MC, Genant HK, Hochberg MC (1993) Reliability of new indices of radiographic osteoarthritis of the hand and hip and lumbar disc degeneration. J Rheumatol 20:1911–1918

33. Bergink AP, van der Klift M, Hofman A, Verhaar JA, van Leeuwen JP, Uitterlinden AG, Pols HA (2003) Osteoarthritis of the knee is associated with vertebral and nonvertebral fractures in the elderly: the Rotterdam study. Arthritis Rheum 49:648–657

34. Arden NK, Croziew S, Smith H, Anderson F, Edwards C, Raphael H, Cooper C (2006) Knee pain, knee osteoarthritis, and the risk of fracture. Arthritis Rheum 55:610–615

35. Roux C, Fechtenbaum J, Briot K, Cropet C, Liu-Léage S, Marcelli C (2008) Inverse relationship between vertebral fractures and spine osteoarthritis in postmenopausal women with osteoporosis. Ann Rheum Dis 67:224–228

36. Kinoshita H, Tamaki T, Hashimoto T, Kasagi F (1998) Factors influencing lumbar spine bone mineral density assessment by dual energy X-ray absorptiometry: comparison with lumbar spinal radiogram. J Orthop Sci 3:3–9

37. Zhang Y, Hannan MT, Chaisson CE, McAlindon TE, Evans SR, Aliabadi P, Levy D, Felson DT (2000) Bone mineral density and risk of incident and progressive radiographic knee osteoarthritis in women: the Framingham study. J Rheumatol 27:1032–1037

38. Hart DJ, Cronin C, Daniels M, Worthy T, Doyle DV, Spector TD (2002) The relationship of bone density and fracture to incident and progressive radiographic osteoarthritis of the knee: the Chingford Study. Arthritis Rheum 46:92–99

39. Muraki S, Oka H, Mabuchi A, Akune T, En-yo Y, Yoshida M, Saika A, Suzuki T, Yoshida H, Ishibashi H, Yamamoto S, Nakamura K, Kawaguchi H, Yoshimura N (2008) Prevalence of radiographic lumbar spondylosis and its association with low back pain in the elderly of population-based cohorts: the ROAD study. Ann Rheum Dis (in press)